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Synergistic effects of harvest and climate drive synchronous somatic growth within key New Zealand fisheries

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8 **Synergistic effects of harvest and climate drive synchronous somatic growth**
9 **within key New Zealand fisheries**

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11 **Running head: Fishing and climate drive growth synchrony**

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25

26 **Abstract**

27 Fisheries harvest has pervasive impacts on wild fish populations, including the truncation of
28 size and age structures, altered population dynamics and density, and modified habitat and
29 assemblage composition. Understanding the degree to which harvest-induced impacts
30 increase the sensitivity of individuals, populations and ultimately species to environmental
31 change is essential to ensuring sustainable fisheries management in a rapidly changing
32 world. Here, we generated multiple long-term (44 to 62 years), annually resolved, somatic
33 growth chronologies of four commercially important fishes from New Zealand's coastal and
34 shelf waters. We used these novel data to investigate how regional- and basin-scale
35 environmental variability, in concert with fishing activity, affected individual somatic growth
36 rates and the magnitude of spatial synchrony among stocks. Changes in somatic growth can
37 affect individual fitness and a range of population and fishery metrics such as recruitment
38 success, maturation schedules and stock biomass. Across all species, individual growth
39 benefited from a fishing-induced release of density controls. For nearshore snapper and
40 tarakihi, regional-scale wind and temperature also additively affected growth, indicating
41 that future climate change-induced warming and potentially strengthened winds will
42 initially promote the productivity of more poleward populations. Fishing increased the
43 sensitivity of deep-water hoki and ling growth to the Interdecadal Pacific Oscillation (IPO). A
44 forecast shift to a positive IPO phase, in concert with current harvest strategies, will likely
45 promote individual hoki and ling growth. At the species level, historical fishing practices and
46 IPO synergised to strengthen spatial synchrony in average growth between stocks separated
47 by 400-600 nm of ocean. Increased spatial synchrony can, however, increase the
48 vulnerability of stocks to deleterious stochastic events. Together, our individual- and
49 species-level results show how fishing and environmental factors can conflate to initially
50 promote individual growth but then possibly heighten the sensitivity of stocks to
51 environmental change.

52 **Keywords**

53 Spatial synchrony; otolith biochronology; climate change; dispersal; fishing; Pacific Ocean;
54 Moran Effect; population dynamics; density-dependent growth

55

56 **Introduction**

57 The degree to which life history parameters are spatially synchronous plays a critical role in
58 determining the abundance and persistence of wild populations, as greater levels of
59 synchrony increase the likelihood that all populations in a region will express good or poor
60 performance simultaneously (Liebhold *et al.*, 2004, Schindler *et al.*, 2010). Spatial synchrony
61 is a common feature of wild populations, with its magnitude and extent dependent on rates
62 of individual dispersal, strength of environmental factors ('Moran Effect'), or the widespread
63 influence of interspecific trophic interactions (reviewed by Liebhold *et al.*, 2004).

64 Importantly, anthropogenic stress, such as that induced by fisheries harvest, can also drive
65 an increase in spatial synchrony (Frank *et al.*, 2016), as age and size truncation decrease
66 demographic buffering and in turn increase instability in population dynamics (Anderson *et al.*,
67 2008, Gamelon *et al.*, 2019, Rouyer *et al.*, 2012). Understanding the causes and
68 consequences of spatiotemporal population dynamics greatly enhances our capacity to
69 predict and manage natural systems and natural resource yields in the face of rapid
70 anthropogenic environmental change (Engen *et al.*, 2018, Rouyer *et al.*, 2012, Schindler *et al.*,
71 2010).

72 Sustainable management of wild fisheries relies on sound understanding of how individual
73 fish, and the populations they comprise, respond to harvest and stochastic or directional
74 environmental forcing such as ocean warming (Audzijonyte *et al.*, 2016, Kjesbu *et al.*, 2014,
75 Le Bris *et al.*, 2018). Historically, the impacts of fishing and environmental fluctuations on
76 population dynamics have been considered in isolation. Harvest is seen as exploitative and
77 affects population averages, whereas climate fluctuations affect population variance
78 (reviewed by Planque *et al.*, 2010). Potential synergies between the two have only come
79 into focus more recently. For example, increased adult mortality due to fishing can decrease
80 recruitment strength and thus increase the sensitivity of recruitment success to

81 environmental variation (Rouyer *et al.*, 2011). Similarly, harvest can select on reproductive
82 phenology and drive potential mismatches between spawning windows and favourable
83 environmental conditions (Rogers & Dougherty, 2019), or constrain the expression of
84 thermally-dependent growth phenotypes through alterations to social hierarchies or density
85 (Morrongiello *et al.*, 2019). Habitat loss and degradation due to fishing activity can reduce
86 carrying capacities and also increase the sensitivity of populations to additional
87 environmental change (Brown *et al.*, 2019). Conversely, fishing-induced reductions in
88 density (Lorenzen & Enberg, 2002) can facilitate fish survival in potentially sub-optimal
89 warmer waters (Ciannelli *et al.*, 2004). Importantly here, both warming (Morrongiello &
90 Thresher, 2015) and harvest (Schindler *et al.*, 2010) are known to induce spatial synchrony
91 in fish populations. Increased synchrony can heighten the vulnerability of individual stocks
92 to deleterious stochastic events as struggling populations will fail to receive demographic
93 rescue from more successful populations (Heino *et al.*, 1997), and could even promote
94 species-wide fishery collapse (Lobón-Cerviá, 2009).

95 The World's oceans have experienced recent rapid warming (Cheng *et al.*, 2019) and
96 prolonged fisheries exploitation (Hilborn *et al.*, 2020). These changes have occurred against
97 a backdrop of large and small-scale natural environmental variations that affect individual
98 fitness and population dynamics. For example, El Niño/Southern Oscillation (ENSO) and the
99 Interdecadal Pacific Oscillation (IPO) are basin-wide climate and oceanographic signals and
100 have been implicated in ecosystem regime shifts (e.g. Hare & Mantua, 2000). On a more
101 regional scale, interannual variation in zonal wind strength affects fish recruitment patterns
102 (Jenkins, 2005) whilst the magnitude of upwelling can influence fish growth (Black *et al.*,
103 2011). It is essential to understand how regional and basin-scale anthropogenically and
104 naturally driven environmental variations interact with fishing activity to influence patterns
105 of spatial synchrony. In turn, this understanding informs the adaptation of fisheries
106 management and fisher behaviour to ensure fisheries viability in the future (Audzijonyte *et*
107 *al.*, 2016, Perry *et al.*, 2010, Planque *et al.*, 2010).

108 Somatic growth rate describes how an individual's size-at-age changes through time and is
109 highly sensitive to a range of intrinsic (within individual), natural and anthropogenic extrinsic
110 factors (Enberg *et al.*, 2012, Morrongiello & Thresher, 2015). Growth is an ideal candidate
111 with which to explore the impacts of fishing and environmental variability because it not

112 only underpins individual fitness (e.g., size-dependent survival and reproduction, Barneche
113 *et al.*, 2018, Sogard, 1997), but also a range of population-level metrics such as recruitment
114 success (Bergenius *et al.*, 2002), maturity schedules (Uusi-Heikkilä *et al.*, 2015), and stock
115 biomass (Stawitz & Essington, 2019). Growth rate is dependent on the amount of acquired
116 resources and subsequent allocation trade-offs among maintenance and reproduction
117 (Enberg *et al.*, 2012). Environmental- and fishing-induced changes to resource availability
118 and competition can thus affect the expression of growth (Morrongiello *et al.*, 2019).
119 Likewise, size selective harvest and increased mortality can select for earlier age at
120 maturation and an associated increase in juvenile growth rate (Uusi-Heikkilä *et al.*, 2015),
121 whilst temperature-dependent metabolic costs can limit energy available to be allocated
122 towards growth (Neuheimer *et al.*, 2011). Therefore, whilst it is intuitive to think of
123 elevated growth rate as being ‘good’ for an individual, it is important to acknowledge that it
124 can also be symptomatic of a population under stress (Trippel, 1995). The growth
125 information naturally archived in fish otoliths offers an opportunity to identify and quantify
126 the drivers of individual growth variation within wild populations (Morrongiello *et al.*, 2012)
127 and spatial synchrony at the species and assemblage level (Martino *et al.*, 2019, Ong *et al.*,
128 2018, Tanner *et al.*, 2020).

129 The waters around New Zealand support significant commercial fisheries on both inshore
130 and deep-water species (Fisheries New Zealand, 2019) and have undergone warming
131 associated with climate change (Sutton & Bowen, 2019). Here, we use the unique biological
132 information naturally archived in fish otoliths collected during routine fisheries assessments
133 to recreate 44- to 62-year-long biochronologies of average annual fish growth for four
134 commercially important fishes. We then relate these estimates of inter-annual growth
135 variability to environmental changes that are manifest on both basin- and regional spatial
136 scales. We adopt a hierarchy of analyses to sequentially ask: 1) do environmental and fishing
137 variables additively or synergistically affect the expression of individual growth within
138 species; 2) what is the level of spatial synchrony in growth within and among focal species;
139 and 3) how does basin and regional-scale environmental variability and fishing pressure
140 affect the magnitude and direction of spatial synchrony? We predict that individual fish
141 growth will become more sensitive to environmental drivers during periods of elevated
142 fishing pressure. Furthermore, spatial synchrony will be greatest within species across areas,

143 followed by between species in the same area and then between species across areas.
144 These synchrony patterns arise because of implicit similarities and differences in
145 environmental and fishing responses within and among species respectively, and potential
146 dispersal between stocks linking population dynamics. By simultaneously assessing how
147 environmental variability and fishing pressure affect individuals, populations and species,
148 we can develop a holistic understanding of the impact of these changes and in turn make
149 more meaningful predictions of future fishery productivity.

150

151 **Methods**

152 *Study species and system*

153 We developed growth biochronologies for two inshore and two deep-water New Zealand
154 fishes. Snapper (*Chrysophrys auratus*) and tarakihi (*Nemadactylus macropterus*) are
155 demersal inshore species occurring in water depths from 5 to 250 m (Anderson *et al.*, 1998).
156 Both species have been sampled regularly since the 1980s from the northeast of North
157 Island's Hauraki Gulf and East Cape region (snapper: HAGU; tarakihi: NENI), and off northern
158 South Island primarily in Tasman Bay and the west coast (snapper: TBGB; tarakihi: TBWC,
159 Figure 1a). Tagging data and analyses of growth rates and year class strengths indicate that
160 both species comprise separate stocks in each sampled area (Fisheries New Zealand,
161 2019), with tarakihi tending to move more than snapper.

162 Hoki (*Macruronus novaezelandiae*) and ling (*Genypterus blacodes*) are predominantly deep-
163 water demersal species abundant around South Island in depths from 200 to 800 m
164 (Anderson *et al.*, 1998). Trawl surveys monitoring these species have been conducted
165 regularly since the 1980s on Chatham Rise and the Campbell Plateau (Figure 1a). Hoki
166 comprise two stocks, one having a home ground on the Campbell Plateau (CAMP) and a
167 spawning ground off the west coast of South Island, and the other spawning in Cook Strait
168 and having a home ground on Chatham Rise (CHAT; Livingston *et al.*, 2015). Ling on
169 Chatham Rise (CHAT) are clearly distinct from Campbell Plateau (CAMP) fish because of
170 marked between-area differences in growth rates and year class strengths (Horn, 2005). All

171 eight nearshore and deep-water stocks are assessed and managed separately (Fisheries
172 New Zealand, 2019). See Table S1 for additional biological and management details.

173 *Otolith samples*

174 For each area-species combination, we followed the sampling design of Morrongiello et al.
175 (2012) and selected otoliths based on their year of collection and previously estimated age.
176 This approach ensured good coverage of year classes (3-78 fish per collection, $\bar{x}=28.7$),
177 replication of fish ages throughout the reconstructed growth time series, and 3681-5330
178 increment measurements from 435-521 fish per area (Figure 1b; Table 1). All otoliths were
179 re-aged following validated otolith interpretation protocols (Horn & Sutton, 2017, Horn,
180 1993, Walsh *et al.*, 2016, Walsh *et al.*, 2014). We sampled snapper aged 3–20 years, tarakihi
181 aged 2–26, hoki aged 4–15, and ling aged 4–19 (i.e., if possible avoiding the oldest fish
182 which were likely to have very narrow, and hence, less informative growth zones near the
183 otolith margin). It was necessary, however, to sample 2–10% of fish from older ages to
184 enable the planned coverage of year classes across sample years. We aimed to sample 9–10
185 fish from each species-area year class, with even proportions by sex.

186 Annual growth increments were measured from images of sagittal otolith cross-sections
187 (Figure 1c) in ImageJ (Abràmoff *et al.*, 2004). Our approach assumes that otolith increments
188 are a proxy of fish growth. We tested and confirmed this assumption for all species by
189 regressing fish length on otolith radius (snapper $R^2 = 0.705$, $n = 938$; tarakihi: $R^2 = 0.643$, $n =$
190 985 ; hoki: $R^2 = 0.429$, $n = 1005$; ling: $R^2 = 0.830$, $n = 1012$).

191 *Fishery and environmental predictors*

192 For each of the eight species-area combinations, we developed a series of environmental
193 and fishery-related predictor variables (defined in Table 2; Figure S1). Environmental
194 predictor series were derived for the basin-scale Southern Oscillation Index (*SOI*) and inter-
195 decadal Pacific Oscillation Index (*IPO*), and the regional scale sea surface temperature (*SST*),
196 Trenberth indices of wind direction and strength (*Z & M*), and offshore *wind stress*
197 magnitude. Temperature-at depth data (e.g. Simple Ocean Data Assimilation [SODA3]
198 reanalysis: Carton *et al.*, 2018) was not available for the full study period and therefore we
199 used *SST* to describe the benthic thermal regime inhabited by our focal species. This is

200 justified for nearshore species as the maximum mixed layer depth in New Zealand waters is
201 250 m (Bowen *et al.*, 2017), which encompasses the modal depth of catch for snapper and
202 tarakihi (~20 m and ~150 m respectfully; Anderson *et al.*, 1998, Fisheries New Zealand,
203 2019). SST variation around New Zealand has also been linked to water temperatures down
204 to 800 m (Bowen *et al.*, 2017, Sutton *et al.*, 2005) which encompasses the environment of
205 our deep-water species (modal catch for hoki ~ 550 m and ling ~ 500 m; Anderson *et al.*,
206 1998, Fisheries New Zealand, 2019). Series of spawning stock biomass (*SSB*) and fishing
207 pressure (*FP*; catch/biomass vulnerable to fishing) were derived from stock assessments
208 (Fisheries New Zealand, 2019).

209 The environmental series were constructed in line with each species' 'growth year', i.e., the
210 year beginning at the fish's birthday. For both hoki and ling, the growth year is September to
211 the following August. For snapper, the growth year was from January to the following
212 December and for tarakihi April to the following March. For each available series, an annual
213 index was developed (e.g., for ling: mean *SST* for the September–August year), as well as
214 four three-monthly indices (e.g., for ling: September–November, December–February,
215 March–May, June–August). We also estimated environmental conditions in the year
216 preceding fish growth to explore potential lagged effects, necessary when examining
217 linkages between the higher-order consumer production observed here and environmental
218 factors that may operate through trophic pathways (Nguyen *et al.*, 2015).

219 **Statistical analyses**

220 Variation in annual growth increments were analysed using mixed effects modelling
221 (Morrongiello & Thresher, 2015). A random intercept for each *Fish* induces a correlation
222 among increments to account for any individual-specific differences in growth and the
223 repeated measures nature of otolith increment data. A random intercept for *Area:year*
224 (unique identifier for each growth year within an area) induces a correlation among
225 increments from different fish living in the same area that were deposited in the same year,
226 and accounts for potential temporal growth differences induced by common environmental
227 effects. Otolith increment width exponentially declines with fish age and this pattern is
228 generally linearised by taking the natural log of both variables. However, a preliminary
229 exploration of model residuals indicated that such an approach did not fully describe the

230 growth-age relationship for the four species. Therefore, we explored whether the inclusion
 231 of a quadratic for log-transformed age improved model fit. Mixed effects models were
 232 developed using the ‘lme4’ package (Bates *et al.*, 2015) in R 3.5.1.

233 *Intrinsic drivers of annual growth and biochronology development*

234 We developed a series of base growth models for each species that included three-way
 235 interactions between *Age* (linear or quadratic) and *Area* and *Sex*. These terms allowed for
 236 spatial and sex-specific differences in the average growth-age relationship. We also included
 237 an interaction between *age-at-capture* (AAC) and *Area* to account for potential biases in
 238 estimated growth rates associated with differential mortality or sampling regime at each
 239 area (see Morrongiello *et al.*, 2012 for discussion). *Age* and *age-at-capture* were scaled
 240 $[(x - \bar{x})/s]$ to facilitate model convergence and interpretation of interactions and quadratic
 241 terms. Base models included random intercepts for *Fish* and *Area:year* and their associated
 242 random *Age* (linear and quadratic) slopes. These complex random slope structures allow for
 243 individual-specific and year -specific deviations from the average growth-age relationship.
 244 The most complex base intrinsic effects model can be represented by:

$$245 \quad y_{ijklmn} = \alpha_k + \alpha_{Yk,l} + \alpha_{Fi} + \beta_A x_{j(k,m)} + \beta_A^2 x_{j(k,m)} + \gamma_{Czn} a_k + b_{FA} x_{ij} + b_{YAx_{j(k),l}} + \varepsilon_{ijk}$$

$$246 \quad \begin{bmatrix} \alpha_{Fi} \\ b_{FAi} \end{bmatrix} N(0, \Sigma_i), \begin{bmatrix} \alpha_{Yk,l} \\ b_{YAk,l} \end{bmatrix} N(0, \Sigma_{k,l}), \varepsilon_{ijk} N(0, \sigma^2)$$

(Eq. 1)

247 where y_{ijklmn} is the annual otolith increment (μm) for the i th fish at age j from area k in year l
 248 and sex m that is caught at age n , α_k is the fixed (area-specific average) annual growth
 249 intercept, $\alpha_{Yk,l}$ is the random extrinsic effect for *Area* k at *growth year* l , α_{Fi} is the random
 250 intrinsic fish effect, $\beta_{Aj(k,m)}$ and $\beta_{Aj(k,m)}^2$ describe the linear and quadratic components of age-
 251 dependent change in growth specific to each *Area* k and *Sex* m , $\gamma_{Cn(k)}$ is the interaction
 252 between *age-at-capture* z and *Area* k , b_{FAij} is the random slope of *Age* for fish i , $b_{YAj,k,l}$ is the
 253 random slope of *Age* for *Area* k at *growth year* l and ε is assumed to be independent errors
 254 with mean zero and common variance σ^2 .

255 We compared the performance of eight random effects structures that included different
 256 combinations of linear and quadratic *Age* random slopes for *Fish* and *Area:year*. These
 257 models were forced to include the maximal intrinsic effects structure (*Age * Area * Sex +*

258 $Age^2 * Area * Sex + age-at-capture * Area$), were fit with restricted maximum likelihood
259 (REML), and compared using Akaike's Information Criterion corrected for small sample sizes
260 (AIC_c). We then applied each species' best random effects structure to models with
261 increasing fixed intrinsic effect complexity, fit with maximum likelihood (ML), and compared
262 these with AIC_c . The best intrinsic model for each species was reanalysed using REML to
263 produce unbiased parameter estimates.

264 *Extrinsic drivers of annual growth: basin-wide versus regional scale*

265 The impact of fishery activity, demography and environmental variation on each species'
266 annual growth was assessed by including additional extrinsic variables to the best intrinsic
267 model structure for each species (from Eq. 1). *SST*, *Wind Stress*, *SSB* and *FP* were scaled
268 $[(x_{ijk} - \hat{x}_{jk})/s_{jk}]$ within each area (j) and species (k). *IPO*, *SOI* and *Trenberth M* and *Z* values
269 were the same across areas for a species and so were just scaled $[(x_{ik} - \hat{x}_k)/s_k]$ within each
270 species (k). We included an interaction term with *Area* for all scaled variables so that growth
271 responses could be conditional on location. We built and compared two broad suites of
272 models: those that included basin-scale environmental drivers (*IPO* or *SOI*, which both affect
273 *SST* and winds) and those that included more regionalised environmental drivers (*SST*, wind
274 indices). We allowed *SSB* and *FP* to be included in both model sets and tested for
275 interactions between these fishing variables and basin-wide and regional environmental
276 drivers. *IPO* and *SOI* both describe large-scale environmental processes, and *Trenberth M*
277 and *Z* and *Wind Stress* (offshore) all describe wind-related phenomena. *SSB* and *FP* are
278 related to fishing activity. Competing models therefore included either *IPO* and *SOI*, either
279 of the three wind variables, and either of the two fishing activity series to allow for
280 meaningful comparisons of their relative importance. We included just linear terms for each
281 extrinsic variable to manage model complexity. Competing environmental models were fit
282 with ML and compared using AIC_c . The best environmental model for each species was then
283 refit with REML. Parameter estimates and 95% credible intervals were derived from the
284 posterior distribution of the fixed effects in the best model using 1000 model simulations
285 generated by the 'arm' package (Gelman & Su, 2020), and model predictions generated
286 using the 'effects' package (Fox & Weisberg, 2019).

287 *Spatial synchrony in annual growth*

288 We extracted the *Area.year* random effect conditional modes (best linear unbiased
289 predictors, BLUPs) from each species' best base intrinsic effects model (model not including
290 fixed extrinsic variables) and used these biochronologies to explore spatial synchrony in
291 growth within and among species (*sensu* Tanner *et al.*, 2020). Temporal trends were
292 removed from each biochronology using a first-order autoregressive model that included a
293 linear term for year and the resultant residuals taken forward for further analysis
294 (Buonaccorsi *et al.*, 2001). Zero-lag cross correlations were performed between each
295 residual growth series to estimate overall levels of spatial synchrony. Then, we calculated
296 21-year centre-aligned running correlations (using the 'running' function from the 'gtools'
297 package in R: Warnes *et al.*, 2020) between the residual growth series within each species to
298 explore temporal patterns in spatial synchrony. A span of 21 years was selected as a
299 compromise between correlation robustness (maximised by greater span) and the need to
300 identify shorter-term changes in synchrony (maximised by shorter span). Running
301 correlation significance at the 0.05 level was assessed by taking 1000 Monte Carlo
302 simulations of 21-year running correlations based on randomly generated data of the same
303 length as that of each growth time series. We extracted the highest running correlation for
304 each simulation and used the resulting distribution to identify the critical threshold of
305 running correlations above which they are significant at $p=0.05$. Repeatedly testing the
306 significance of each running correlation without simulation would greatly inflate Type 1
307 error rates due to multiple comparisons.

308 Where necessary, we scaled and then took time-dependent averages of predictor variables
309 that differed between areas for a species (*SSB*, *FP*, *SST*, *wind stress*) to generate a single
310 series. We then calculated 21-year centre-aligned rolling means (using the 'rollmean'
311 function from the 'zoo' package in R) for each annually derived fishery and environmental
312 variable to match the growth running correlation data described above. We explored how
313 these time-averaged annual predictor variables influenced the degree of within-species
314 spatial synchrony using a first order autoregressive model that accounted for temporal
315 autocorrelation. Each environmental variable was interacted with fishing variables and
316 competing models, along with the null which included just a model intercept, were
317 compared using AICc.

318

319 Results

320 *Intrinsic drivers of growth*

321 The best supported random effects structure for all species included random linear and
322 quadratic age slopes for *Fish* (Table S2), indicating that the growth-age relationship varied
323 among individual fish. Snapper and ling models also included random linear and quadratic
324 age slopes for *Area:year* (Table S2), indicating that there were differences in age-dependent
325 growth across years. There was some variation in the composition of the best intrinsic
326 effects model among species (Table S3). For all species, growth declined with age and varied
327 across regions (Table 3; Figure 2). Sex-dependent differences were only evident for ling,
328 where females consistently grew faster than males from about 7 years of age (Table 3; Figure
329 3). Age-dependent biases in estimated growth rates were detected, and accounted for, in all
330 species using the *age-at-capture* term. Specifically, snapper and tarakihi caught at older
331 ages consistently grew slower than fish caught when younger, whereas ling displayed the
332 opposite pattern with fish caught at older ages being the faster growers (Table 3). The
333 strength of these patterns differed between regions for snapper and ling.

334 *Extrinsic drivers of annual growth: basin-wide versus regional scale*

335 The growth of all four species was best explained by combinations of environmental factors
336 and fishing activity calculated over either the winter-spring or lagged annual time periods
337 (Table S4). Nearshore (snapper and tarakihi) and deep-water (hoki and ling) species
338 systematically differed in the importance of regional versus basin-wide environmental
339 drivers, and the synergistic impacts of fishing and the environment, on growth variation.

340 Snapper growth increased at a rate of $2.5\%.\text{°C}^{-1}$ following warmer years (Figure 4a) and
341 decreased at higher spawning stock biomasses (Figure 4b). There was a regional difference in
342 how snapper responded to the previous years' *Trenberth Z1*, with a negative response in
343 TBGB compared to a negligible response in HAGU (Figure 4c). The strength, but not direction,
344 of the tarakihi growth response to *SST* during July-September differed among areas, with
345 the poleward TBWC stock responding more positively ($1.7\%.\text{°C}^{-1}$) than the more equatorial
346 NENI stock ($0.6\%.\text{°C}^{-1}$) (Figure 4d). Tarakihi grew faster in years with higher fishing pressure
347 (Figure 4e) and windier winters (Figure 4f).

348 CHAT hoki growth was negligibly affected by the previous year's spawning stock biomass
349 (**Figure 5a**). The growth of CAMP hoki was, however, much more sensitive to *SSB* with a
350 steep decline in growth in years following large stock sizes (**Figure 5a**). We detected
351 evidence for a synergistic impact of fishing and *IPO* on hoki growth (**Figure 5b**). At low
352 spawning stock biomass, growth was positively related to the previous year's *IPO* index. At
353 median biomass there was negligible effect of *IPO*, whereas at high biomass fish grew
354 slower in years following those with more positive *IPO* indices.

355 CHAT ling growth was lower in years following large spawning stock biomass (**Figure 5c**). In
356 contrast, *SSB* had a negligible impact on CAMP ling growth. Like hoki, we detected
357 synergistic impacts of fishing and *IPO* on ling growth. Ling growth was positively related to
358 *IPO* at low biomass, but negatively related to *IPO* at high biomass (**Figure 5d**).

359 *Spatial synchrony in annual growth*

360 HAGU and TBGB snapper experienced elevated growth during the mid-1980s, mid- to late-
361 1990s, and mid- to late-2000s, and depressed growth during the mid- to late-1960s and
362 1992–1994 (Figure 6a). TBWC tarakihi displayed considerable interannual growth variation,
363 with a period of pronounced growth depression from 1992–1996 which was followed by a
364 period of elevated growth from 2000–2002 (Figure 6b). There was little evidence for strong
365 consistent inter-annual growth variability synchrony among individuals in NENI tarakihi.
366 Annual growth of CAMP and CHAT hoki increased in the decade following an early 1980s
367 low point, and again for five years from a low point in the early 2000s (Figure 6c). Ling
368 growth steadily declined for over a decade from 1980, and then remained depressed from
369 1993 onwards (Figure 6d). We detected no evidence for temporal trends in either snapper or
370 tarakihi growth (area-specific OLS regressions, $p > 0.242$). Hoki CHAT growth increased by
371 $6.5\%.\text{decade}^{-1}$ ($p = 0.002$), whilst ling CAMP and ling CHAT decreased by $4.5\%.\text{decade}^{-1}$ ($p <$
372 0.001) and $4\%.\text{decade}^{-1}$ ($p < 0.001$) respectively.

373 Overall, fish growth rates were temporally synchronous across areas within species for
374 snapper, hoki and ling ($r(\text{raw}): 0.395\text{--}0.694$, $p \leq 0.009$) and these patterns persisted after
375 accounting for temporal trends and autocorrelation ($r(\text{res}): 0.375\text{--}0.698$, $p \leq 0.013$, Figure
376 6a,c,d). We identified consistent habitat-specific differences in the pattern of 21-year

377 residual growth synchrony through time. Spatial synchrony in nearshore snapper and
378 tarakihi residual growth time series steadily increased through time (Figure 6e-f). Conversely,
379 spatial synchrony in the deep-water hoki and ling initially increased to a peak in the mid to
380 late 1980 and early 1990s, then steadily declined to the early 2000s with some evidence for
381 a recent increase in synchrony thereafter (Figure 6g-h). The best model explaining each
382 species' temporal variation in spatial synchrony included *IPO*, and in three species *SSB* or *FP*
383 (Table 4). Snapper spatial synchrony was highest in 21-year periods characterised by low *SSB*
384 and a negative *IPO* phase (**Figure 7a-b**). Tarakihi spatial synchrony was best explained by an
385 interaction between *IPO* and *FP* whereby correlations were strongest during periods of high
386 fishing pressure and negative *IPO* (**Figure 7c**). Temporal variation in hoki synchrony was
387 primarily related to positive *IPO* conditions (**Figure 7d**), and ling growth synchrony became
388 more sensitive to *IPO* at higher *SSB* (**Figure 7e**). Between species, the only significant
389 correlation in growth time series was between CAMP hoki and CHAT ling growth ($r(\text{res})$:
390 0.350, $p = 0.022$). There was little other evidence for strong growth synchrony among
391 species.

392

393

394 **Table 1:** Summary of otolith sample data used to generate each growth biochronologies.

	Snapper		Tarakihi		Hoki		Ling	
	HAGU	TBGB	NENI	TBWC	CHAT	CAMP	CHAT	CAMP
Number of individuals	446	447	521	435	482	499	504	507
Number of increments	4020	4837	4431	4492	3681	4063	5330	5240
Growth biochronology	1954-2013	1960-2014	1954-2015	1962-2016	1972-2015	1973-2016	1965-2015	1968-2016
span	(60 years)	(55 years)	(62 years)	(55 years)	(44 years)	(44 years)	(51 years)	(49 years)

395

396 **Table 2:** Description of environmental and fishing-related predictor variables used in spatial synchrony and annual growth models for four
 397 species of New Zealand fish.

Variable type	Variable	Description	Available time span
Basin-wide environment	Southern Oscillation Index (SOI)	Difference in surface atmospheric pressure between Tahiti and Darwin; a diagnostic Trade Wind strength and the El Niño/Southern Oscillation (ENSO) system. El Niño's are normally associated with increased westerly and southerly winds (Salinger <i>et al.</i> , 1995) and cooler ocean temperatures around New Zealand (Sutton & Roemmich, 2001).	1950-2016
	Interdecadal Pacific Oscillation (IPO)	A filtered tripole index from Henley <i>et al.</i> (2015). Cooler ocean and stronger south-westerly wind conditions around New Zealand in positive IPO phases (Hannah & Bell, 2012).	1950-2016
Regional environment	Sea surface temperature (SST)	Two SST products are used in this analysis: 1) reconstructed SST (NOAA Extended Reconstructed Sea Surface Temperature: ERSST.v5; available via http://www.esrl.noaa.gov/psd/) with a 2° latitude and longitude spatial resolution and monthly temporal resolution extending back to 1854 (Huang <i>et al.</i> , 2017); 2) objectively-analysed SST based on satellite measurement (NOAA OI	1960-2016

SST V2 High Resolution Dataset). The latter data are higher resolution (0.25° latitude and longitude spatial and daily temporal resolution) but are only available for the period since September 1981 (Reynolds *et al.*, 2007). SST series for each stock were developed for areas bounded by boxes in Figure 1a.

	Trenberth (Z & M)	Changes in large-scale wind fields measured as normalised differences in mean sea level air pressure between points (Trenberth, 1976). By geostrophic balance, the pressure difference between two points is a direct proxy for the average strength of the surface wind perpendicular to that pressure difference. ‘Z’ indices describe zonal (westerly) winds and ‘M’ indices describe Meridional (northerly) winds. The indices and the pairs of geographical points defining them were for snapper and tarakihi: Z1 (Auckland – Christchurch) and M1 (Hobart – Chatham Islands); for hoki and ling: Z2 (Christchurch – Campbell Island) and M2 (Hokitika – Chatham Islands).	1950-2016
	Wind stress (offshore)	Mean offshore <i>wind stress</i> (i.e., the shear stress exerted by the wind on the surface of large bodies of water) in the areas of interest extracted from the JRA-55 reanalysis (http://jra.kishou.go.jp/JRA-55/index_en.html) (Kobayashi <i>et al.</i> , 2015). Related to regional changes in wind forcing and thus regional upper-ocean circulation and mixing.	1960-2016
Fishing	Spawning stock biomass (SSB)	Combined weight of all individuals in a stock that have reached sexual maturity (tonnes). Measure of adult population abundance and, assuming relatively stable distribution, density-dependent competition. Data from Fisheries New Zealand (2019).	Snapper: 1950-2015 Tarakihi: 1950-2017 Hoki: 1972-2016
	Fishing pressure (FP)	Annual catch/ biomass vulnerable to fishing. Measure of fishing intensity and selectivity. Data from Fisheries New Zealand (2019).	Ling: 1971-2016

399 **Table 3:** Fixed effect parameter estimates ($\pm 95\%$ confidence intervals) and random effect variance components for best intrinsic models
 400 explaining annual growth variation in four New Zealand fishes.

	Snapper			tarakihi			hoki			ling		
Fixed effects	estimate (95% CI)			estimate (95% CI)			estimate (95% CI)			estimate (95% CI)		
Intercept	4.284 (4.263, 4.305)			3.812 (3.795, 3.829)			4.439 (4.421, 4.458)			5.804 (5.784, 5.824)		
Age	-0.426 (-0.440, -0.413)			-0.493 (-0.506, -0.480)			-0.479 (-0.490, -0.467)			-0.592 (-0.612, -0.573)		
Age ²	0.156 (0.148, 0.166)			0.019 (0.012, 0.025)			0.069 (0.060, 0.078)			-0.151 (-0.168, -0.135)		
Age-at-capture	-0.051 (-0.066, -0.035)			-0.021 (-0.030, -0.011)			-			0.016 (0.006, 0.026)		
Area	0.205 (0.178, 0.233)			-0.023 (-0.044, -0.001)			0.032 (0.005, 0.057)			-0.096 (-0.121, -0.068)		
Sex	-			-			-			0.049 (0.024, 0.075)		
Age * Area	-0.075 (-0.094, -0.055)			-0.039 (-0.053, -0.025)			-0.020 (-0.036, -0.005)			-0.173 (-0.200, -0.150)		
Age ² * Area	-0.087 (-0.099, -0.074)			-			-0.011 (-0.023, 0.002)			-0.029 (-0.050, -0.009)		
Age * Sex	-			-			-			0.066 (0.044, 0.086)		
Age ² * Sex	-			-			-			0.032 (0.017, 0.048)		
Area * Sex	-			-			-			-0.024 (-0.050, 0.004)		
Age-at-capture * Area	0.031 (0.011, 0.049)			-			-			-0.008 (-0.021, 0.006)		
Random effects	σ	corr		σ	corr		σ	corr		σ	corr	
Fish	0.127			0.149			0.156			0.140		
Age Fish	0.075	0.06		0.104	0.09		0.079	0.19		0.120	0.28	
Age ² Fish	0.049	-0.59	0.42	0.061	-0.49	0.6	0.054	-0.71	0.53	0.091	-0.68	0.4
						3						3
Area:year	0.047			0.017			0.022			0.020		
Age Area:year	0.020	-0.19								0.030	-0.90	

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Age ² Area:year	0.012	-0.60	-0.11			0.032	-0.99	0.8
Residual	0.129			0.136		0.192		6
						0.218		

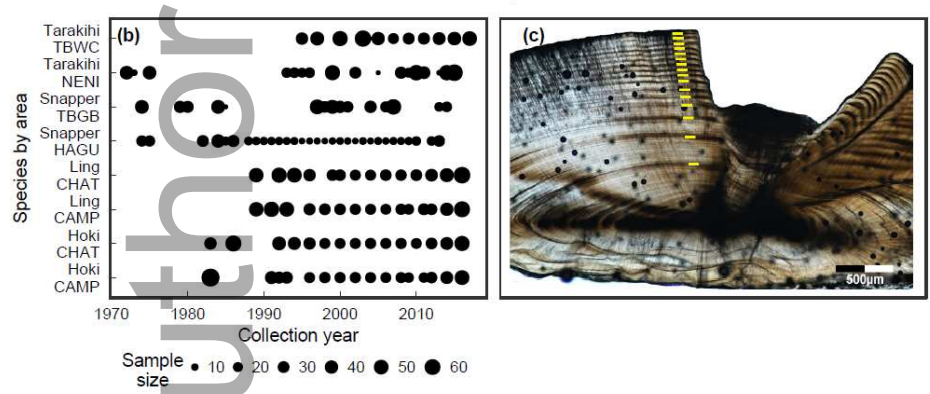
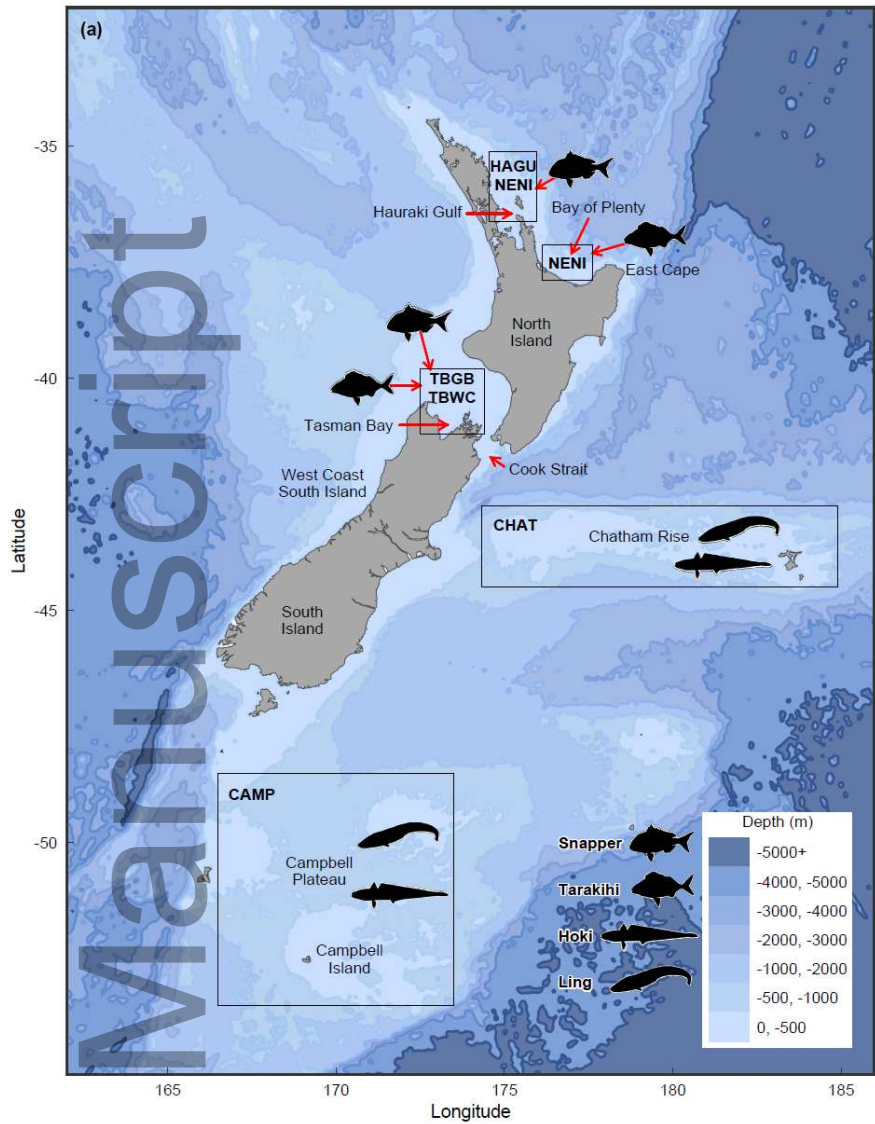
402 **Table 4:** Results of environmentally driven spatial synchrony model selection for four New
 403 Zealand fishes. Shown are the top three models (based on AICc) for each species. The best
 404 supported model for each species has a ΔAICc of 0 and is highlighted in bold. df = degrees of
 405 freedom.

Species and model	df	Log likelihood	AICc	ΔAICc
Snapper				
$\text{IPO}_{\text{ann}} + \text{SSB}_{\text{ann}}$	5	52.659	-93	0
SSB_{ann}	4	51.008	-92.5	0.48
$\text{SSB}_{\text{ann}} + \text{TrenZ1}_{\text{ann}}$	5	52.336	-92.4	0.65
Intercept only	3	46.694	-86.5	6.5
Tarakihi				
IPO_{ann}	4	36.993	-63.8	0
$\text{IPO}_{\text{ann}} * \text{SSB}_{\text{ann}}$	6	40.446	-63.6	0.12
$\text{IPO}_{\text{ann}} * \text{FP}_{\text{ann}}$	6	39.666	-62.1	1.68
Intercept only	3	30.632	-54.3	9.5
Hoki				
FP_{ann}	4	36.993	-63.8	0
$\text{IPO}_{\text{ann}} * \text{SSB}_{\text{ann}}$	6	40.446	-63.6	0.12
$\text{IPO}_{\text{ann}} * \text{FP}_{\text{ann}}$	6	39.666	-62.1	1.68
Intercept only	3	32.002	-56.7	7.0
Ling				
$\text{IPO}_{\text{ann}} * \text{SSB}_{\text{ann}}$	6	44.881	-72.5	0
$\text{IPO}_{\text{ann}} * \text{FP}_{\text{ann}}$	6	42.447	-67.6	4.87
$\text{SOI}_{\text{ann}} * \text{FP}_{\text{ann}}$	6	36.923	-56.6	15.92
Intercept only	3	28.998	-50.7	21.8

406

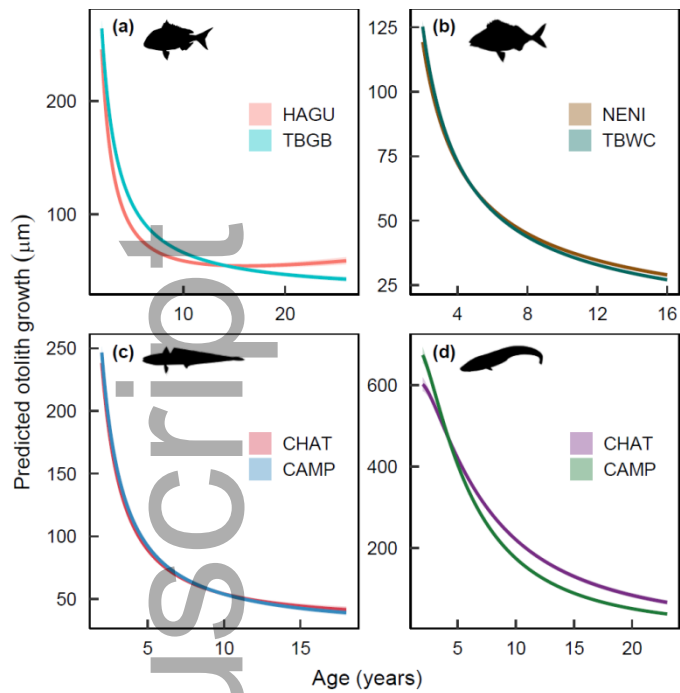
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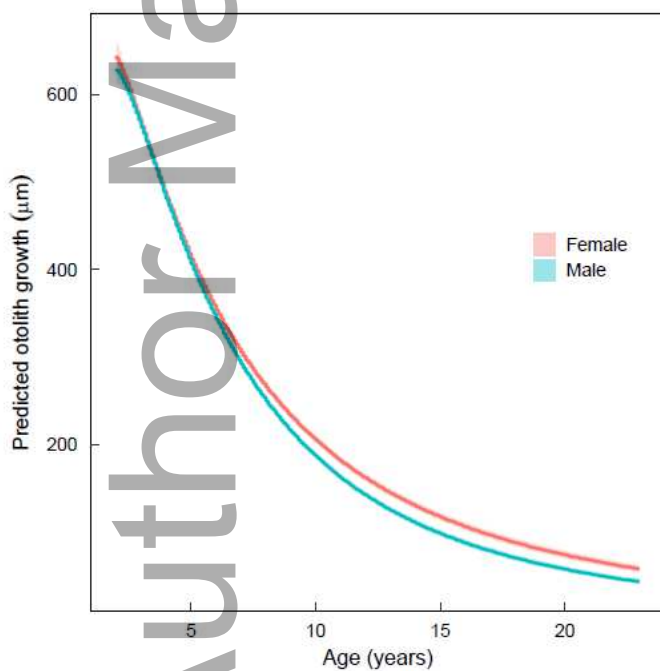
409

410 **Figure 1:** (a) Map of New Zealand showing fish sample locations; (b) graphical summary of
 411 otolith samples for each species by area combination used in this study; and (c) transverse
 412 section of a 16+ year old snapper's sagittal otolith, illustrating growth increments along the
 413 primordium to ventro-sulcal ageing axis.



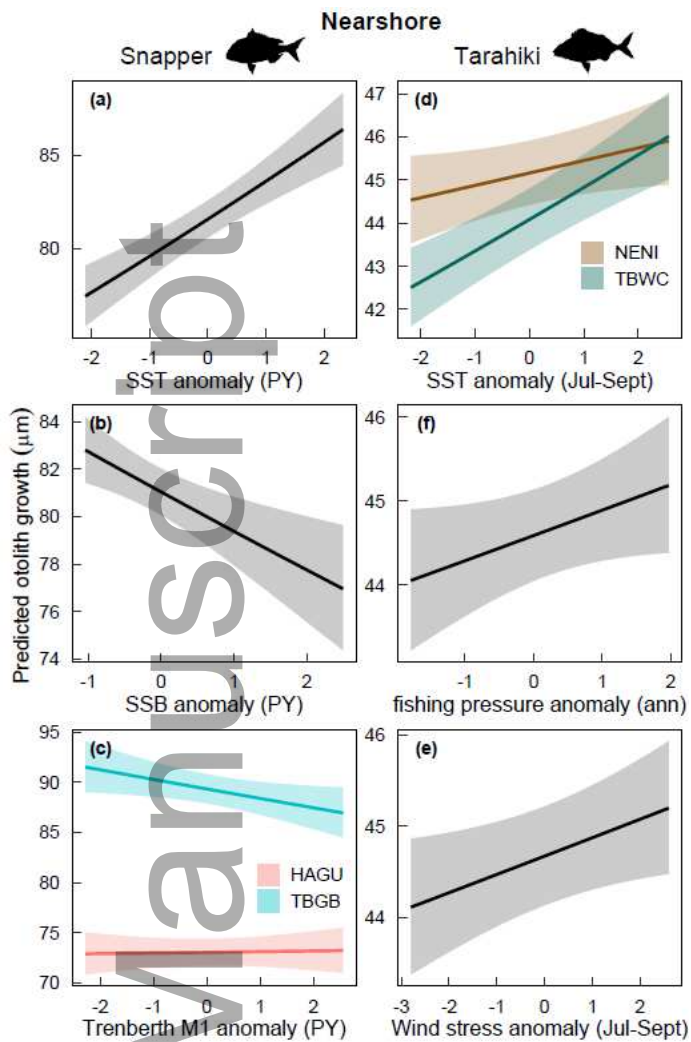
414

415 **Figure 2:** Age-dependent annual otolith growth (μm , $\pm 95\%$ CI) for each sample area in (a)
 416 snapper, (b) tarakihi, (c) hoki, and (d) ling.



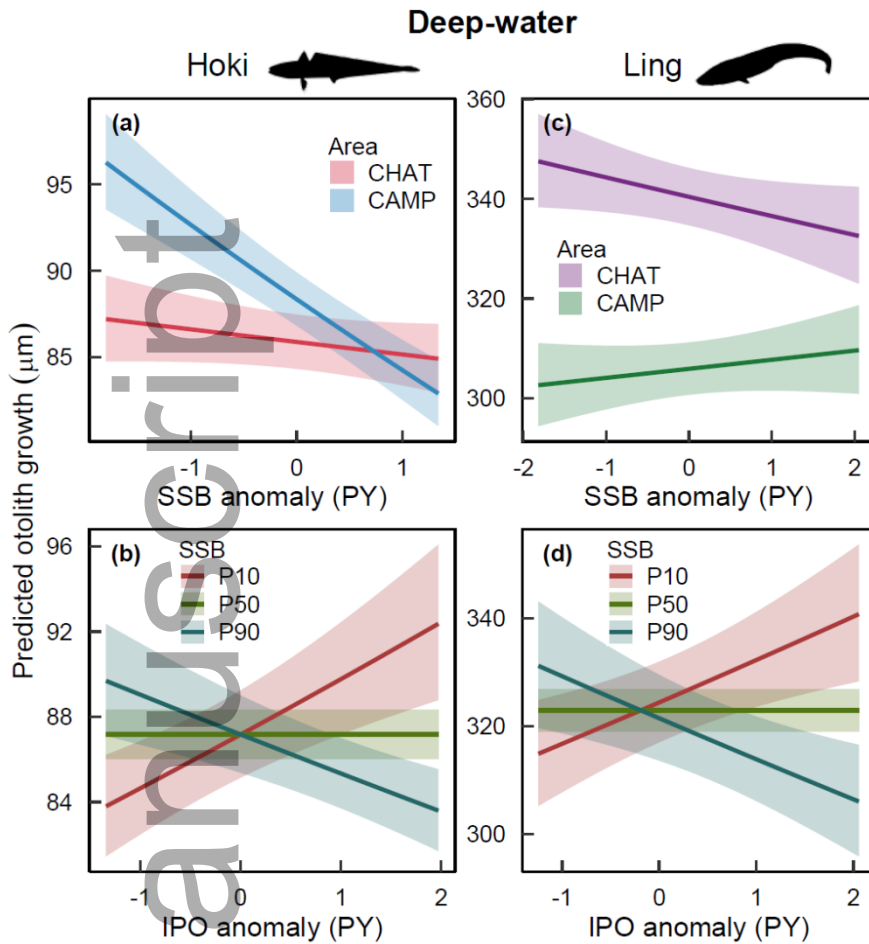
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418 **Figure 3:** Age-dependent annual otolith growth (μm , $\pm 95\%$ CI) for female and male ling.



419

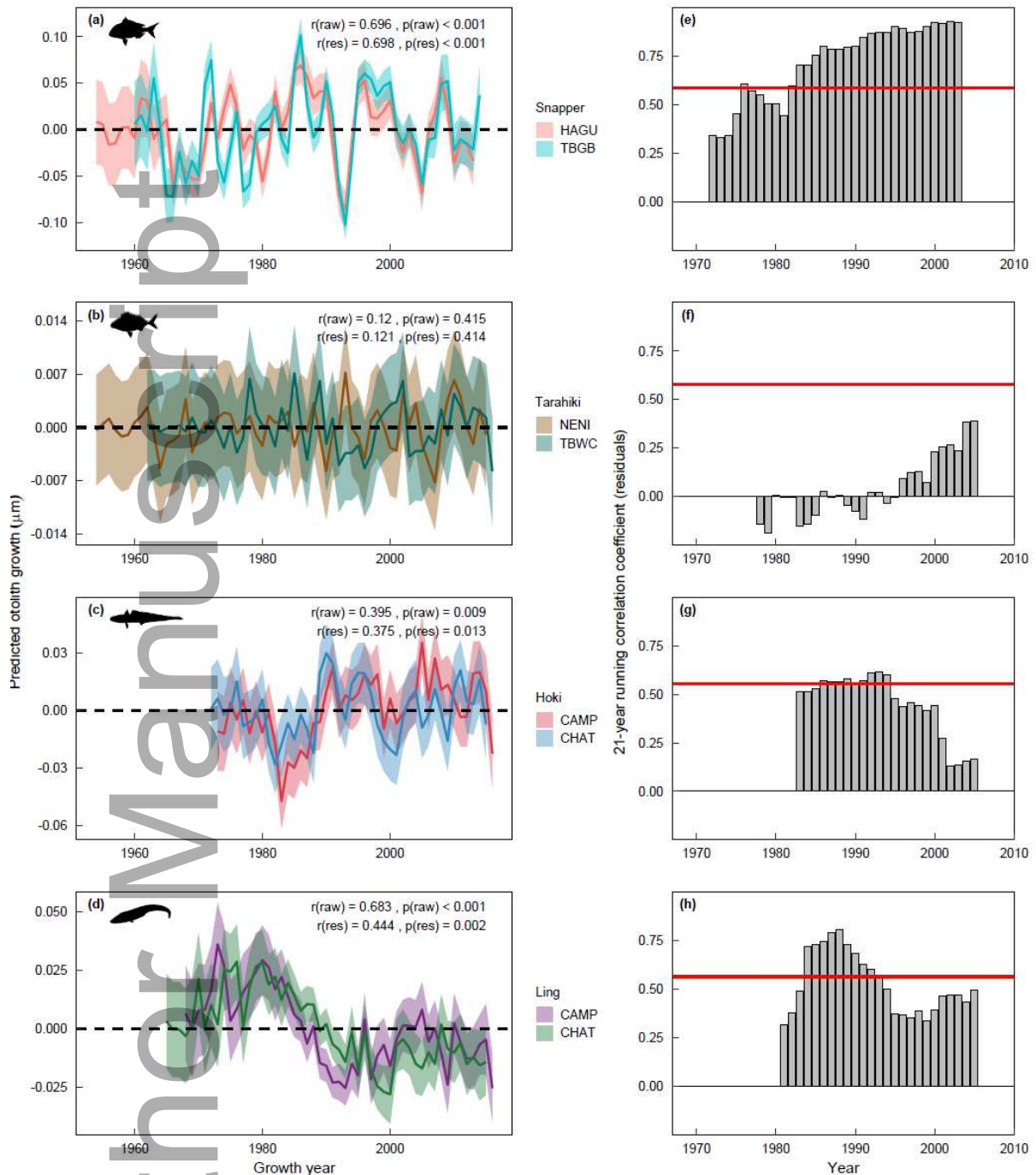
420 **Figure 4:** Annual growth variation (μm of otolith growth, $\pm 95\%$ CI) of nearshore species as a
 421 function of extrinsic factors. (a) Snapper growth response to the previous year's SST, (b) SSB,
 422 and (c) area-specific snapper growth responses to the previous year's meridional winds. (d)
 423 Tarahiki area-specific growth response to July–September SST, (e) fishing pressure, and (f)
 424 July–September wind stress. PY: previous year; SST: sea surface temperature; SSB: spawning
 425 stock biomass.



426

427 **Figure 5:** Annual growth variation (μm of otolith growth, $\pm 95\%$ CI) of deep-water species as
 428 a function of extrinsic factors. (a) Area-specific hoki growth response to the previous year's
 429 SSB, and (b) the interaction between the previous year's SSB and IPO. (c) Area-specific ling
 430 growth response to the previous year's SSB, and (d) the interaction between the previous
 431 year's SSB and IPO. PY: previous year; SSB: spawning stock biomass; IPO: Interdecadal Pacific
 432 Oscillation. Note: P10, P50 and P90 refer to the 10% 50% and 90% percentile value of
 433 observed SSB.

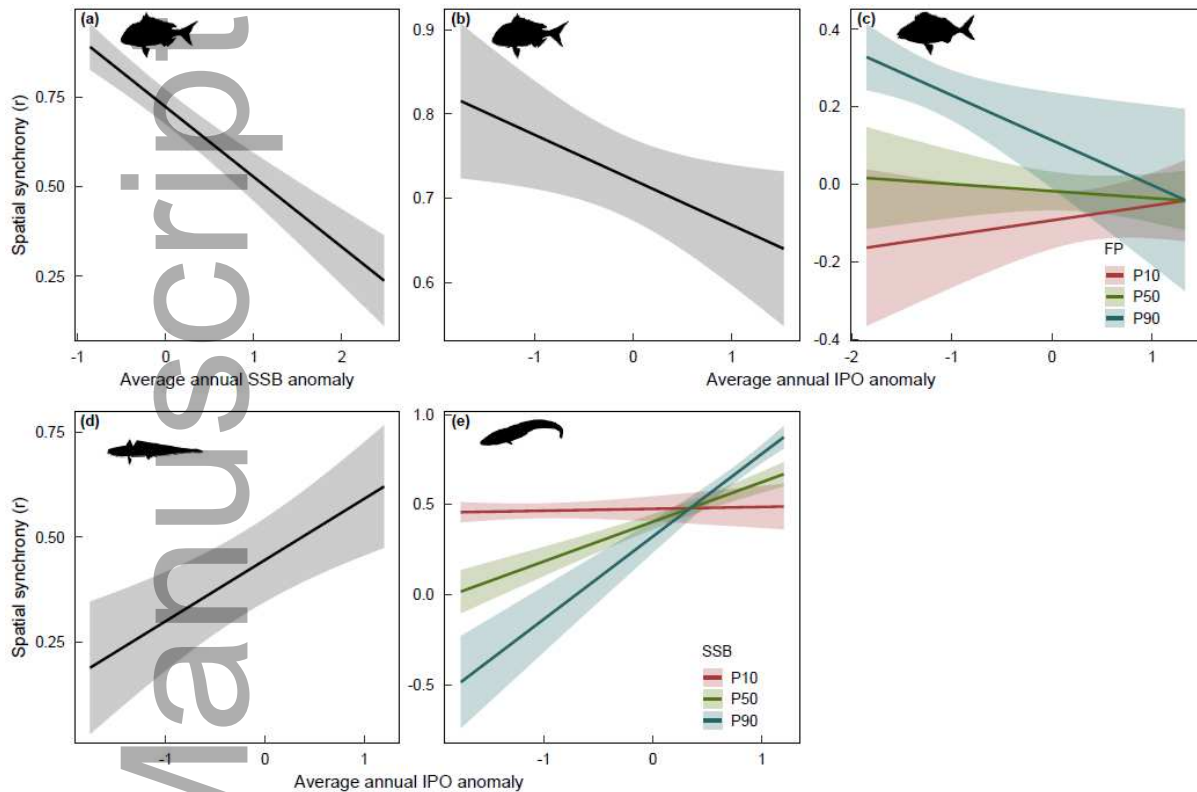
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435

436 **Figure 6:** (a–d) Predicted area-specific annual growth biochronologies (μm of otolith growth,
 437 ± standard error) with raw and residual (temporal trend and auto-correlation removed; see
 438 text) within species zero-lag cross correlations for snapper, tarakihi, hoki, and ling. Dotted
 439 horizontal lines represent long-term mean growth for each species (model fixed effect
 440 intercept). (e–h) 21-year within-species running correlations of annual growth residuals (see
 441 text; bars), plotted on the central year of each window. The red horizontal line represents
 442 the threshold above which running correlations are significant at the 0.05 level. Note, each

443 species' running correlations are based on a subset of years for which annual sample depth
444 was at least five increment measurements (snapper: 1962-2013; tarakihi: 1968-2015; hoki:
445 1973-2015; ling: 1971-2015).



446

447 **Figure 7:** Predicted relationships ($\pm 95\%$ CI) between estimates of time-varying spatial
448 synchrony (21-year running correlation r) and 21-year rolling means of annual extrinsic
449 factors. Spatial synchrony responses of: snapper to (a) SSB and (b) IPO; (c) tarakihi to IPO
450 and fishing pressure; (d) hoki to IPO; and (e) ling to IPO and SSB. SSB: spawning stock
451 biomass; IPO: Interdecadal Pacific Oscillation; FP: fishing pressure. Note: P10, P50 and P90
452 refer to the 10% 50% and 90% percentile value of observed FP and SSB.

453 Discussion

454 Natural and anthropogenically driven changes to the marine environment, coupled with
455 harvest-related shifts in population size, additively and synergistically influenced the degree
456 of average individual growth and population-level spatial synchrony across our four study
457 species. Within stocks, surviving individuals have benefited from historical and current
458 fishing practices as it released them from regulating density-dependent processes.
459 Conversely, fishing-induced abundance reductions have increased stock synchrony and thus

460 potentially increased the vulnerability of species to environmental change. The strong
461 environmental signals in our data indicate that individual growth will initially increase in a
462 warmer future and forthcoming positive IPO phase. However, the implications for this
463 change on fishery productivity remain less clear due to the contrasting effects of fishing at
464 the within and among stock scales.

465 *Intrinsic drivers of annual growth*

466 Spatial differences in age-dependent growth are likely to be driven primarily by
467 environmental factors, given the lack of any marked genetic heterogeneity across our
468 putative stocks (Fisheries New Zealand, 2019). Previous work using size-at-age data has
469 identified differences between male and female age-dependent growth for ling and hoki
470 (females attain larger size: Horn, 2005, Horn, 1993, Horn & Sullivan, 1996), but not snapper
471 and tarakihi (Walsh *et al.*, 2016, Walsh *et al.*, 2014). Our increment-based results generally
472 conform with these expectations, although the lack of detectable sex-specific differences in
473 hoki could reflect the relatively small differences in length-at-age (Horn & Sullivan, 1996).

474 Natural and artificial selection both impact the expression of an individual's average lifetime
475 growth rate. For our inshore species (snapper and tarakihi), individuals caught at younger
476 ages had, on average, grown at a faster rate than fish caught when older. We believe that
477 these patterns are representative of fishing selectivity, whereby faster growing individuals
478 within a cohort become vulnerable to capture at a younger age, or overall increases in
479 mortality risk are selecting for earlier maturation and thus faster growth (Audzijonyte *et al.*,
480 2016). Conversely, relatively more slow-growing ling were captured at younger ages during
481 the trawl surveys. Such a pattern is suggestive of size selective natural mortality (Sogard,
482 1997) favouring faster growing individuals to live longer (e.g. Tanner *et al.*, 2019).

483 Regardless of the cause, our modelling approach ensured that we could identify and
484 account for any ecological, evolutionary, or sample regime-dependent biases in annual
485 growth variation (Morrongiello *et al.*, 2012).

486 *Extrinsic drivers of annual growth*

487 The internal body temperature of ectotherms like fish largely depends on environmental
488 temperature. This means that key biological processes such as metabolic rate, digestion

489 rate, muscle activity and reproductive allocation- all linked to somatic growth rate
490 (Audzijonyte *et al.*, 2019)- can be directly regulated by water temperature. Ectotherm
491 growth can also be indirectly affected by temperature, such as when warming causes
492 changes in food quality or quantity or alters interspecific interactions (Behrenfeld *et al.*,
493 2006, Kordas *et al.*, 2011). Snapper and tarakihi share a similar diet (Table S1; Stevens *et al.*,
494 2011) and their annual growths were positively related to SST. Tarakihi growth was
495 correlated with winter SST (stronger in the cooler TBWC population) and we believe this is
496 indicative of a direct temperature effect. Whilst empirical studies on tarakihi thermal
497 tolerance have not been performed, inference can be drawn from their spatial distribution
498 (Stuart-Smith *et al.*, 2015). Tarakihi commonly experience annual temperatures ranging
499 from 12.7°C to 20.3°C (10th and 90th percentile; Kaschner *et al.*, 2019). Winter
500 temperatures in our more poleward (TBWC) and equatorial (NENI) populations ranged from
501 12.6-14.2°C and 14.9-16.8°C respectively. A greater growth response to warming would
502 therefore be expected in TBWC as conditions move toward a thermal optimum whereas
503 NENI fish will be closer to this point already.

504 Snapper growth was correlated with annual SST from the previous growth year. We
505 interpret this lagged SST response as an indirect effect of warming that promoted the
506 recruitment or survival of snapper food (crabs, shellfish, echinoderms and polychaetes).
507 Such a temperature-induced productivity pulse may take a year to become available to
508 adult snapper (e.g. Frederiksen *et al.*, 2006, Nguyen *et al.*, 2015). Previous work has shown
509 that snapper growth elsewhere in its range is positively related to SST up to an optimal of c.
510 19–20°C, whereafter growth declines with additional warming (Martino *et al.*, 2019,
511 Wakefield *et al.*, 2016). HAGU snapper and NENI tarakihi populations are already
512 experiencing temperatures close to their potential thermal optimum (annual maximum in
513 our study of 19°C). Depending on the relative importance of direct and indirect mechanisms,
514 projected increases in SST may initially cause an increase in growth for both species in their
515 more equatorial populations, followed by a decline (Morrongiello & Thresher, 2015,
516 Neuheimer *et al.*, 2011). Warming will continue to promote the somatic growth of fish in
517 more poleward, cooler populations.

518 Wind direction and wind strength play an important role in determining water column
519 mixing, sea surface temperature, and associated rates of primary productivity (e.g. Longdill

520 *et al.*, 2008, Lozier *et al.*, 2011, Schofield *et al.*, 2018). Annual snapper growth was generally
521 higher, especially in the south, following years of weaker meridional (northly) winds. These
522 conditions promote stratification and may lead to warmer, more favourable temperatures
523 for snapper prey (Chang *et al.*, 2003). Tarakihi growth was higher in years where the winter
524 had greater offshore wind stress. Offshore winds aid mixing which promotes primary
525 productivity in New Zealand coastal waters (Longdill *et al.*, 2008), which in turn may support
526 tarakihi food (Stevens *et al.*, 2011). There has been a recent increase in Southern
527 Hemisphere mid latitude westerly winds attributable to climate change (Thompson &
528 Solomon, 2002), and this trend is expected to continue over the next 50 years (Mullan *et al.*,
529 2001). We speculate that these stronger zonal winds will promote primary productivity in
530 coastal waters (*sensu* Sydeman *et al.*, 2014) and lead to an increase in snapper and tarakihi
531 growth rates.

532 The effects of spawning stock biomass and fishing pressure on growth were apparent for all
533 species. The patterns observed were consistent with individual fish growing faster when
534 biomass was lower, suggesting a release from density-dependent growth control (Lorenzen
535 & Enberg, 2002, Morrongiello *et al.*, 2019), and/or reflective of size selective harvest and
536 increased mortality driving the expression of elevated juvenile growth (Audzijonyte *et al.*,
537 2016). Contrasting spatial differences in the effect of spawning stock biomass were detected
538 for hoki and ling, with the growth of CAMP hoki (larger absolute biomass) and CHAT ling
539 (smaller absolute biomass) more sensitive to changes in this variable. It is likely that relative
540 changes in biomass or habitat alteration associated with trawl fishing, not absolute changes
541 in stock size, drives the observed growth variation. Interestingly, the growth of snapper and
542 tarakihi experienced only additive effects of fishing and environmental variation despite
543 both species being heavily harvested and there being stock-specific differences in fishing
544 pressure (Figure S1; Fisheries New Zealand, 2019). These results suggest that regional
545 environmental variation has always had a strong influence on the productivity of New
546 Zealand's nearshore fish stocks, regardless of whether they are fished or not (e.g. Gillanders
547 *et al.*, 2012).

548 The growth of deep-water species is often driven by factors and at time scales different to
549 those operating at the surface (Nguyen *et al.*, 2015, Tanner *et al.*, 2020). These differences
550 arise because deep-water environments are relatively constant and the assimilation of

551 surface-derived primary productivity into deep-water food webs takes time (Dickey, 2001,
552 Morales-Nin & Panfili, 2005). The IPO is a complex, multi-decadal and Pacific-wide, climate
553 phenomena. During positive phases of the IPO, the south Pacific convergence zone (SPCZ)
554 shifts southwest and the waters around New Zealand become cooler than average, south-
555 westerly winds become stronger, the weather stormier, and waves bigger (Folland *et al.*,
556 2002, Godoi *et al.*, 2018, Salinger *et al.*, 2001). These weather conditions can result in
557 deeper mixed layers (Carranza *et al.*, 2018) and potentially a greater subsidy of surface
558 productivity to the deep (Pinkerton *et al.*, 2019). We detected a synergistic effect of lagged
559 IPO and spawning stock biomass on the annual growth of hoki and ling. The importance of
560 lagged IPO could arise if its multi-decadal cycle of influence at the surface (Henley *et al.*,
561 2015) allows this ocean-wide phenomenon to be more strongly integrated into higher
562 trophic levels of the food web at depth (*sensu* Di Lorenzo & Ohman, 2013). Positive IPO
563 phases are characterised by favourable environmental conditions for growth, and this
564 benefit is likely maximised when intraspecific competition (related to biomass) is low.
565 Therefore, our results suggest that density dependence can limit the capacity of individuals
566 to respond to favourable conditions. Under current fisheries management practice, future
567 SSB will likely remain around the 40% reference point while IPO is forecast to enter a
568 positive phase (Henley *et al.*, 2017, Meehl *et al.*, 2016). These conditions indicate that hoki
569 and ling annual growth rate will increase over the next decade.

570 *Spatial synchrony in annual growth*

571 Spatial synchrony in populations is widely believed to be driven by three non-exclusive
572 natural processes: rates of individual dispersal, strength of environmental factors, and
573 trophic interactions (Liebhold *et al.*, 2004). The strong link between all four species' spatial
574 synchrony time series and IPO is intuitive given the IPO has a periodicity of 10-30 years
575 (Parker *et al.*, 2007) and our running correlations were calculated over a 21 year window.
576 The increasing spatial synchrony in nearshore species during the recent negative IPO phase
577 could be, in part, related to the associated warming of coastal waters (see above).
578 Conversely, predicted reductions in region-wide surface-to-deep trophic subsidies during
579 the same negative IPO phases may lead to a dominance of local scale processes on growth
580 (e.g. habitat complexity; Brown *et al.*, 2019) and thus the observed reduction in hoki and
581 ling spatial synchrony.

582 Temporal patterns in the spatial synchrony of our stock somatic growth rates is most likely
583 driven by unifying environmental conditions (i.e. 'Moran Effect'), with this sensitivity further
584 amplified by fishing activity in some species. For tarakihi and ling, increased fishing pressure
585 (and its correlate, reduced spawning stock biomass) amplified the effect of IPO on spatial
586 synchrony. For tarakihi, this occurred despite both stocks experiencing very different fishing
587 mortalities, albeit with a similar trend (Figure S1). Snapper synchrony was affected by fishing
588 in a similar, but solely, additive way. Fishing reduces population density (Lorenzen &
589 Enberg, 2002) and truncates a population's demographic structure (Barnett *et al.*, 2017).
590 Both these fishing-induced mechanisms, reduced density dependence and demographic
591 truncation, can impact on an individual or population's environmental sensitivity and hence
592 its likelihood of spatial synchrony. Less competition for resources can be beneficial and
593 allow surviving individuals to have greater capacity to track favourable conditions
594 (Morrongiello *et al.*, 2019). Conversely, a loss of phenotypic diversity can erode the capacity
595 of populations to buffer against environmental change (Gamelon *et al.*, 2019).

596 Despite evidence of within-species spatial synchrony in growth, correlations between
597 species in the same location were generally weak. This result contrasts those from a recent
598 study that detected among-species growth synchrony in the deep-water fish from the
599 Azores (Tanner *et al.*, 2020). Our analyses of extrinsic drivers of growth showed strong
600 similarities in the importance of environmental variables for nearshore and deep-water
601 species. It is possible then that different fishing histories have induced varying levels of
602 impact on each species' demographic structure (Fisheries New Zealand, 2019), and in turn
603 reduced apparent among-species spatial synchrony. Previous work has found that dispersal
604 can induce spatial synchrony among fish population abundances over distances
605 commensurate with ecologically relevant dispersal (Cheal *et al.*, 2007). In our case, fish
606 populations were between 400-600nm apart and so we suspect that not enough fish
607 regularly move between populations to homogenise growth.

608 *Conclusions*

609 Overall, environmental change and fishing activity are having an important impact on the
610 growth rate, and thus productivity of four of New Zealand's most important nearshore and
611 deep-water fisheries. Historical and current fishing practices increased stock synchrony

612 (Moran Effect) and thus potentially the vulnerability of species to environmental change.
613 Conversely, within stocks individuals benefited from fishing as it released them from
614 regulating density-dependent processes. Nearshore species were most strongly linked to
615 regional-scale variables, while deep-water species responded to basin-scale drivers. The
616 functional relationships identified here between environmental parameters and a proxy for
617 somatic growth can be used to increase the reality of future stock assessments. The
618 inclusion of environmentally driven growth patterns could increase the predictive power of
619 models and will also allow for fisheries scenario testing under a range of warming futures.

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631 **Data availability statement**

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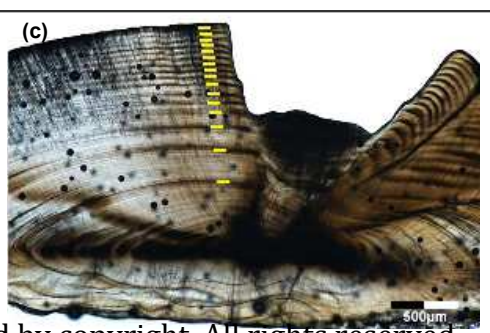
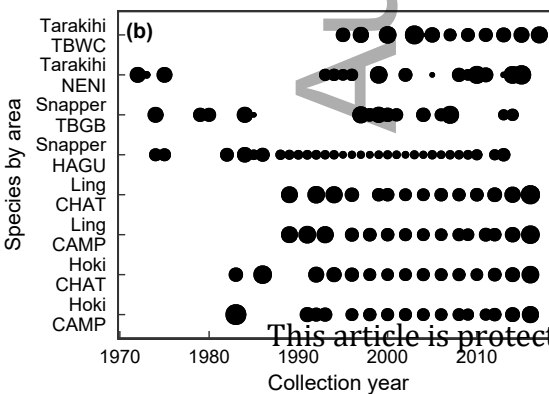
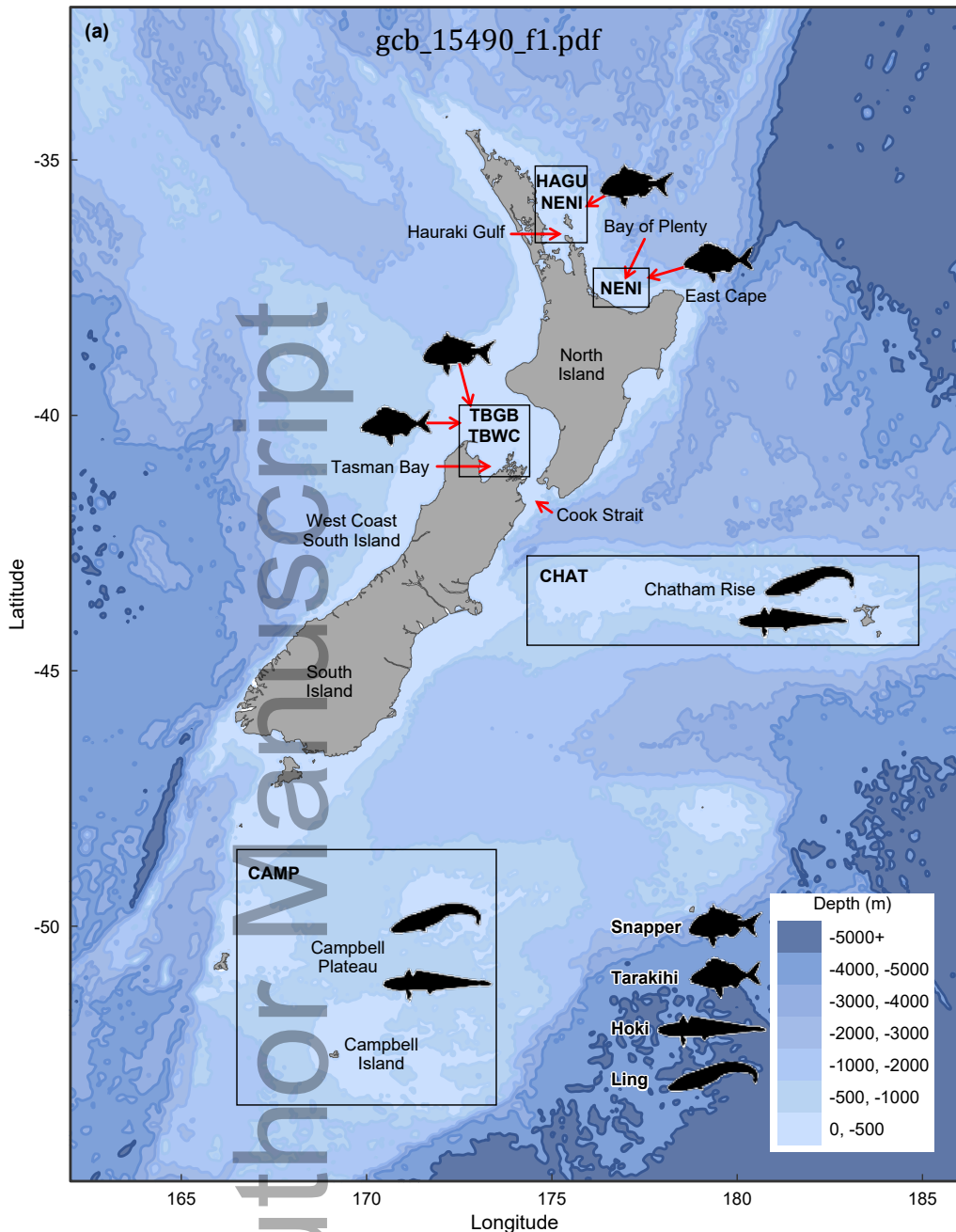
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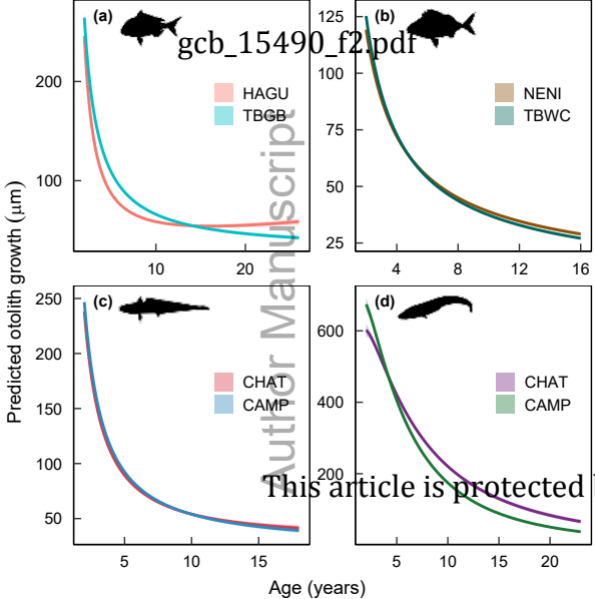
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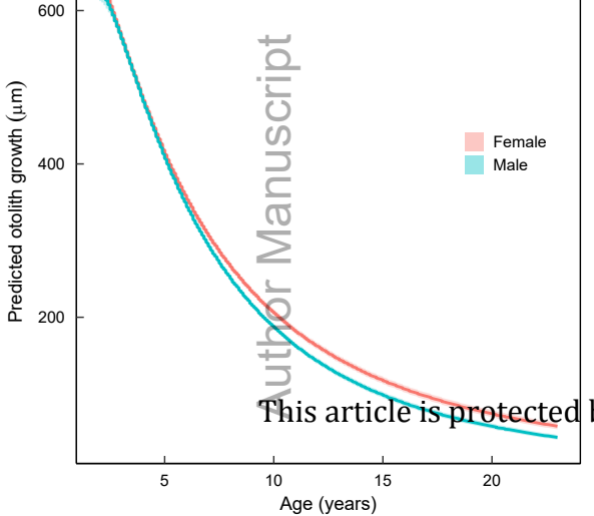
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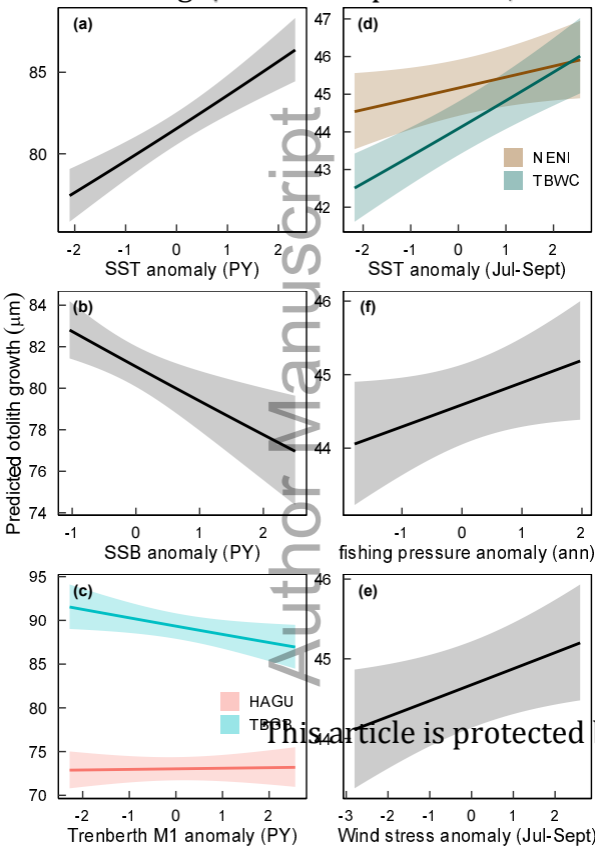
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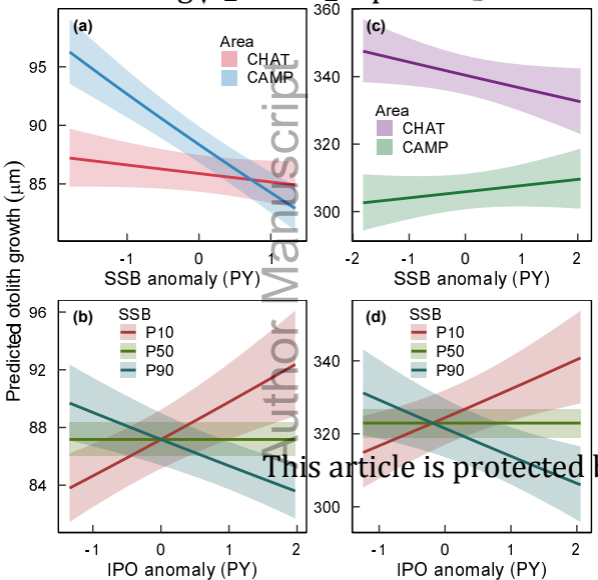


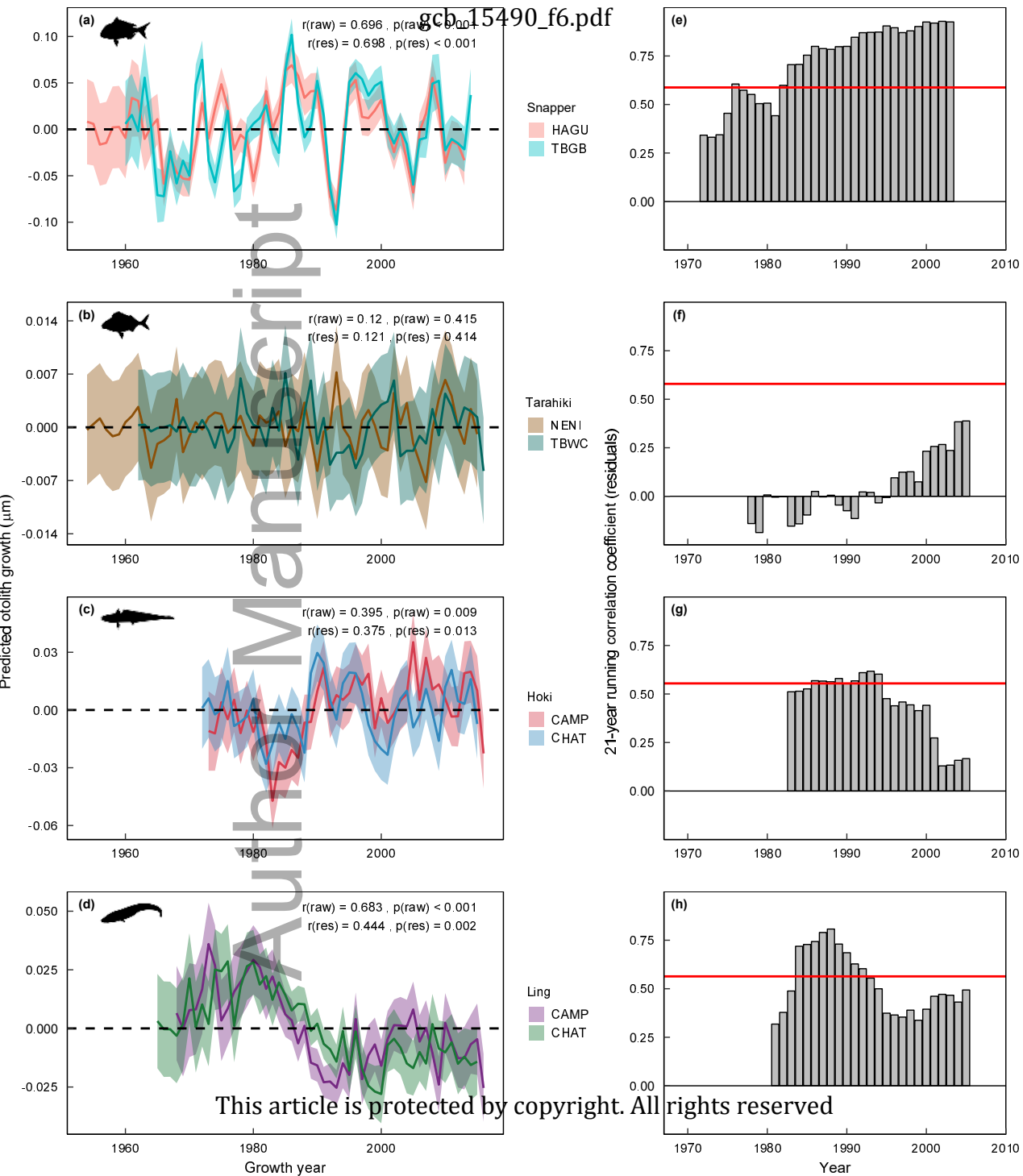


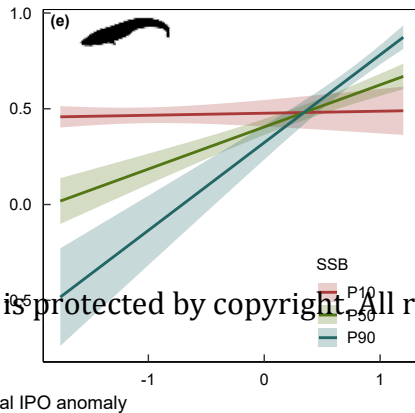
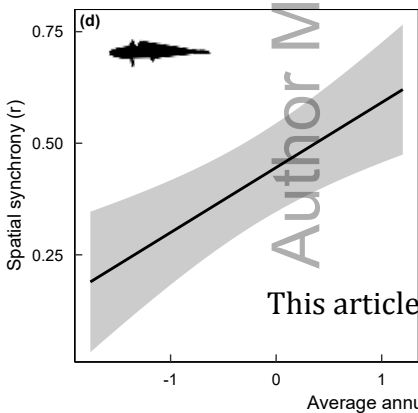
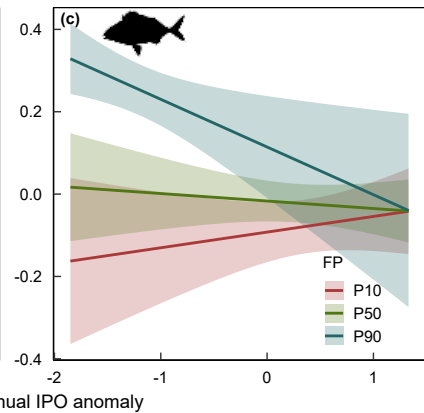
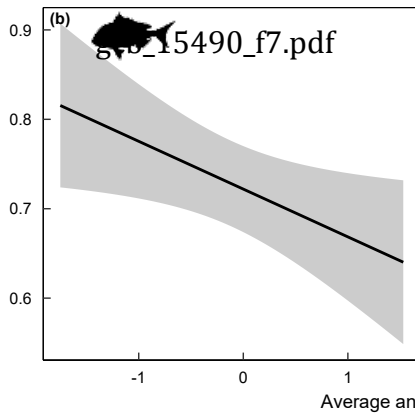
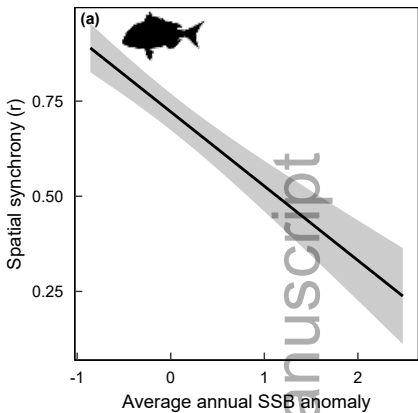
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