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The impact of the COVID-19 pandemic on air pollution: A global assessment using machine learning techniques

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Abstract

In response to the COVID-19 pandemic, most countries implemented public health ordinances that resulted in restricted mobility and a resultant change in air quality. This has provided an opportunity to quantify the extent to which carbon-based transport and industrial activity affect air quality. However, quantification of these complex effects has proven to be difficult, depending on the stringency of restrictions, country-specific emission source profiles, long-term trends and meteorological effects on atmospheric chemistry, emission levels and in-flow from nearby countries. In this study, confounding factors were disentangled for a direct comparison of pandemic-related reductions in absolute pollution levels, globally. The non-linear relationships between atmospheric processes and daily ground-level NO₂, PM₁₀, PM_{2.5} and O₃ measurements were captured in city- and pollutant-specific XGBoost models for over 700 cities, adjusting for weather, seasonality and trends. City-level modelling allowed adaptation to the distinct topography, urban morphology, climate and atmospheric conditions for each city, individually, as the weather variables that

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were most predictive varied across cities. Pollution forecasts for 2020 in absence of a pandemic were generated based on weather and formed an ensemble for country-level pollution reductions. Findings were robust to modelling assumptions and consistent with various published case studies. NO_2 reduced most in China, Europe and India, following severe government restrictions as part of the initial lockdowns. Reductions were highly correlated with changes in mobility levels, especially trips to transit stations, workplaces, retail and recreation venues. Further, NO_2 did not fully revert to pre-pandemic levels in 2020. Ambient $\text{PM}_{2.5}$ pollution, which has severe adverse health consequences, reduced most in China and India. Since positive health effects could be offset to some extent by prolonged exposure to indoor pollution, alternative transport initiatives could prove to be an important pathway towards better health outcomes in these countries. Increased O_3 levels during initial lockdowns have been documented widely. However, our analyses also found a subsequent reduction in O_3 for many countries below what was expected based on meteorological conditions during summer months (e.g., China, United Kingdom, France, Germany, Poland, Turkey). The effects in periods with high O_3 levels are especially important for the development of effective mitigation strategies to improve health outcomes.

Key words: COVID-19, Nitrogen dioxide, Particulate matter, Ozone, Mobility, Decision trees

1. Introduction

At the end of 2019, a new, highly infectious and deadly coronavirus was detected, transmitted via human-to-human contact (Riou and Althaus, 2020). Since then, the rapid spread of various variants of SARS-CoV-2 has put unprecedented pressure on economies and healthcare systems, resulting in significant morbidity and a death toll of several million people (World Health Organization, 2021). To slow disease transmission, many countries implemented public health ordinances that included an array of measures such as mask wearing, quarantin-

9 ing of positive cases and reduced population mobility across and between cities.
10 Even before travel and work restrictions were enacted and the World Health
11 Organization (WHO) declared a pandemic on 11 March 2020, mobility dropped
12 dramatically (Google, 2020). By April 2020, more than half of the world’s pop-
13 ulation had reduced their travel by more than 50% (Forster et al., 2020; Le
14 Quéré et al., 2020). Heavy restrictions resulted in deserted cities, empty roads
15 and clear skies in cities frequently blanketed in air pollution (e.g., He et al.,
16 2020). The enforced reductions in mobility and industrial activity provide a
17 natural experiment to explore the effects of these activities on air pollution. For
18 example, it provides an opportunity to investigate the effects on air quality of
19 decarbonising the road transport system as part of climate change mitigation.

20 *1.1. Quantifying the impact of COVID-19 lockdowns*

21 During the first seven months following the start of the pandemic, over 200
22 papers were accepted for publication that explored changes in air pollution lev-
23 els (Gkatzelis et al., 2021). Not all pollutants were affected equally, due to the
24 different sources of pollution. For example, secondary pollutants such as O_3
25 form through chemical reactions of NO_x and volatile organic compounds under
26 the influence of solar radiation (Seinfeld and Pandis, 2016). NO_2 is mainly emit-
27 ted through fuel combustion from road transport and fossil fuel power plants.
28 $PM_{2.5}$ is also emitted in these sectors, but mostly from diesel motor vehicles
29 (e.g., commercial heavy-duty diesel trucks) and coal-fired power stations. It
30 is also generated in other sectors, for example, through industrial activity and
31 coal-fired winter heating (He et al., 2020). Thunis et al. (2018) allocated $PM_{2.5}$
32 emissions in Europe to agriculture (23%), industry (20%), natural (19%), trans-
33 port (14%) and residential (13%) sources. In Scotland, $PM_{2.5}$ pollution depends
34 mostly on natural and non-traffic sources (Dobson and Semple, 2020). Although
35 road transport strongly declined during lockdown periods, some locations were
36 mainly exposed to pollution sources that were affected to a limited extent by
37 the pandemic (e.g., agriculture, natural sources). Hence, besides the stringency
38 of restrictions, the mix of emission sources in a specific country influences the

39 reduction in air pollution that might be expected.

40 *1.1.1. Nitrogen dioxide*

41 Studies used either ground-level measurements or satellite remote sensing to
42 quantify the reduction in NO₂ during the initial lockdown period (i.e., March/April
43 2020 for most countries). In the United States, reported NO₂ reductions were
44 20% compared to the corresponding period in 2010–2019 (Bekbulat et al., 2021)
45 and 21.6% compared to 2019 (Goldberg et al., 2020). Other studies found av-
46 erage reductions of 1.3 parts per billion (reported as –27% compared to 2015–
47 2019, before de-trending) (Archer et al., 2020) to 4.8 parts per billion (reported
48 as –25.5% compared to 2017–2019) (Berman and Ebisu, 2020), with the largest
49 reductions in urban counties. Sharma et al. (2020) found an average NO₂ reduc-
50 tion of 18% compared to 2017–2019 in 22 cities across India, while Mahato et al.
51 (2020) reported a 53% reduction compared to the pre-lockdown period in Delhi.
52 In Wuhan, reported reductions were 53.3% compared to the pre-lockdown pe-
53 riod (Lian et al., 2020) and 57% compared to 2017–2019 (Sicard et al., 2020).
54 Sicard et al. (2020) also found a 53% NO₂ reduction in selected European cities
55 with high pollution levels (i.e., Nice, Rome, Valencia and Turin), while Bauwens
56 et al. (2020) reported reductions of 27% compared to 2019 in Western Europe
57 and 40% in China. Reductions varied substantially between cities (Goldberg
58 et al., 2020). Connerton et al. (2020) found NO₂ reductions of 24% compared
59 to 2015–2019 in New York City, 25% in São Paulo, 38% in Los Angeles, and
60 39% in Paris.

61 NO₂ reductions were short-lived and returned towards normal ranges after
62 the lockdowns ended. For example, Dentener et al. (2020) observed reductions
63 in the average NO₂ tropospheric column of over 50% during March (compared to
64 2019) in a selection of major Asian cities. However, over the three-month period
65 from March to May 2020, reductions were much lower at 15–20%. Similarly, the
66 observed three-month reductions were 20% in Germany and the Benelux, 15%
67 in Italy, 10–15% in North America, Spain, France, the United Kingdom, Poland
68 and Czech Republic, and 8% in Romania.

69 *1.1.2. Particulate matter*

70 For particulate matter, results were mixed with large reductions in some
71 countries and no significant effects in others. Substantial PM_{2.5} reductions of
72 36% compared to 2017–2019 (Sicard et al., 2020) and 36.9% compared to the
73 pre-lockdown period (Lian et al., 2020) were reported for Wuhan. He et al.
74 (2020) found PM_{2.5} reductions of 21.1 and 7.1 $\mu\text{g}/\text{m}^3$ compared to 2019 in
75 Chinese cities with and without formal lockdowns, respectively. Effects were
76 larger in colder, richer, and more industrialised cities. Similarly, Giani et al.
77 (2020) found an average reduction of 14.5 $\mu\text{g}/\text{m}^3$ in population-weighted PM_{2.5}
78 across China, compared to 2016–2019. In the areas most affected by COVID-19,
79 two-month average reductions up to 40 $\mu\text{g}/\text{m}^3$ were observed. In India, Sharma
80 et al. (2020) found an average reduction in PM_{2.5} of 43% compared to 2017–
81 2019 across 22 cities, while Mahato et al. (2020) reported a reduction of 39%
82 compared to 2019 in Delhi.

83 In contrast, the average PM_{2.5} reduction in Europe was found to be only
84 2.2 $\mu\text{g}/\text{m}^3$ compared to 2016–2019 (Giani et al., 2020). Reductions in Scotland
85 were very limited (i.e., within 1 $\mu\text{g}/\text{m}^3$ compared to 2017 and 2018), even though
86 motor vehicle journeys reduced by 65% (Dobson and Semple, 2020). Further,
87 Sicard et al. (2020) reported an average reduction of 4% compared to 2017–
88 2019 in four cities with high pollution levels in Southern Europe. In the United
89 States, a slight increase in PM_{2.5} was observed of 0.28 $\mu\text{g}/\text{m}^3$ compared to April
90 2015–2019 (Archer et al., 2020). Bekbulat et al. (2021) found a 10% increase in
91 PM_{2.5} compared to 2010–2019 and concluded this was within the normal range
92 of variability.

93 *1.1.3. Ozone*

94 In contrast to the reductions observed for other pollutants, many studies
95 found increased O₃ during lockdown periods. For example, O₃ increases of 17%
96 compared to 2017–2019 were reported for both Europe (Sicard et al., 2020) and
97 India (Sharma et al., 2020). Sicard et al. (2020) reported a 36% increase com-
98 pared to 2017–2019 in Wuhan, while Lian et al. (2020) found an increase of

99 116.6% compared to the pre-lockdown period. Increased O₃ levels were also re-
100 ported for São Paulo (30% compared to 2015–2019), Paris (12%) and New York
101 City (7%), while a decrease of 10% was observed for Los Angeles (Connerton
102 et al., 2020). Many studies discussed atmospheric chemistry for O₃ formation as
103 the potential cause of observed ozone increases. O₃ titration occurs particularly
104 during winter months if the NO_x level is high, reducing the O₃ level (Sillman,
105 1999). Hence, reduced NO_x during the initial lockdowns could lead to higher
106 O₃ pollution.

107 1.2. Limitations and opportunities

108 Many papers present case studies for a specific country (e.g., China (He
109 et al., 2020), Scotland (Dobson and Semple, 2020)) or a selection of cities (e.g.,
110 Connerton et al., 2020; Sicard et al., 2020). However, very few studies have per-
111 formed city-level analyses at a global scale. Global studies have the advantage
112 that they allow for a consistent comparison between countries. One example of
113 a global analysis is the study by Venter et al. (2020), finding substantial varia-
114 tions in country-level temporal and spatial pollution anomalies for NO₂, PM_{2.5}
115 and O₃ that remain unexplained. Although not considering cities individually,
116 they compared pollution changes across 34 different countries.

117 Multiple confounding factors complicate the attribution of air pollution to
118 pandemic-related changes. For example, long term trends in air pollution can
119 be observed for some pollutants, as government policies shift towards renewable
120 energy sources. Several studies did not account for long-term trends in pollution
121 levels, which could lead to incorrectly attributing a decrease in air pollution to
122 actions taken during the pandemic. Further, various studies compared 2020
123 measurements to equivalent periods in previous years (e.g., 2019, or 2017–2019)
124 without adjustments for weather effects (e.g., Bauwens et al., 2020; Berman
125 and Ebisu, 2020; Sicard et al., 2020). This approach can have limitations if the
126 weather in 2020 differs from atmospheric conditions in the equivalent historical
127 period (e.g, more rainfall or higher temperatures). Goldberg et al. (2020) found
128 that weather variations between years can cause fluctuations in monthly NO₂

129 levels of approximately 15%. Further, Wang et al. (2020) found that benefits
130 of emission reductions in China during the lockdown period were sometimes
131 overwhelmed by adverse meteorology, leading to severe air pollution events.
132 These examples indicate that, in absence of weather adjustments, pollution
133 estimates need to be averaged over extended time periods to reduce weather
134 influences; an approach that is taken by many studies (as mentioned above). The
135 importance of weather adjustments is frequently acknowledged, but not often
136 addressed. Specifically, Gkatzelis et al. (2021) found that two thirds of studies
137 that quantified changes in air pollution did not correct for weather effects.

138 Studies that did adjust for weather influences used a variety of approaches,
139 including difference-in-difference methods (e.g., Navinya et al., 2020), gener-
140 alised additive models (e.g., Ropkins and Tate, 2021) and atmospheric chemical
141 transport models (e.g., Zhao et al., 2020). Further, several studies have investi-
142 gated the use of machine learning approaches. For example, Grange and Carslaw
143 (2019) investigated the normalisation of time series with respect to weather ef-
144 fects using a random forest model, obtaining time series under average weather
145 conditions. The use of non-parametric machine learning methods for meteoro-
146 logical normalisation has considerable advantages over parametric alternatives
147 such as regression analysis, as atmospheric processes are complex, non-linear,
148 and variables frequently exhibit high multicollinearity (Grange and Carslaw,
149 2019). Further, Petetin et al. (2020) quantitatively illustrated the benefits of
150 using a machine learning approach to incorporate meteorological variability, as
151 opposed to directly comparing measurements in previous years to 2020 levels.
152 Importantly, machine learning is especially beneficial for accurately predicting
153 pollution at finer spatial and temporal scales (Petetin et al., 2020). Therefore,
154 for a robust comparison to daily city-level mobility data (i.e., fine spatial and
155 temporal scale), accurate weather-normalised pollution estimates are a prereq-
156 uisite.

157 With respect to the selection of a particular machine learning approach, Ma
158 et al. (2020a), Zamani Joharestani et al. (2019) and Ren et al. (2020) all found
159 that XGBoost (Chen and Guestrin, 2016) performed better for air pollution

160 modelling than other machine learning techniques such as random forest and
161 support vector regression. Further, Ma et al. (2020b) showed that an XGBoost
162 model significantly improved the $PM_{2.5}$ forecasts of an operational atmospheric
163 chemical-transport model at the Shanghai Meteorological Service. Our research
164 should be seen in light of these developments. In this paper, we present a
165 city-level analysis at a global scale, including adjustment for weather variables
166 such as temperature, wind speed and precipitation. City- and pollution-specific
167 modelling based on historical atmospheric data and annual trends, provides
168 air pollution estimates for 2020 in absence of a pandemic (i.e., a counterfac-
169 tual, 'business as usual' scenario). Various methods were explored to optimise
170 forecast accuracy. The identified pollution anomalies were then compared to
171 country-specific government policies intended to limit the spread of COVID-19.
172 Hence, our research investigates air pollution both at the micro- and the macro-
173 level. It explores how different countries obtained different results based on the
174 stringency of restrictions and associated changes in city-level mobility patterns.

175 **2. Materials and methods**

176 *2.1. Data*

177 The following global data sources were selected to provide information on air
178 pollution levels, weather, the severity of government restrictions, and mobility
179 patterns of city residents. Data was checked for completeness, cleaned and
180 processed using Java, Python and R.

181 *2.1.1. Pollution*

182 The World Air Quality Index project (AQICN, 2021) provides ground-level
183 readings of pollutants for cities in 132 countries sourced from world-wide envi-
184 ronmental protection agencies. Daily measurements of NO_2 , PM_{10} , $PM_{2.5}$ and
185 O_3 were downloaded over the period 2015–2020 for approximately 900 cities
186 (i.e., a subset of the 1692 largest cities in the world (United Nations, 2015) for
187 which sufficient data was available). All measurements were a 24-hour average

188 of hourly readings, accounting for the local time zone. Since the data provider
189 quoted pollution measurements in terms of AQI levels based on standards of
190 the Environmental Protection Agency (Gilliam and Hall, 2016), measurements
191 were converted back to their original unit for comparative analysis in this study.
192 Our final dataset consisted of NO₂ and O₃ measurements in parts per billion
193 (ppb) and PM₁₀ and PM_{2.5} measurements in $\mu\text{g}/\text{m}^3$.

194 *2.1.2. Weather*

195 Meteorological data was obtained from the ERA5 reanalysis (Hersbach et al.,
196 2020) of the European Centre for Medium-Range Weather Forecasts, providing a
197 large range of atmospheric, land and oceanic climate variables. Importantly, the
198 ERA5 reanalysis combines past observations with models to generate globally
199 consistent time series, while limiting missing data. Hourly (solar radiation and
200 precipitation only) and 8-hourly ERA5 data was downloaded at a $0.25^\circ \times 0.25^\circ$
201 resolution via the Copernicus Climate Change Service Climate Data Store. Us-
202 ing each city’s latitude and longitude coordinates (United Nations, 2015), daily
203 city-level information was extracted from the grid cell containing the location
204 of the city centre. The following atmospheric and land-based variables were
205 obtained for each 24-hour period based on the local time zone:

- 206 – mean air temperature at 2 meters above the surface (K);
- 207 – total net solar radiation at the surface (Jm^{-2});
- 208 – total precipitation (m);
- 209 – mean wind speed (ms^{-1});
- 210 – mean wind direction (degrees); and
- 211 – mean leaf area index of vegetation (m^2m^{-2}).

212 *2.1.3. Stringency index*

213 Governments adopted a broad variety of (initial) approaches to deal with the
214 COVID-19 pandemic, ranging from trying to maintain business as usual (e.g.,
215 Brazil, Sweden), to imposing strict lockdowns and border closures in attempts

216 to eliminate the virus (e.g., Taiwan, New Zealand). A globally consistent as-
217 sessment of the stringency of these policy measures has been provided by the
218 Oxford COVID-19 Government Response Tracker (Hale et al., 2021). This daily,
219 country-level stringency index is based on a combination of containment and clo-
220 sure policies, providing an overall score of a country’s physical distancing policies
221 between 0 (no restrictions) and 100 (most severe restrictions). Components of
222 the stringency index include the severity of restrictions on gatherings, cancella-
223 tion of public events, stay-at-home requirements, school, university, workplace
224 and public transport closures, local, regional and international travel controls,
225 and the presence of public information campaigns. The stringency index purely
226 evaluates the strictness of government policies, without providing a qualitative
227 assessment on the appropriateness of restrictions.

228 *2.1.4. Mobility*

229 Google’s COVID-19 Community Mobility Reports (Google, 2020) were down-
230 loaded to assess the mobility patterns of city residents. These reports are based
231 on GPS location information of users who had turned on Location History on
232 their mobile phone. Google assigns location information to places using six
233 categories: retail and recreation, grocery and pharmacy, parks, transit stations,
234 workplaces, and residences. The total number of visitors is recorded for each cat-
235 egory, except for residences which is measured by average time spent. Reports
236 provide a daily percentage change, comparing actual mobility to levels before
237 widespread disruptions due to the pandemic. Baseline mobility levels represent
238 the median value for the corresponding category and day of week during the
239 five-week period between January 3 and February 6, 2020. The Community
240 Mobility Reports consist of county/province-level data to preserve privacy. In
241 this research, county-level data was fully allocated to a single city when a city
242 was located inside the county area (i.e., Clark County for Las Vegas), or com-
243 bined when the city extended across multiple counties (i.e., Bronx, Kings, New
244 York, Queens, and Richmond counties for New York City).

245 *2.2. Modelling*

246 As air pollution levels strongly depend on meteorological conditions, pol-
 247 lution forecasts should be adjusted for weather-related effects before analysis.
 248 In this study, the relationship between weather and air pollution was modelled
 249 based on historical data from January 2015 until December 2019. A separate
 250 model was calibrated (i.e., fitted) for each city, as city topography and the pres-
 251 ence of pollution sources in or near the city have a strong influence on the impact
 252 of changes in weather. For example, southerly winds in Melbourne (Australia)
 253 transport clean maritime air masses to the city, while the same southerly winds
 254 in continental Europe might bring in pollution from nearby cities. Further, four
 255 separate models were calibrated for pollutants NO_2 , PM_{10} , $\text{PM}_{2.5}$ and O_3 . To
 256 ensure a sufficient amount of ground-level pollution data is available for model
 257 calibration and analysis, cities were excluded if less than 365 training samples
 258 were available or more than 30 pollution measurements in 2020 were missing.
 259 This led to a sample of 720, 710, 751 and 707 cities for NO_2 , PM_{10} , $\text{PM}_{2.5}$ and
 260 O_3 , respectively. Fig. 1 provides a schematic overview of the modelling process
 261 for each city and pollutant, which is described in more detail in the following
 262 paragraphs.

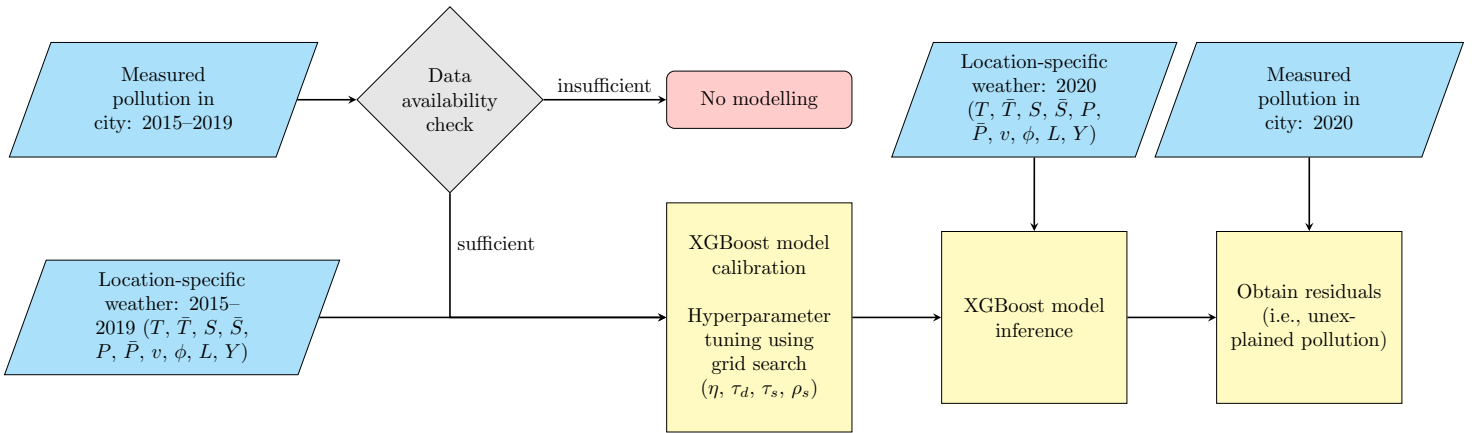


Fig. 1. Flowchart describing the modelling process that was followed for each city and pollutant.

263 City-specific model forecasts based on the actual weather in 2020 provided
264 an indication of pollution levels if the COVID-19 pandemic had not occurred.
265 Model forecasts for 2020 were then compared to observed air pollution levels
266 in 2020 to obtain the forecast errors (i.e., residuals). Residuals of the model
267 consist of (i) imperfections in the statistical model for the weather—pollution
268 relationship, (ii) the impact of pandemic-related government restrictions and
269 reduced mobility on air pollution, and (iii) other confounders such as bushfires.
270 Therefore, it is important to model the weather—pollution relationship with
271 high accuracy, reducing noise in subsequent analyses.

272 *2.2.1. XGBoost*

273 Machine learning was used to model city-specific air pollution with weather-
274 related variables. Specifically, this study used XGBoost, an eXtreme Gradi-
275 ent Boosting algorithm based on decision trees. XGBoost is an adaptation of
276 gradient boosting machines (Friedman, 2001), which perform additive optimi-
277 sation in functional space. This means that a sequence of decision trees is
278 calibrated, where each decision tree aims to explain the residuals left by the
279 previous tree. Compared to gradient boosting machines, XGBoost adds a regu-
280 larisation constraint to the objective function to prevent overfitting and includes
281 several methodological improvements to enhance scalability. Model calibration
282 was performed on the University of Melbourne’s high-performance computing
283 system (Lafayette et al., 2016), using the XGBoost implementation in R.

284 *2.2.2. Hyperparameter tuning*

285 XGBoost has several hyperparameters that influence the learning process.
286 For example, the shrinkage parameter η controls the learning rate of the algo-
287 rithm, by scaling down the contribution of each decision tree to the pollution
288 estimate. Generally, a low η leads to a more robust model by limiting the influ-
289 ence of a single decision tree, although resulting in longer calibration times (i.e.,
290 more trees are required). Further, constraints on the structure of each decision
291 tree can limit model complexity. These constraints include the maximum depth

292 of a tree (τ_d) and the minimum number of samples required in a single node of
 293 a tree (τ_s). The latter reduces complexity by preventing further splits when the
 294 data does not provide sufficient evidence (i.e., only a small number of samples
 295 is available at the node). Further, the calibration of each decision tree using
 296 a random sample of the data can improve model performance by reducing the
 297 correlation of subsequent trees. Sub-sampling can be controlled using hyper-
 298 parameters for the ratio of training samples (ρ_s) and the ratio of features (ρ_f)
 299 that is used for model calibration.

300 In this study, hyperparameters were tuned using grid search. Initial ex-
 301 periments used a large grid for a small sample of models ($n = 45$; 15 cities
 302 and 3 pollutants). The initial grid, presented in Table 1, led to a total of 720
 303 combinations of hyperparameter values. However, extensive experimentation
 304 on the high-performance computing cluster was not feasible using a grid search
 305 this large. Note that a single experiment involved calibrating 2888 XGBoost
 306 models, capturing over 700 cities and four different pollutants. Further, each of
 307 these model calibrations involved a grid search over the hyperparameter space,
 308 including a five-fold cross-validation for each possible combination of the hyper-
 309 parameters and up to 1000 decision trees per sub-model.

| Hyperparameter | Value set |
|----------------|---|
| η | { <u>0.01</u> , <u>0.05</u> , 0.10} |
| τ_d | {2, 4, <u>5</u> , <u>6</u> , <u>8</u> , 10} |
| τ_s | {1, 3, <u>5</u> , <u>7</u> } |
| ρ_s | { <u>0.5</u> , <u>0.6</u> , 0.7, 0.8, 1} |
| ρ_f | {0.8, <u>1</u> } |

Table 1. Reduced (underlined) and initial hyperparameter set for grid search.

310 Experiments using the full grid aimed to narrow down the range of hyper-
 311 parameter values that performed well on our dataset. Hyperparameter com-
 312 binations were ranked based on lowest root mean squared error (RMSE) in

313 cross-validation. The highlighted subset of hyperparameter values in Table 1
314 were frequently present in the top models of the sampled cities. Therefore, a
315 grid search over these values, a total of 24 hyperparameter combinations, was
316 used for the full-scale experiments in this study. In particular, a low learning
317 rate ($\eta \in \{0.01, 0.05\}$) reduced overfitting by making the boosting process more
318 conservative. Further, the minimum number of instances required in each node
319 was increased from the default setting (i.e., 1) to improve generalisability ($\tau_s \in$
320 $\{5, 7\}$). Sub-sampling was applied to samples, but not to features. That is,
321 either 50% or 60% of observations was selected at random to calibrate each de-
322 cision tree, while a tree could use all weather-related variables. Besides limiting
323 overfitting, sub-sampling also reduced computation time for a single tree, aiding
324 the processing of many cities. Strongly restricting the maximum depth of each
325 tree did not lead to good overall accuracy. Hence, values for this hyperparameter
326 were set to $d \in \{5, 6, 8\}$.

327 *2.2.3. Feature selection*

328 Various feature selection experiments explored the best approach to incor-
329 porate the meteorological variables in the XGBoost models, based on the mean
330 absolute error (MAE) of out-of-sample forecasts for the first two weeks of Jan-
331 uary 2020. This indicated that some variables have increased prediction power
332 at different time lags. It was explored whether to include the value on the day
333 itself, day $t-1$, $t-2$, $t-3$, or the average over the last or preceding three days ($t-2$
334 to t , or $t-3$ to $t-1$), taking into account interactions with the other features.
335 Adding additional features capturing weather during the preceding three days
336 improved model accuracy the most. Specifically, the total precipitation, mean
337 temperature and mean solar radiation over the preceding three days were added
338 as features. Wind speed and direction were most predictive when measured on
339 the day itself. Since the leaf area index did not fluctuate substantially from
340 day to day, only the measurement on day t was used. Finally, a time variable
341 was included to capture any annual trends in air pollution levels, as government
342 policies sometimes resulted in a negative trend over the 2015–2019 period (e.g.,

343 cleaner fuels, gradual introduction of electric vehicles, etc.). An overview of the
 344 final selection of variables is provided in Table 2.

| Variable | Description |
|-----------|---|
| T | Air temperature at 2 meters altitude, day t |
| \bar{T} | Air temperature at 2 meters altitude, mean($t-3, t-2, t-1$) |
| S | Net solar radiation at the surface, day t |
| \bar{S} | Net solar radiation at the surface, mean($t-3, t-2, t-1$) |
| P | Total precipitation, day t |
| \bar{P} | Total precipitation, sum($t-3, t-2, t-1$) |
| v | Wind speed, day t |
| ϕ | Wind direction, day t |
| L | Leaf area index of vegetation, day t |
| Y | Year of observation |

Table 2. Selected features for XGBoost models.

345 Some variables that have potential to further improve pollution forecasts
 346 have not been included in the modelling on purpose. For example, a 'day of
 347 week' variable can explain weekly patterns in air pollution caused by mobility
 348 fluctuations (i.e., limited travel during weekends). However, since mobility will
 349 be explored later on, models did not incorporate any mobility-related variables
 350 even if this could have led to higher accuracy. The only features used in the
 351 XGBoost models were atmospheric, vegetation-related and temporal variables,
 352 capturing weather, seasonality and any annual trends in air pollution levels.

353 *2.2.4. Model forecasts and assessment*

354 Final XGBoost models were calibrated using the full training set, containing
 355 the features in Table 2. Models used the city- and pollutant-specific hyperpa-
 356 rameter set that resulted in the lowest RMSE in cross-validation. To improve
 357 explainability, variable importance in the final models was assessed through
 358 ranking. Specifically, tree-based models such as XGBoost allow for an investi-

359 gation of the contribution of each feature to the model (i.e., gain). The gain
360 quantifies how important a feature is in making a branch of a decision tree more
361 pure, based on the sum of improvements in squared error over all internal nodes
362 where the feature was chosen as a splitting variable (Hastie et al., 2009, p. 367–
363 368). The averaged feature importance characteristics over all decision trees in
364 one XGBoost model provide an overall score of feature importance. Feature
365 importance was explored in various case studies to provide further insights into
366 the modelling approach.

367 Each calibrated model was used to predict air pollution levels for the corre-
368 sponding city and pollutant throughout 2020. These weather-based forecasts for
369 2020 do not incorporate any pandemic-related reductions in emissions. Hence,
370 forecasts provide a counterfactual indication of pollution levels had a pandemic
371 not occurred (i.e., the XGBoost models are agnostic of the pandemic). The
372 differences between 2020 forecasts and the actual pollution measurements in
373 2020 are referred to as 'unexplained pollution', representing the meteorology-
374 normalised reduction in air pollution. Importantly, a separate XGBoost model
375 was calibrated for each city, independently. Therefore, city-specific air pollution
376 models provide a multi-model ensemble forecast for the unexplained pollution in
377 a country (i.e., for air pollution in large cities with a population over 300,000).

378 Finally, the unexplained NO₂ pollution was selected to explore associations
379 with mobility patterns using the Google Mobility Reports. Motor-vehicle traf-
380 fic is a major source of NO₂ pollution, while many emission sources/processes
381 contribute to PM₁₀, PM_{2.5} and O₃ levels. Since the NO₂ pollution has been ad-
382 justed for weather influences, seasonality and annual trends, unexplained pollu-
383 tion estimates could theoretically be allocated to mobility levels on a day-to-day
384 basis. However, the Google Mobility Reports used a different baseline for each
385 day of the week, complicating the analysis of mobility data for consecutive days.
386 Therefore, 7-day smoothing was applied to both time series for a more robust
387 comparison. The Google Mobility Reports stratify mobility reductions based
388 on trip purpose. A correlation analysis was performed to determine which type
389 of trips were most highly associated with unexplained NO₂ levels. Correlations

390 were computed between 15 February and 15 April 2020, capturing more regular
391 mobility levels during the pre-lockdown period and subsequent reductions dur-
392 ing the lockdown period. Analysis was performed at a country-level, showing
393 how these relationships varied, globally.

394 **3. Results**

395 *3.1. Micro level*

396 *3.1.1. Model fit*

397 In total, 2888 XGBoost models for the weather—pollution relationship were
398 calibrated, corresponding to 720, 710, 751 and 707 cities for NO₂, PM₁₀, PM_{2.5}
399 and O₃, respectively. The average explained variation in training data, as mea-
400 sured by R^2 , was 86.3%, 85.8%, 86.2% and 86.3% for each of these four pol-
401 lutants. A wide variety of seasonal patterns were observed for different cities
402 and pollutants, which were adequately captured by the XGBoost models. For
403 example, Fig. 2 provides out-of-sample XGBoost model forecasts for 2020, show-
404 ing NO₂ in Chongqing, PM₁₀ in Antwerpen and PM_{2.5} in Ulaanbaatar. The
405 assessment of out-of-sample performance using 2020 measurements is compli-
406 cated, because of the influences of the pandemic. However, results of further
407 statistical model validation have been provided in Section 3.3.

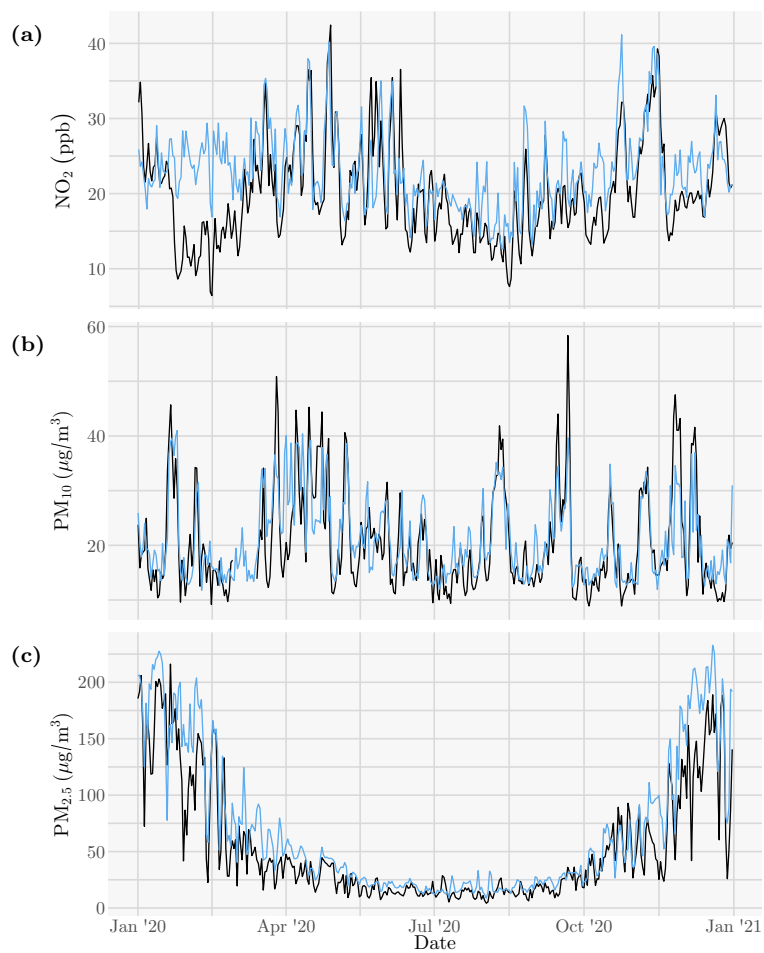


Fig. 2. Actual (black) and forecasted (blue) pollution levels for (a) NO_2 in Chongqing, China; (b) PM_{10} in Antwerpen, Belgium; and (c) $\text{PM}_{2.5}$ in Ulaanbaatar, Mongolia.

408 *3.1.2. Feature importance*

409 Tables 3 and 4 show the feature importance statistics per pollutant, com-
410 puted across all calibrated XGBoost models. For each feature in Table 3, the
411 corresponding gains of all 2888 models were averaged to obtain the mean gain.
412 Note that the gain is an indicator of relative feature importance (i.e., how im-
413 portant is a feature compared to the other features) and the gains across all
414 features sum to 1. The three-day temperature average and the leaf area index
415 were the most useful features for NO_2 , PM_{10} and $\text{PM}_{2.5}$, while the current solar
416 radiation was most important for O_3 . Variations in temperature and the amount
417 of foliage throughout the year are both indicators of seasonal influences that af-
418 fect air pollution levels. Urban vegetation also impacts pollution levels directly
419 by removing a substantial amount of NO_2 and particulate matter from the air
420 (Nowak et al., 2006). In contrast, temperature levels have a significant effect
421 on the emission of pollutants into the atmosphere. For example, the increased
422 pollution levels in Ulaanbaatar, Mongolia (see Fig. 2c) were mainly caused by
423 the use of heating stoves and heat only boilers during winter (The World Bank,
424 2009). The XGBoost models were in line with these observations, as \bar{T} was the
425 most important feature to predict air pollution in Ulaanbaatar (i.e., \bar{T} gain =
426 0.695 for $\text{PM}_{2.5}$). High temperatures can also result in increased pollution levels
427 through atmospheric chemistry (von Schneidemesser et al., 2015) or increased
428 energy consumption such as through the use of air conditioning (Davis and
429 Gertler, 2015). For example, Fig. 2b shows increased PM_{10} levels during the
430 early August 2020 heat wave in Belgium, captured accurately by the XGBoost
431 model.

432 The top features per pollutant were not always the best-performing features,
433 as considerable variation in feature importance was observed across cities (see
434 Table 4). For example, wind direction was the most important feature for sev-
435 eral coastal cities (e.g., Liverpool, Nantong, Jakarta). In these cities, clean
436 ocean air can substantially reduce pollution levels, while unfavourable wind
437 conditions transport pollution from other cities or nearby power plants into the

| Feature | Mean gain | NO ₂ | PM ₁₀ | PM _{2.5} | O ₃ |
|-----------|-----------|-----------------|------------------|-------------------|----------------|
| L | 0.134 | 0.137 | 0.126 | 0.142 | 0.131 |
| \bar{T} | 0.127 | <u>0.152</u> | <u>0.136</u> | <u>0.153</u> | 0.064 |
| T | 0.110 | 0.105 | 0.094 | 0.108 | 0.133 |
| S | 0.107 | 0.063 | 0.076 | 0.076 | <u>0.218</u> |
| v | 0.105 | 0.129 | 0.108 | 0.109 | 0.071 |
| \bar{P} | 0.092 | 0.075 | 0.118 | 0.122 | 0.052 |
| \bar{S} | 0.088 | 0.076 | 0.071 | 0.076 | 0.131 |
| P | 0.086 | 0.095 | 0.111 | 0.062 | 0.076 |
| ϕ | 0.082 | 0.081 | 0.095 | 0.088 | 0.065 |
| Y | 0.069 | 0.087 | 0.066 | 0.064 | 0.059 |

Table 3. Average feature importance of temperature (T , \bar{T}), leaf area index (L), precipitation (P , \bar{P}), wind speed (v) and direction (ϕ), solar radiation (S , \bar{S}) and year (Y). The most important feature per pollutant is underlined.

city. This assertion is supported by air pollution, wind speed and wind direction observations during 2015–2019. Bivariate polar plots of these observations (created using the polarplotr/openair R package by Carslaw and Ropkins, 2012) illustrate how air pollution levels in these cities varied based on wind direction and speed. In Liverpool (Fig. 3a), wind directions between East-northeast (ENE) and South-southeast (SSE) transport particulate matter from nearby cities Manchester, Birmingham and London, and continental Europe into the city (Graham et al., 2020). Further, wind directions from the ocean transport clean maritime air masses, substantially reducing air pollution in Liverpool and Nantong (Figs. 3a–b).

Wind speed was a slightly more informative feature in NO₂ models than in PM₁₀, PM_{2.5} and O₃ models. In contrast, the total precipitation over the past three days (\bar{P}) was more useful for predicting PM₁₀ and PM_{2.5} than NO₂ and O₃. This finding is consistent with the larger washout effect of precipitation for particulate matter compared to NO₂ (Yoo et al., 2014). Due to many null

| Feature | Overall | NO ₂ | PM ₁₀ | PM _{2.5} | O ₃ |
|-----------|---------|-----------------|------------------|-------------------|----------------|
| L | 631 | 160 | 146 | <u>203</u> | 122 |
| \bar{T} | 549 | <u>169</u> | <u>182</u> | 188 | 10 |
| T | 329 | 101 | 46 | 76 | 106 |
| S | 361 | 1 | 25 | 8 | <u>327</u> |
| v | 261 | 109 | 52 | 91 | 9 |
| \bar{P} | 193 | 18 | 85 | 85 | 5 |
| \bar{S} | 124 | 31 | 9 | 21 | 63 |
| P | 182 | 58 | 87 | 12 | 25 |
| ϕ | 125 | 14 | 60 | 46 | 5 |
| Y | 133 | 59 | 18 | 21 | 35 |

Table 4. As per Table 3, but showing the number of cities where the feature was the most important feature.

453 observations, \bar{P} and P are not a differentiating feature in cities with a mostly
 454 dry climate (e.g., $\{\bar{P}, P\}$ gain = $\{0.023, 0.019\}$ and $\{0.024, 0.010\}$ for PM₁₀
 455 and PM_{2.5} in the semi-arid climate of Aguascalientes, Mexico). However, the
 456 precipitation variables rank highly for some cities with a sufficient amount of
 457 rainfall. For example, \bar{P} was the most influential predictor for PM_{2.5} in Kath-
 458 mandu (Nepal), which receives a substantial amount of rain annually while the
 459 local topography limits the ability of wind to clear the area of pollution (see
 460 Fig. 3c). Overall, the results of the feature importance analysis support our
 461 approach to model air pollution separately for each city and pollutant. This
 462 allowed the XGBoost models to adapt to the distinct urban features, climate
 463 and atmospheric conditions for each city, individually.

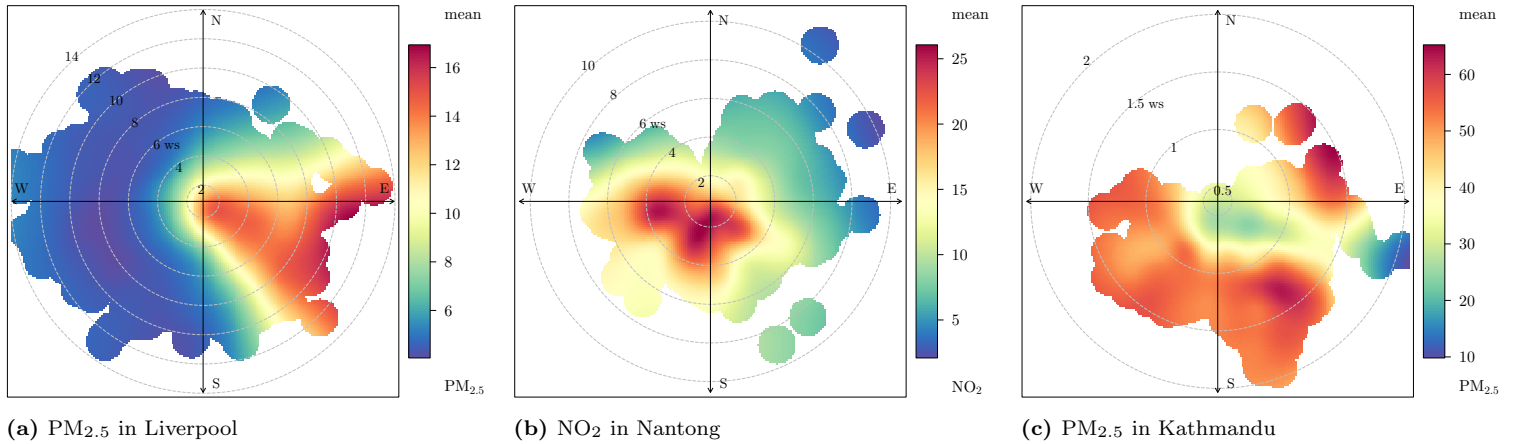


Fig. 3. Bivariate polar plots of pollution compared to wind direction and speed, as observed during 2015–2019.

464 *3.2. Macro level*

465 After exploring micro-level effects in Sections 3.1.1 and 3.1.2, this section
 466 will focus on macro-level effects. Figs. 4–6 show the results of the ensemble
 467 forecasts for the unexplained pollution in several countries, using the XGBoost
 468 models of all corresponding cities. The figures present the absolute reduction in
 469 NO_2 , $\text{PM}_{2.5}$ and O_3 , stated in ppb, $\mu\text{g}/\text{m}^3$ and ppb, respectively. For brevity,
 470 PM_{10} has not been included, as country-level results showed similar patterns
 471 for unexplained $\text{PM}_{2.5}$ and PM_{10} (even though $\text{PM}_{2.5}$ and PM_{10} models were
 472 calibrated independently). The unexplained pollution is the difference between
 473 2020 measurements and the forecasts of the XGBoost models using weather,
 474 vegetation and temporal variables. Hence, these outcomes are agnostic of the
 475 pandemic and the associated restrictions on mobility. However, the unexplained
 476 pollution matched the timing of the restrictions in different countries very well.
 477 In particular, the unexplained pollution in China peaked between late January
 478 and early February, consistent with the lockdown during this period (see first
 479 column of Fig. 4). In other countries, the peaks were observed around early
 480 April. In larger countries such as China and India, regional lockdowns led to

481 larger confidence intervals, as restriction levels varied across cities (He et al.,
482 2020). In smaller countries restrictions generally applied more uniformly, re-
483 sulting in narrower confidence intervals. As the XGBoost models did not in-
484 corporate mobility-related variables, many of the unexplained pollution charts
485 still showed 7-day temporal patterns. This is especially visible for NO₂ in the
486 United States and several European cities, showing 52 peaks of reduced pollu-
487 tion corresponding to weekends in 2020. The same patterns can be observed for
488 the mobility time series, for example, capturing different effects of the pandemic
489 on workplace travel during weekdays and weekends.

490 Besides differences in timing, the amplitude of the pollution reductions also
491 varied substantially between countries. For NO₂, the largest reductions were ob-
492 served in China, India and Europe. Further, the amplitude of NO₂ reductions
493 was in line with the magnitude of observed reductions in mobility for the cor-
494 responding country. Large reductions in PM_{2.5} were mainly observed for China
495 and India. Note that in absolute terms only slight reductions in PM_{2.5} were
496 apparent for some countries. However, these reductions were still substantial in
497 relative terms, as the regular ranges of PM_{2.5} levels were already significantly
498 lower than in China and India to start with. Since negative health effects are
499 computed based on absolute levels, less benefits may be obtained in these coun-
500 tries by interventions targeting reductions in mobility, or alternative transport
501 initiatives.

502 Most existing studies have investigated air pollution during the initial lock-
503 down periods in March and April 2020. Our study found that after the initial
504 lockdowns, NO₂ and PM_{2.5} emissions did not fully return to pre-pandemic lev-
505 els. Reduced mobility during most of 2020 led to a sustained reduction in air
506 pollution, although not to the same extent as during the initial lockdowns. For
507 O₃, many studies reported an increase during the initial lockdowns (see Sec-
508 tion 1.1.3), consistent with our results. However, from June/July to September,
509 our analyses found a subsequent reduction in O₃ for many countries below what
510 was expected based on meteorological conditions (e.g., China, United Kingdom,
511 France, Germany, Poland, Turkey). This period corresponds to summer in the

512 Northern Hemisphere, when O₃ levels are normally elevated. Note that photochemical production of O₃ at mid and high latitudes in the Northern Hemisphere is low in winter due to reduced sunlight and temperatures (Dentener et al., 2020). Most studies reported increases in O₃ during March and April, from already low levels. However, the implications in periods with high O₃ levels are especially important with respect to health consequences. This should be investigated further in future research.

519 Figs. 4–6 also illustrate the effects of different baselines for each day of the week in the mobility dataset. This resulted in positive peaks during weekends, corresponding to lower mobility reductions with respect to already low baseline levels (e.g., see workplace travel). In contrast, the unexplained NO₂ pollution estimates showed negative peaks during weekends. Table 5 reports correlations after 7-day smoothing between the NO₂ and mobility time series from mid-February to mid-April, linking mobility patterns for each country to the unexplained pollution levels. Correlations were reported for all countries where at least one city was modelled, noting that one city with complete measurements resulted in $n = 61$ data points. For various countries, the total unexplained NO₂ pollution was highly correlated with changes in mobility levels. Exposures for each trip type were different per country (e.g., travel to workplaces may constitute the majority of trips), indicating that policy interventions in some categories may be more promising to explore than others. Trips to transit stations, workplaces, retail and recreation venues generally had a higher correlation with unexplained pollution than trips to groceries and pharmacies. Further, park visits varied significantly between countries. In most of the Nordic countries (Denmark, Finland, Norway) the correlations between park usage and unexplained pollution were opposite to the rest of the world, presumably because the baseline levels for park visits during winter were very low due to the cold weather conditions in these countries.

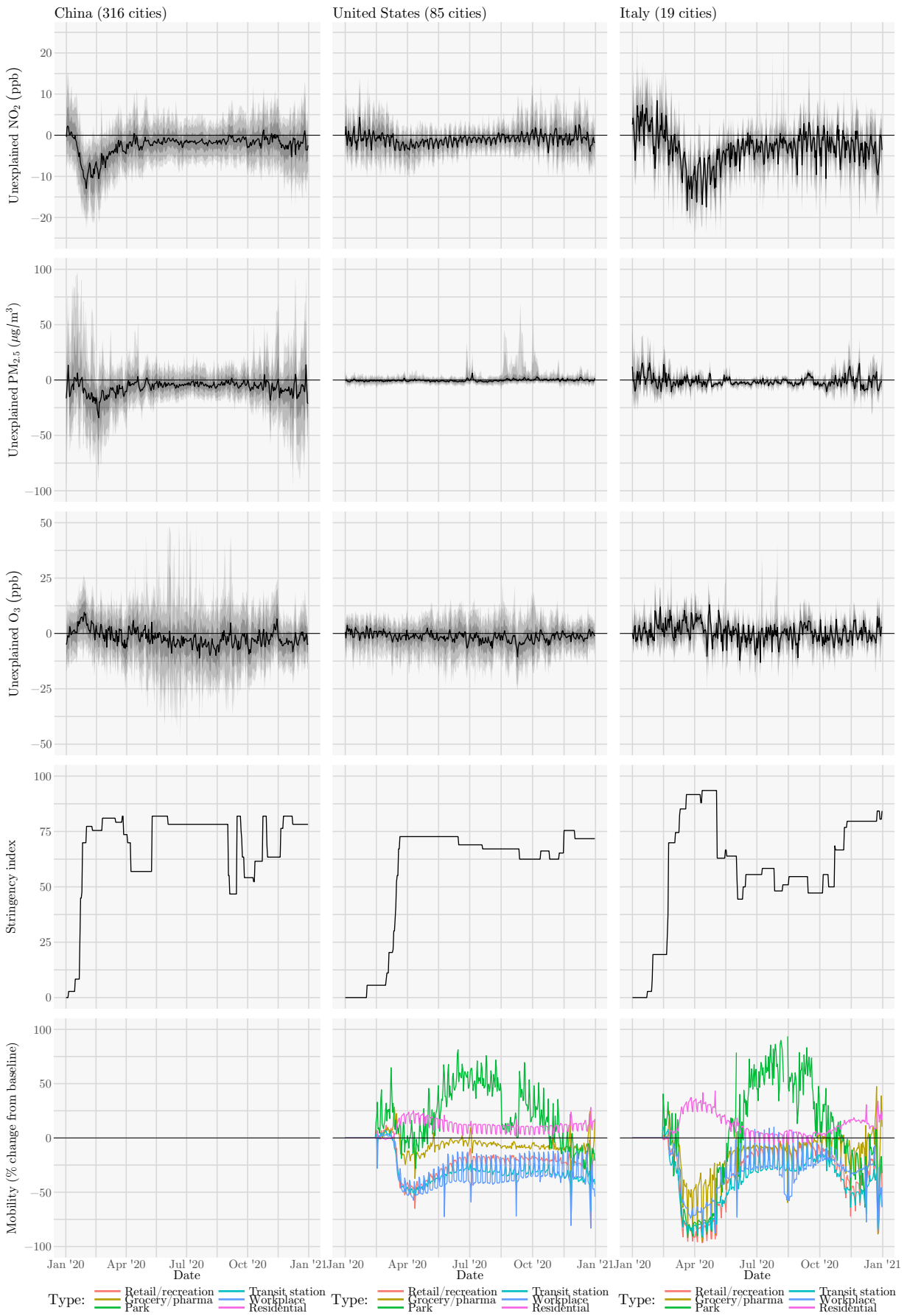


Fig. 4. Unexplained NO_2 (ppb) and PM_{10} ($\mu\text{g}/\text{m}^3$) across cities, stringency of COVID-19 restrictions, and mobility patterns for China, United States and Italy. Shading indicates 50, 80 and 90% confidence intervals. Google mobility data was not available for China; US park visits between mid-August and September had data quality issues.

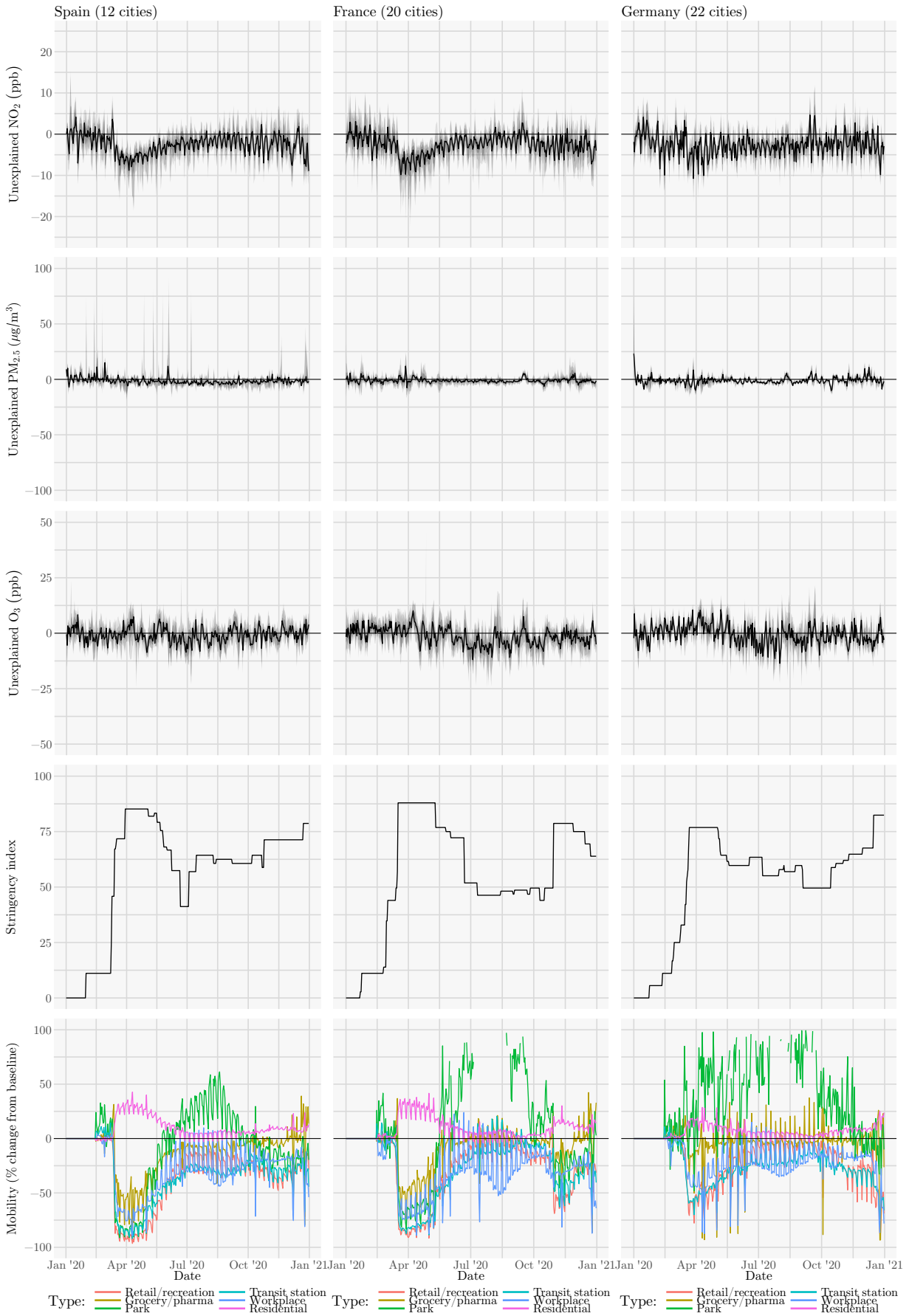


Fig. 5. As Fig. 4, but for Spain, France and Germany. Park usage in France and Germany exceeds +100% during summer months, compared to the January baseline.

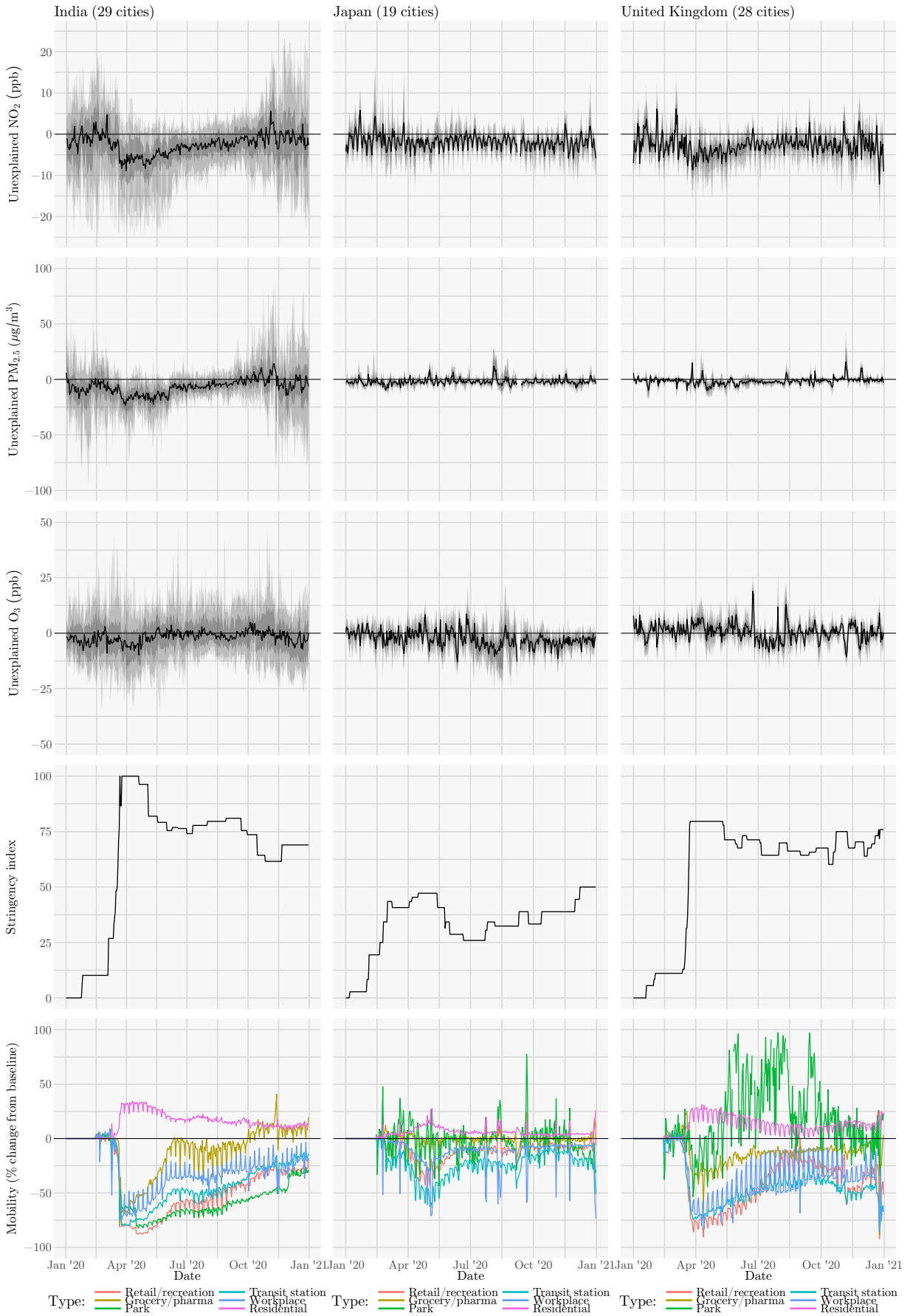


Fig. 6. As Fig. 4, but for India, Japan and Mexico.

| Country | Maximum stringency | n | Transit stations | Workplaces | Retail and recreation | Grocery and pharmacy | Parks | Residential |
|------------------------|--------------------|------|------------------|------------|-----------------------|----------------------|-------|-------------|
| Croatia | 96.3 | 46 | 0.98 | 0.97 | 0.98 | 0.94 | 0.91 | -0.97 |
| Colombia | 88.0 | 61 | 0.97 | 0.97 | 0.97 | 0.94 | 0.97 | -0.97 |
| Serbia | 100.0 | 61 | 0.96 | 0.96 | 0.95 | 0.94 | 0.92 | -0.96 |
| Denmark | 72.2 | 61 | 0.92 | 0.91 | 0.92 | 0.84 | -0.60 | -0.92 |
| Bosnia and Herzegovina | 92.6 | 61 | 0.90 | 0.89 | 0.88 | 0.84 | 0.83 | -0.90 |
| Israel | 94.4 | 177 | 0.86 | 0.82 | 0.82 | 0.80 | 0.75 | -0.84 |
| Finland | 67.6 | 122 | 0.80 | 0.78 | 0.76 | 0.80 | -0.48 | -0.81 |
| Bolivia | 84.3 | 36 | 0.78 | 0.78 | 0.79 | 0.78 | 0.77 | -0.78 |
| Indonesia | 71.8 | 61 | 0.72 | 0.79 | 0.78 | 0.74 | 0.72 | -0.81 |
| Spain | 85.2 | 693 | 0.76 | 0.77 | 0.75 | 0.73 | 0.74 | -0.80 |
| Romania | 87.0 | 183 | 0.77 | 0.76 | 0.76 | 0.74 | 0.75 | -0.77 |
| Hungary | 76.9 | 61 | 0.73 | 0.75 | 0.74 | 0.84 | 0.77 | -0.75 |
| Belgium | 81.5 | 60 | 0.76 | 0.73 | 0.76 | 0.74 | 0.57 | -0.75 |
| Austria | 81.5 | 61 | 0.73 | 0.72 | 0.72 | 0.69 | 0.73 | -0.73 |
| Portugal | 88.0 | 122 | 0.68 | 0.71 | 0.68 | 0.74 | 0.67 | -0.70 |
| France | 88.0 | 1213 | 0.70 | 0.72 | 0.69 | 0.68 | 0.69 | -0.72 |
| New Zealand | 96.3 | 122 | 0.69 | 0.69 | 0.72 | 0.74 | 0.66 | -0.66 |
| Switzerland | 73.2 | 181 | 0.73 | 0.64 | 0.71 | 0.66 | 0.21 | -0.68 |
| Italy | 93.5 | 1085 | 0.66 | 0.70 | 0.67 | 0.69 | 0.66 | -0.71 |
| Czech Republic | 82.4 | 94 | 0.59 | 0.60 | 0.55 | 0.64 | 0.25 | -0.58 |
| China | 81.9 | 427 | 0.45 | 0.59 | 0.51 | 0.70 | 0.29 | -0.69 |
| Brazil | 74.5 | 329 | 0.61 | 0.62 | 0.57 | 0.51 | 0.63 | -0.62 |
| Germany | 76.9 | 1303 | 0.54 | 0.57 | 0.58 | 0.65 | -0.06 | -0.57 |
| Ireland | 90.7 | 28 | 0.56 | 0.58 | 0.59 | 0.59 | 0.57 | -0.57 |
| Estonia | 77.8 | 61 | 0.61 | 0.52 | 0.59 | 0.57 | 0.49 | -0.59 |
| South Africa | 88.0 | 271 | 0.57 | 0.58 | 0.55 | 0.51 | 0.49 | -0.58 |
| Vietnam | 96.3 | 108 | 0.66 | 0.46 | 0.57 | 0.67 | 0.79 | -0.42 |
| United Kingdom | 79.6 | 1387 | 0.57 | 0.58 | 0.55 | 0.48 | 0.37 | -0.59 |
| Norway | 79.6 | 61 | 0.56 | 0.57 | 0.57 | 0.48 | -0.34 | -0.55 |
| Poland | 83.3 | 488 | 0.50 | 0.56 | 0.49 | 0.54 | 0.49 | -0.52 |
| Bulgaria | 73.2 | 105 | 0.46 | 0.41 | 0.41 | 0.28 | 0.31 | -0.40 |
| Mongolia | 65.7 | 61 | 0.42 | 0.23 | 0.46 | 0.32 | 0.20 | -0.35 |
| United States | 72.7 | 3961 | 0.34 | 0.36 | 0.37 | 0.20 | 0.02 | -0.37 |
| India | 100.0 | 1594 | 0.30 | 0.32 | 0.30 | 0.32 | 0.29 | -0.33 |
| Argentina | 100.0 | 94 | 0.32 | 0.28 | 0.30 | 0.35 | 0.30 | -0.30 |
| Chile | 73.2 | 183 | 0.34 | 0.34 | 0.27 | 0.17 | 0.15 | -0.37 |
| Macedonia | | 61 | 0.25 | 0.30 | 0.25 | 0.31 | 0.11 | -0.24 |
| Japan | 45.4 | 1038 | 0.41 | 0.29 | 0.34 | -0.06 | -0.06 | -0.35 |
| Canada | 74.5 | 755 | 0.29 | 0.25 | 0.27 | 0.19 | 0.22 | -0.27 |
| Netherlands | 79.6 | 305 | 0.22 | 0.24 | 0.30 | 0.19 | 0.14 | -0.26 |
| Mexico | 82.4 | 781 | 0.16 | 0.16 | 0.18 | 0.27 | 0.21 | -0.14 |
| Thailand | 76.9 | 549 | 0.16 | 0.16 | 0.17 | 0.15 | 0.17 | -0.17 |
| Turkey | 77.8 | 774 | 0.22 | 0.12 | 0.10 | 0.07 | 0.15 | -0.10 |
| Slovakia | 87.0 | 61 | 0.09 | 0.11 | 0.11 | 0.15 | 0.30 | -0.10 |
| Sweden | 64.8 | 122 | 0.08 | 0.06 | 0.02 | 0.03 | 0.04 | -0.06 |
| South Korea | 82.4 | 1464 | -0.05 | 0.21 | -0.03 | -0.42 | -0.42 | -0.07 |

Table 5. Correlations between mobility and unexplained NO₂, computed per country from mid-February to mid-April.

540 3.3. Robustness of results

541 To assess the impact of modelling assumptions used in this research, various
542 alternative approaches were explored. Specifically, the weather—pollution rela-
543 tionship was modelled using linear regression, generalised linear models (GLM)
544 using various link functions, and random forests. The goodness-of-fit on train-
545 ing data, averaged across all cities and pollutants, varied from linear regression
546 ($R^2 = 36.7\%$), Gamma GLM with inverse link function ($R^2 = 40.6\%$), random
547 forest ($R^2 = 50.3\%$), to XGBoost ($R^2 = 86.2\%$). As the goodness-of-fit of each
548 model was different, Table 6 shows the variation of model fit across cities. Fur-
549 ther, a time series modelling approach without weather variables was explored.
550 As the time series forecasting approach differed substantially from our preferred
551 XGBoost method described in Section 2.2, it is described here in more detail.
552 First, one time series was created per pollutant and city based on pollution mea-
553 surements during 2015–2019. Missing values were imputed, or the time series
554 was disregarded if there were too many missing values. Each remaining time
555 series was decomposed into seasonal, trend and remainder components using
556 Loess (STL) (Cleveland et al., 1990). The trend and remainder components
557 were then recombined and used to calibrate an ARIMA model. A 2020 forecast
558 was computed as the prediction of the ARIMA model plus the seasonal compo-
559 nent obtained earlier. This resulted in city- and pollutant-specific 2020 forecasts
560 based purely on historical air pollution data, without adjustments for weather
561 variables. Since pollution data until the end of 2019 was used for model cali-
562 bration, the time series forecasting approach is also agnostic to the COVID-19
563 pandemic.

564 Machine learning methods such as XGBoost tend to overfit on training data.
565 To provide an indication of out-of-sample performance, Fig. 7 provides examples
566 of the actual pollution levels in 2020 and predictions using the various modelling
567 approaches for NO_2 in Ürümqi (China) and Santiago (Chile), and PM_{10} in Pune
568 (India). All methods correctly captured the seasonal pollution patterns of the
569 different cities. Notably, Ürümqi experienced multiple COVID-related lock-
570 downs. Besides the initial lockdown in China during February and early March,

| Percentile | Linear model | Gamma GLM | Random forest | XGBoost |
|------------------|--------------|-----------|---------------|---------|
| 10 th | 18.4% | 23.4% | 31.1% | 74.8% |
| 20 th | 25.3% | 29.6% | 38.7% | 80.9% |
| 30 th | 29.8% | 34.0% | 43.6% | 84.4% |
| 40 th | 33.6% | 38.0% | 47.5% | 86.7% |
| 50 th | 36.6% | 41.2% | 51.2% | 88.5% |
| 60 th | 40.1% | 44.4% | 54.5% | 90.0% |
| 70 th | 43.6% | 47.8% | 58.5% | 91.5% |
| 80 th | 47.9% | 51.6% | 62.7% | 92.9% |
| 90 th | 54.4% | 56.3% | 68.2% | 94.7% |

Table 6. Variation of model fit across cities, measured using R^2 .

571 residents of Ürümqi were locked inside their homes from 18 July to 31 August
572 as part of very severe lockdown restrictions. Based on the weather during these
573 periods, all methods forecasted substantially higher levels of NO₂ than mea-
574 sured. In Santiago (Chile), NO₂ pollution normally increases during winter in
575 the Southern Hemisphere. Chile experienced a severe outbreak of COVID-19
576 during 2020 and residents in Santiago were subject to an extended 143-day lock-
577 down period from March to mid-August. The various methods were consistent
578 in predicting substantially higher NO₂ levels for the counterfactual scenario,
579 based on the observed weather. Finally, the large seasonal PM₁₀ fluctuations in
580 Pune (India) were captured accurately, including pollution spikes in November
581 2020, although forecasts also showed some variations between methods.

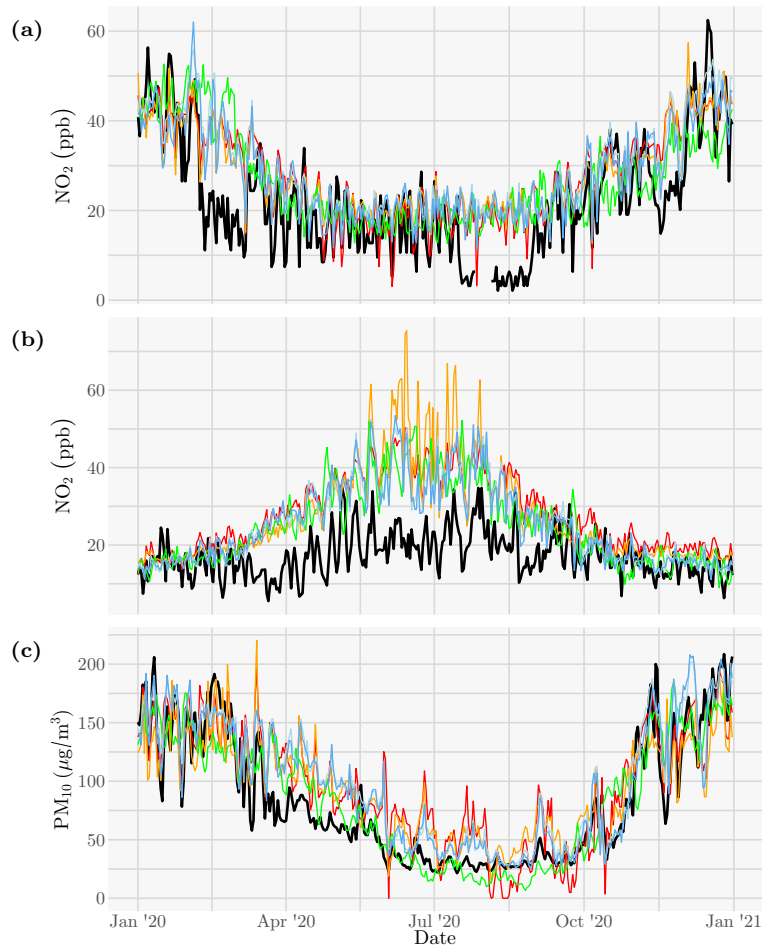


Fig. 7. Actual (black), XGBoost (blue), linear regression (red), Gamma GLM (orange), random forest (light blue) and time series (green) forecasts for (a) NO_2 in Ürümqi, China; (b) NO_2 in Santiago, Chile; and (c) PM_{10} in Pune, India.

582 Tables 7 and 8 provide a statistical assessment of out-of-sample performance.
583 The accuracy of pollution forecasts for the first two weeks of January 2020
584 was computed based on MAE and RMSE across all cities. The effects of the
585 pandemic on mobility levels were not as widespread in early January 2020.
586 Hence, forecasting errors during this period can, to a large extent, be attributed
587 to the quality of the statistical models, except for instances where increased
588 pollution was not related to weather effects (e.g., the large-scale bushfires in
589 Australia in January 2020, or reduced travel during the holiday period). Results
590 show that the XGBoost models obtained the best out-of-sample performance,
591 except for PM_{2.5} and O₃, where the random forest approach led to a slightly
592 lower MAE and RMSE.

| | Time series | Linear model | Gamma GLM | Random forest | XGBoost |
|--|-------------|--------------|-----------|---------------|---------|
| NO ₂ (ppb) | 5.45 | 5.01 | 5.19 | 4.40 | 4.18 |
| PM ₁₀ ($\mu\text{g}/\text{m}^3$) | 28.02 | 21.76 | 21.78 | 21.78 | 21.09 |
| PM _{2.5} ($\mu\text{g}/\text{m}^3$) | 20.01 | 16.69 | 17.96 | 16.58 | 16.68 |
| O ₃ (ppb) | 6.68 | 7.52 | 8.20 | 6.05 | 6.28 |

Table 7. MAE of forecasts during the first two weeks of January 2020, for different modelling approaches.

| | Time series | Linear model | Gamma GLM | Random forest | XGBoost |
|--|-------------|--------------|-----------|---------------|---------|
| NO ₂ (ppb) | 6.54 | 5.95 | 6.47 | 5.37 | 5.13 |
| PM ₁₀ ($\mu\text{g}/\text{m}^3$) | 34.07 | 26.81 | 27.89 | 26.50 | 26.18 |
| PM _{2.5} ($\mu\text{g}/\text{m}^3$) | 24.44 | 21.20 | 24.09 | 20.65 | 21.02 |
| O ₃ (ppb) | 8.08 | 8.93 | 9.72 | 7.37 | 7.60 |

Table 8. RMSE of forecasts during the first two weeks of January 2020, for different modelling approaches.

593 Importantly, when aggregating results to country level, the alternative ap-
594 proaches led to similar results, but with an increased level of noise (see Fig. 8).

595 The XGBoost method presented in this study had the highest prediction accu-
 596 racy (based on training and out-of-sample data) and hence, provides the most
 597 accurate assessment of the impact of the pandemic on pollution levels. It also
 598 led to smaller confidence intervals compared to a time series or linear modelling
 599 approach. Replication of the main results using alternative methodologies pro-
 600 vides further confidence in the robustness of the presented results. Deep learning
 601 methods have not been investigated in this study, but could potentially further
 602 improve accuracy.

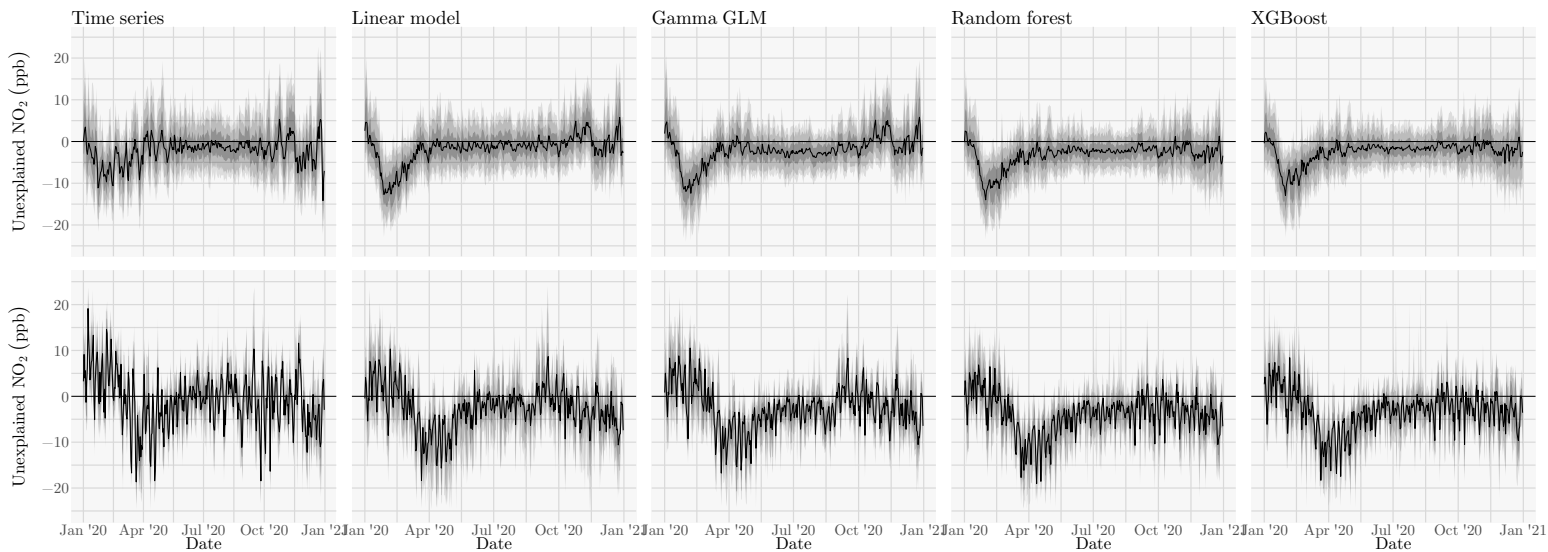


Fig. 8. Sensitivity analysis, showing unexplained NO_2 (ppb) across cities in China (top row) and Italy (bottom row) for different modelling approaches. Shading indicates 50, 80 and 90% confidence intervals.

603 Some XGBoost models had a lower goodness-of-fit on training data than the
 604 86% average R^2 , for example, due to remaining data quality issues or external
 605 events impacting pollution levels during 2015–2019. To assess the sensitivity
 606 of the results to model quality, various experiments were performed with dif-
 607 ferent exclusion criteria for XGBoost models with a reduced quality of fit. For
 608 example, experiments were conducted that excluded XGBoost models with an
 609 R^2 on training data below 50%. This resulted in the exclusion of only 21 out

610 of 2888 models (i.e., 0.7%) and had limited effects on the overall results. Fi-
611 nally, various experiments were performed using air pollution measurements
612 obtained via satellite remote sensing, as opposed to ground-level measurements.
613 Ground-level measurements were preferred, leading to improved correlations of
614 the unexplained pollution with the mobility time series.

615 **4. Discussion**

616 The attribution of changes in air pollution levels for primary and secondary
617 pollutants due to COVID-19 is complex. For example, the impact on primary
618 pollutants is affected by the source profile of emissions applicable in a spe-
619 cific country. Government restrictions during the pandemic generally impacted
620 transport emissions heavily and industrial activity to a lesser extent. In con-
621 trast, other sources such as agriculture were largely unaffected, while residen-
622 tial emissions mostly increased due to stay-at-home orders. Attribution is fur-
623 ther complicated by government policies in response to climate change, extreme
624 events such as large-scale bushfires, and weather conditions impacting atmo-
625 spheric chemistry, emission levels, and the in-flow of particulate matter from
626 nearby countries. Our study attempted to disentangle some of these confound-
627 ing factors to directly compare absolute reductions in pollutions levels due to
628 the COVID-19 pandemic, globally.

629 *4.1. Consistency with other publications*

630 Daily meteorology-normalised pollution estimates were obtained for a wide
631 variety of countries. As the temporal and spatial coverage is a superset of
632 most other studies, results can easily be compared to those reported elsewhere.
633 Table 9 provides a comparison to studies that investigated a specific country,
634 region, or city. Our estimates are a slice of the full results for the same lo-
635 cation/region and time period as reported by the respective study. Note that
636 differences are expected due to variations in methodology, confounding factors
637 that were adjusted for (if any), as well as the historical baseline period for re-

638 porting the change. Table 9 is presented for reference purposes and not intended
639 to be an exhaustive comparison to all studies published in this domain.

640 Overall, results of our global analysis are broadly consistent with the pat-
641 terns found in published region-specific case studies. For example, Archer et al.
642 (2020) and Berman and Ebisu (2020) both investigated air pollution anomalies
643 in the United States. Archer et al. (2020) compared the average NO₂ levels in
644 April to the corresponding period in 2015–2019, observing a mean reduction of
645 2.02 ppb and reductions up to 8 ppb for some locations. Adjustments for the
646 annual trend (but not for weather effects) led to a final estimate of -1.3 ppb.
647 Berman and Ebisu (2020) found NO₂ reductions of 4.8 ppb in daily 1-h maxi-
648 mum NO₂ between 13 March and 21 April, compared to a baseline of 2017–2019,
649 without adjustments for trends or weather effects. Our study found reductions
650 in between these studies: a 2.5 ppb weather- and trend-adjusted reduction in
651 24-h average NO₂ between 1 April and 30 April, and 2.9 ppb between 13 March
652 and 21 April.

| Study | Region | Time period | Pollutant | Reported estimate | Our estimate |
|------------------------------|-------------------|---------------|-------------------|---|-------------------------|
| Adams (2020) | Ontario, Canada | 22-Mar–25-Apr | NO ₂ | −2 ppb | −1.9 ppb |
| Adams (2020) | Ontario, Canada | 22-Mar–25-Apr | PM _{2.5} | 0 μg/m ³ | −0.6 μg/m ³ |
| Adams (2020) | Ontario, Canada | 22-Mar–25-Apr | O ₃ | −1 ppb | 1.1 ppb |
| Archer et al. (2020) | United States | 1-Apr–30-Apr | NO ₂ | −2.02, −1.3 ppb ^a | −2.5 ppb |
| Archer et al. (2020) | United States | 1-Apr–30-Apr | PM _{2.5} | 0.05, 0.28 μg/m ³ ^a | −0.1 μg/m ³ |
| Berman and Ebisu (2020) | United States | 8 Jan–12 Mar | NO ₂ | −1.17 ppb | −0.6 ppb |
| Berman and Ebisu (2020) | United States | 13 Mar–21 Apr | NO ₂ | −4.76 ppb | −2.9 ppb |
| Berman and Ebisu (2020) | United States | 8 Jan–12 Mar | PM _{2.5} | −0.29 μg/m ³ | −0.6 μg/m ³ |
| Berman and Ebisu (2020) | United States | 13 Mar–21 Apr | PM _{2.5} | −0.28 μg/m ³ | −0.1 μg/m ³ |
| Connerton et al. (2020) | Los Angeles, USA | 1-Mar–31-Mar | PM _{2.5} | −2.99 μg/m ³ | −1.9 μg/m ³ |
| Connerton et al. (2020) | New York, USA | 1-Mar–31-Mar | PM _{2.5} | −2.03 μg/m ³ | −4.5 μg/m ³ |
| Connerton et al. (2020) | Paris, France | 1-Mar–31-Mar | PM _{2.5} | −2.56 μg/m ³ | −2.4 μg/m ³ |
| Connerton et al. (2020) | São Paulo, Brazil | 1-Mar–31-Mar | PM _{2.5} | −0.54 μg/m ³ | −0.3 μg/m ³ |
| Jia et al. (2020) | Memphis, USA | 25-Mar–4-May | PM _{2.5} | 0.3 μg/m ³ | 0.1 μg/m ³ |
| Jia et al. (2020) | Memphis, USA | 25-Mar–4-May | O ₃ | −1.9 ppb | −1.6 ppb |
| Ordóñez et al. (2020) | Europe | 15-Mar–30-Apr | NO ₂ | −9.2, −13.1 μg/m ³ ^b | −10.7 μg/m ³ |
| Ordóñez et al. (2020) | Europe | 15-Mar–30-Apr | O ₃ | 6.2, 0.1 μg/m ³ ^b | 4.5 μg/m ³ |
| Petetin et al. (2020) | Spain | 14-Mar–29-Mar | NO ₂ | −3.4, −5.6 ppb ^c | −5.9 ppb |
| Petetin et al. (2020) | Spain | 30-Mar–9-Apr | NO ₂ | −5.2, −7.4 ppb ^c | −6.8 ppb |
| Petetin et al. (2020) | Spain | 10-Apr–23-Apr | NO ₂ | −4.3, −6.8 ppb ^c | −6.1 ppb |
| Ropkins and Tate (2021) | United Kingdom | 10-Mar–10-Apr | NO ₂ | −4.16, −7.58 μg/m ³ ^d | −8.1 μg/m ³ |
| Ropkins and Tate (2021) | United Kingdom | 10-Mar–10-Apr | PM _{2.5} | 4.79, 5 μg/m ³ ^d | −1.7 μg/m ³ |
| Ropkins and Tate (2021) | United Kingdom | 10-Mar–10-Apr | O ₃ | 6.96, 7.39 μg/m ³ ^d | 5.8 μg/m ³ |
| Tanzer-Gruener et al. (2020) | Pittsburgh, USA | 14 Mar–30 Apr | PM _{2.5} | −2.8 μg/m ³ | −1.7 μg/m ³ |
| Venter et al. (2020) | China | 24-Jan–15-May | PM _{2.5} | −16 μg/m ³ | −9.2 μg/m ³ |
| Venter et al. (2020) | India | 29-Feb–15-May | PM _{2.5} | −15 μg/m ³ | −15.6 μg/m ³ |
| Zheng et al. (2020) | Wuhan, China | 23-Jan–22-Feb | PM _{2.5} | −24.8 μg/m ³ | −28.9 μg/m ³ |

^a comparison to 2015–2019 and trend-adjusted change, respectively.

^b comparison to 2015–2019 and meteorology-adjusted change, respectively.

^c meteorology-normalised changes at background and traffic sites, respectively.

^d break-point/segment analyses for urban background and traffic sites, respectively.

Table 9. Comparison of results to other studies that quantified air pollution changes during lockdowns in 2020.

653 *4.2. Limitations*

654 Our study has a few limitations. For example, smaller countries did not
655 have as many XGBoost models available for a multi-model ensemble forecast,
656 resulting in larger observed variations. Further, major cities with a population
657 over 300,000 were investigated in this study to provide country-level pollution
658 estimates. Since pollution is comparatively high in urban locations, other studies
659 may find lower reductions in pollution when also including rural areas. For the
660 Google COVID-19 Community Mobility Reports, it should be noted that the
661 accuracy of place categorisation may vary from country to country. However,
662 as mobility is reported as the percentage change from baseline in the same
663 country, the measurements are consistent within a country. Alternative data
664 sources include Apple Maps Mobility Trends Reports (Apple, 2020), providing
665 daily percent changes in Apple Maps routing requests from baseline levels at
666 January 13, 2020. However, a limitation of this data source is that travellers
667 mainly require routing information for new routes and the full extent of daily
668 mobility on routes that are well-known to a traveller may not be captured.

669 *4.3. Impact*

670 It is well known that exposure to air pollution can lead to adverse health out-
671 comes, such as respiratory and cardiovascular illnesses. Every year, ambient air
672 pollution results in an estimated 4.2 million and 254,000 premature deaths from
673 PM_{2.5} and O₃ exposure, respectively (Cohen et al., 2017). Indoor air pollution
674 leads to an additional 3.8 million deaths per year, caused by the use of solid
675 fuels and kerosene for heating and cooking (World Health Organization, 2018).
676 Globally, over half of deaths related to PM_{2.5} exposure occur in China and In-
677 dia. PM_{2.5} emissions related to residential energy use have the largest impact
678 on mortality in these countries (Lelieveld et al., 2015). Our study shows the
679 largest reductions in ambient particulate matter during the pandemic occurred
680 in China and India, where adverse health consequences of air pollution are most
681 prevalent. Other studies have shown that mobility reductions also influence the
682 intraday patterns of air pollutants. For example, mobility restrictions could

683 eliminate the morning rush-hour peak in $\text{PM}_{2.5}$, reducing acute exposures in
684 high-traffic environments (Tanzer-Gruener et al., 2020). However, He et al.
685 (2020) found that $\text{PM}_{2.5}$ levels in China during lockdown periods were still four
686 times higher than the WHO air quality guidelines for annual mean $\text{PM}_{2.5}$ of 10
687 $\mu\text{g}/\text{m}^3$ (World Health Organization, 2006). Lockdown periods also resulted in
688 elevated exposure to indoor particulate matter. For example, residents in India
689 spent over 30% more time at home during the initial lockdown (see Fig. 6).
690 Although targeting mobility reductions in these countries can be effective in
691 reducing exposure to ambient air pollution, the unintended consequences of in-
692 creased exposure through indoor pollution should be carefully considered in any
693 interventions.

694 With respect to O_3 , reduced titration with NO_x and volatile organic com-
695 pounds was discussed in many studies as an explanation for increased O_3 levels.
696 This process is the dominant mechanism affecting O_3 levels in winter (Yang
697 et al., 2019). The majority of studies have investigated pollution levels during
698 the initial lockdown, which occurred during late winter and early spring in the
699 Northern Hemisphere. Our study also observed increased O_3 in various coun-
700 tries during these initial lockdowns. However, a subsequent reduction in O_3
701 was observed during summer months in the Northern Hemisphere (i.e., below
702 levels expected based on meteorological conditions), which was outside the pe-
703 riod of investigation of many other studies. As these reductions occurred when
704 O_3 levels typically peak, this period is important to investigate with respect to
705 the health consequences of altered O_3 levels. Hence, the formation of O_3 is an
706 important research question to investigate further, especially in relation to the
707 development of effective mitigation strategies.

708 **5. Conclusions**

709 The COVID-19 pandemic has prompted extraordinary measures to limit
710 disease transmission, restricting mobility of citizens around the world. These
711 restrictions also influenced air pollution, although quantification has proven to

712 be complex. The aim of our study was to disentangle confounding factors in a
713 city-level analysis at a global scale. Non-parametric machine learning methods
714 were used to model the non-linear relationship between atmospheric processes
715 and air pollution, adjusting for weather, seasonality, trends and city charac-
716 teristics such as topography. Daily pollution levels throughout 2020 were pre-
717 dicted in a counterfactual scenario, agnostic of the pandemic, and compared
718 to the observed air pollution. Finally, city-level predictions were assessed for
719 each country, providing a country-specific multi-model ensemble forecast for the
720 meteorology-normalised pollution reductions. It was shown that results were
721 broadly consistent with those reported elsewhere. Hence, one of the contribu-
722 tions of this study is showing how case studies could be scaled up to a global
723 level.

724 The unexplained pollution matched the timing of government restrictions in
725 different countries very well, based on data from the Oxford COVID-19 Govern-
726 ment Response Tracker. The largest reductions in NO_2 were observed during
727 the initial lockdowns in China, Europe and India. Further, pollution did not
728 fully revert to pre-pandemic levels throughout 2020. For particulate matter,
729 the largest absolute reductions occurred in China and India. In many coun-
730 tries, increased O_3 levels were observed during the initial lockdowns, which has
731 been documented widely (see Section 1.1.3). Reduced titration of O_3 with NO_x
732 and volatile organic compounds was generally postulated as the main driver;
733 this process is the dominant mechanism affecting O_3 levels in winter. However,
734 during the summer period from June/July to September, our analyses found a
735 subsequent reduction in O_3 for many countries below what was expected based
736 on meteorological conditions (e.g., China, United Kingdom, France, Germany,
737 Poland, Turkey). This was outside the period of investigation of many other
738 studies, but important for the development of effective mitigation strategies
739 with respect to health consequences. The adverse health consequences of air
740 pollution are most prevalent in China and India. Although targeting mobility
741 reductions in these countries can be effective in reducing exposure to ambient
742 air pollution, the unintended consequences of increased exposure through indoor

743 PM_{2.5} pollution should be carefully considered in any intervention. Note that
744 in many countries, NO₂ reductions were highly correlated with changes in mo-
745 bility levels, especially trips to transit stations, workplaces, retail and recreation
746 venues. Therefore, alternative transport initiatives for these type of trips, such
747 as electric vehicles, bicycles and ride sharing, could be an important pathway
748 towards improved health outcomes.

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