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Costs and Benefits of the Melbourne Mobile Stroke Unit Compared With Standard Ambulance: Causal Analysis Using Observational Linked Data

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BACKGROUND: Evidence of the cost implications and health outcomes associated with the use of mobile stroke units (MSU) is required to support their utilization. We aimed to evaluate the causal effect of the use of an MSU compared with a standard ambulance on hospitalization costs and 90- to 180-day health outcomes.

METHODS: Causal effect estimation was performed using patient-level data from a cohort of patients with stroke in 2018 identified from the Australian Stroke Clinical Registry (Victoria) and Melbourne MSU. These data were linked to Ambulance Victoria and government-held administrative data sets. In total, linked data from 8657 patients were available. Propensity score matching was used to define comparator groups within a target trial framework. Costs included emergency department and hospital admission costs in the first 180 days after stroke. Multivariable regression analyses of the matched data were used to compare costs and outcomes (mortality and modified Rankin Scale) between MSU and standard ambulance groups.

RESULTS: The target trial sample included 96 patients transported by the MSU (intervention) and 198 patients transported by standard ambulance services (control). Of these, the mean age was 76 years and 157 (53%) were men. A greater proportion of patients received mechanical thrombectomy in the intervention group than the control group (40% versus 23%; $P < 0.001$). The adjusted hospital costs were \$17 949 greater in the intervention group than the control group (95% CI, \$4682–\$31 214; $P = 0.01$). Patients in intervention group doubled the odds of achieving nondisability (modified Rankin Scale scores of 0–1, adjusted odds ratio of 2.11 [95% CI, 1.07–4.18]) and halved the mortality rate (adjusted hazard ratio, 0.53 [95% CI, 0.32–0.86]) within 90 to 180 days poststroke compared with the control group.

CONCLUSIONS: There are important cost implications and improved outcomes from using the MSU that are likely related to increased provision of reperfusion therapy.

GRAPHIC ABSTRACT: A [graphic abstract](#) is available for this article.

Key Words: ambulances ■ cost-effectiveness analysis ■ information storage and retrieval ■ outcome assessment, health care ■ registries ■ routinely collected data ■ stroke

Stroke is a major cause of death and disability,¹ but improving the time to treatment after a stroke increases the likelihood of better health outcomes.² In Australia, only 1 in 3 patients are treated within

the benchmark of 60 minutes from the time of hospital arrival.³ Delays in accessing hospitals after stroke symptoms are common and influenced by patient factors (do not seek urgent assistance) and system factors

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Nonstandard Abbreviations and Acronyms

AuSCR	Australian Stroke Clinical Registry
IQR	interquartile range
IVT	intravenous thrombolysis
mRS	modified Rankin Scale
MSU	mobile stroke unit
MT	mechanical thrombectomy
NIHSS	National Institutes of Health Stroke Scale

(resources, lack of a coordinated and streamlined process including hospital prenotification by paramedics). Bringing treatment directly to patients with suspected stroke via a mobile stroke unit (MSU) is a solution for reducing timelines to treatment, but investment to service an MSU can be substantial.^{4–7}

MSUs are modified ambulances equipped with a brain imaging system for stroke diagnosis, staffed with skilled stroke clinicians (advanced practice nurses and radiographers) to provide medications, point-of-care laboratory testing, or access specialist medical support via telemedicine.^{8–10} MSUs are being used in some high- and low-income countries with growing evidence that their use leads to reductions in the time-to-treatment, improvements in prehospital triage, and better clinical outcomes.^{11–16} Furthermore, there is evidence from economic modeling that MSU services are cost-effective and can reduce downstream stroke-related costs with variation according to the volume of patients treated within the MSU.^{7,17–19}

The first Australian MSU was launched in Melbourne for use in November 2017.^{9,10} Large investments are required to have these services,^{6,7} and an economic evaluation was planned to support the expansion or duplication of the MSU service to other ambulance services within Victoria or other similar settings. Previous estimates of the cost-effectiveness relied on modeling using preliminary activity data and data published in the peer-reviewed literature.¹⁰ In this study, we analyzed patient-level data to evaluate the costs and benefits of the use of MSU (intervention) compared with standard ambulance (control) for the treatment of stroke. The findings will be used to inform future economic modeling of the costs and benefits of the Melbourne MSU over patients' lifetimes.

METHODS

Data Availability and Ethics

Due to ethical and legal restrictions, personal-level data from this study cannot be shared. However, certain aggregated data and coding that support the findings of this study are available upon reasonable request from the corresponding author, after approval from the relevant data custodians. Human Research Ethics Committee approval was obtained

from Melbourne Health (HREC/17/MH/375) and Monash University (20158). Approvals for linkage and release of deidentified data were also obtained from relevant data custodians, including Ambulance Victoria, Melbourne Mobile Stroke Unit Steering Committee, Victorian Department of Health (Data Linkage Unit), the Australian Stroke Clinical Registry (AuSCR) Research Task Group, and the AuSCR Steering Committee.

Study Design

Ideally, a randomized controlled trial design provides confidence in making causal inferences. However, in the real world, for comparative effectiveness studies, alternative methods must be used since it is not always ethical or feasible to conduct randomized controlled trials, particularly when assessing the adoption of new models of health care against usual care.²⁰ In this study, to ensure the robustness of our findings, the causal effect of treatment in the MSU on costs and health outcomes was estimated by conducting a causal analysis of observational linked data using an emulated target trial methodology (target trial framework).²¹ The target trial framework provides a counterfactual theory for comparing the effects of treatment strategies and helps to avoid methodological pitfalls of observational studies using a structured process.²¹ This study is reported in accordance with both the STROBE (Strengthening the Reporting of Observational Studies in Epidemiology) and CHEERS (Consolidated Health Economic Evaluation Reporting Standards 2022) checklists (Supplemental Material; Appendices S1 and S2).^{22,23}

Study Setting

Seven Victorian hospitals were capable of providing stroke unit care in 2018, and participants had access to both MSU and standard ambulances. The details of the human and material resources needed to establish the Melbourne MSU service and the operational system have been described previously.^{9,10} However, the start-up and operational costs for MSU were not considered in this study.

Data Sources

Patient-level data were linked and merged from different sources for this study, after approval from all data custodians. The data linkage was undertaken by the data custodians, and deidentified data sets were then transferred to Monash University for merging using a unique project identification number.²⁴ Only authorized authors J.K. and M.M.B. accessed the merged data set as they did not have the ability to reidentify any cases from the primary data sources. Data collected by the Melbourne MSU and the AuSCR were linked to Ambulance Victoria data and government-held administrative data sets (emergency episodes, admitted episodes, clinical costing, and death registry). Where patients had >1 record in the Melbourne MSU or AuSCR data sets, only the first instant the patient appeared in the data set was considered for analysis. Participants who were transferred from another hospital and had in-hospital strokes, aged <18 years and patients who attended out of the catchment areas were excluded (Figure S1).

Melbourne MSU

The Melbourne MSU team systematically collected data on demographics, clinical characteristics, and treatment, including the time metrics for dispatch, arrival, and treatment.⁹ In addition,

outpatient follow-up was conducted, including an assessment of the modified Rankin Scale (mRS) score at 90 to 180 days after stroke during a clinic visit with a stroke specialist; otherwise, this was assessed by nurses from medical records or via phone.

Australian Stroke Clinical Registry

The AuSCR is used in participating hospitals to monitor the quality of stroke care in public hospitals (eg, prioritized evidence-based treatments including reperfusion therapies, care in stroke units, and discharge medications, and survivors are followed up between 90 and 180 days to complete a self-reported health outcomes survey [including the modified Rankin Scale, mRS]).²⁵ In 2018, the median time to completion of follow-up for adult responders was 144 days after admission.²⁶ Each year the cohort is linked to the National Death Index to obtain information on the fact of death and cause of death made available from the Australian Institute of Health and Welfare.

Ambulance Victoria Data

Ambulance Victoria is the emergency medical services provider in the state of Victoria and captures data including demographic information and time metrics for dispatch, arrival, and treatment for patients who provided ambulance dispatch and transport.²⁴

Department of Health, Victoria

The Victorian Emergency Minimum Datasets collects deidentified data, including demographic, administrative, and clinical information, from emergency department presentations at Victorian public hospitals. The Victorian Admitted Episodes Datasets comprise data for admitted episodes of care occurring in Victorian hospitals, rehabilitation centers, extended care facilities, and day procedure centers.

Patient-level costs related to emergency department and hospital admissions were obtained from the Victorian Cost Data Collection. Patient-level data on presentations to hospitals are conducted at Victorian hospitals and are reported to the Victorian Cost Data Collection and, in turn, to the Independent Health and Aged Care Pricing Authority for funding of hospital services. From 2018 to 2019, 96% of admitted acute episodes had cost data. However, costs for 6 months do not exclude costs unrelated to stroke. Data on the fact of death and cause of death from the Victorian Death Index were also obtained.

Treatment Strategies

Patients treated in the Melbourne MSU (intervention group) were compared with standard ambulance transport to the hospital (control group).

Eligibility Criteria

To emulate randomized controlled trials, only those with a confirmed stroke (ie, ischemic or hemorrhagic stroke) in hospital discharge and those aged ≥ 40 years were included since there were differences in distributions of patient diagnoses (Table S1) and age between groups (Figure S2). Since patients with transient ischemic attack (TIA) by definition return to their usual functional status, they were excluded from the main analysis since MSU-specific interventions would have no opportunity to influence the mRS outcomes at 90 to 180 days. In addition, the Melbourne MSU will less often attend patients with TIA

than patients with stroke as evidenced by the greater imbalance of the proportion between the groups (Table S1). The control group was selected if transported during the days and times of MSU operation (Figure S3). Participants in the control group were limited to those presenting to a predefined set of hospitals that served by MSU. In addition, due to the lack of socioeconomic status information for all participants, we used the postcode of the patients' pick-up location and the hospital destination as a proxy measure. Those with missing values for the National Institutes of Health Stroke Scale (NIHSS) or mRS score at 90 to 180 days were excluded (Table S2).

Exposure Assignment

To ensure reliable comparisons between groups, we used a propensity-score matched causal effect estimation approach. We ensured the positivity assumption requirements (ie, all treatments of interest to be observed with nonzero [positive] probability in every patient subgroup).²⁷ Violation of these assumptions was indicated by nonoverlapping data. To ensure balance between groups given our nonrandomized design, we first profiled the MSU cohort and then matched it to standard ambulance users with similar characteristics (Figure S1). The matching was conducted using age, sex, stroke diagnosis, hospital destination, day of admission, and NIHSS. The standard mean difference of 10% was considered a meaningful imbalance in the baseline covariate.^{28–30}

Follow-Up Period

The participants were followed up from their first stroke event to the first 90 to 180 days to assess their disability and readmission status. The costs of hospital presentation were assessed within 180 days of stroke. Time-to-death was defined as the time from the stroke event to the date of death or the last follow-up time, whichever came first in the first 180 days of follow-up.

Outcomes

The primary outcome was the per patient hospitalization costs including emergency department and hospital admissions from stroke onset to 180 days. Costs of hospital presentations within 180 days of stroke were categorized as urgent and nonurgent emergency department presentations, acute admissions, and subacute admissions (rehabilitation, palliative care, and other subacute care). All costs were expressed in 2018 Australian dollars (AUD). The secondary outcomes were the nondisability rate, which was measured using the mRS^{31–33} score at 90 to 180 days, and mortality. Patients were categorized as nondisabled with mRS scores of 0 or 1 at 90 to 180 days. Further analysis was also done with the 3-tier disability scale at 90 to 180 days (ie, independent [0–2 mRS], ambulatory [0–3 mRS], and requires either constant care or deceased [a worse outcome; 5–6 mRS]).

Causal Contrasts of Interest

Our primary analysis was based on the intention-to-treat analysis via a comparison of cost and health outcome measures among patients transported by MSU or standard ambulance.

Statistical Analyses

The intervention and control groups were compared using Wilcoxon rank-sum test (Mann-Whitney U test) for continuous variables, while categorical variables were compared using χ^2 tests and logistic regression. Fisher exact test was used when the frequency had a cell count <5 .

Multiple imputation with chained equations was used to estimate costs for hospital presentations without cost data. Data on costs were missing for $<2\%$ of records related to the hospital records for presentation related to stroke (on event), and for $<15\%$ of records related to hospital presentations following the presentation related to stroke (after event). Data were imputed based on the patient's age and the primary diagnosis code (*International Classification of Diseases, 10th Revision*).

Generalized linear regression with gamma distributions and a log link was used for comparisons of hospital costs. Where possible, medians are presented because of the non-normal distribution. However, means were presented where adjusted generalized linear regressions were used. We used logistic regression for measures of nondisability and hospital admission and Cox proportional hazards regression for mortality. Secondary analysis of the utility-weighted mRS was conducted using linear regression.³⁴ All the multivariable regression analyses were adjusted for potential confounders (age, sex, stroke diagnosis, hospital destination, day of admission, and NIHSS of the patient). Statistical analyses were performed in Stata V.18 (STATA BE/18, StataCorp). A 2-tailed $P \leq 0.05$ was considered statistically significant.

RESULTS

Cohort Description

In total, linked data from 8657 patients were available (median age, 76 years; 4637 [54%] were men). Of these, 1143 (13.2%) were from the MSU and 7514 (86.8%) were from the standard ambulance group (Figure S1). The day of the week that the admission occurred was similar between groups (Figure S3). After applying exclusion criteria and propensity score matching (ie, using age, sex, stroke diagnosis, hospital destination, day of admission, and NIHSS of the patient as confounding factors; Figure S4), 294 patients were included in the causal inference analysis (intervention group, ie, transported by MSU [$n=96$] and control group, ie, standard ambulance users [$n=198$]). The sociodemographic characteristics of the intervention group and control group were similar (Table 1). The median age of the matched cohort was 79 years, 157 (53%) were men, 231 (79%) had ischemic stroke, and 63 (21%) had intracerebral hemorrhage.

Provision of Acute Stroke Care

Among patients with ischemic stroke, a greater proportion of patients in the intervention group arrived at the hospital within 4.5 hours (77.9% versus 56.3%;

Table 1. Sociodemographic and Other Characteristics of Stroke Survivors Transported by Ambulance Among Our Cohort Within the Target Trial, Victoria, 2018

Characteristics	Standard ambulance, n=198	MSU group, n=96	P value
Age, y; median (25th–75th percentile)	79.5 (69.1–87.5)	78.4 (66.1–86.0)	0.37
Male, n (%)	109 (55.1)	48 (50.0)	0.42
Diagnosis			
Ischemic, n (%)	161 (81.3)	70 (72.9)	0.10
Intracerebral hemorrhage, n (%)	37 (18.7)	26 (27.1)	
NIHSS, median (25th–75th percentile)	10 (4–18)	15 (7–21)	<0.001
Hospital destination*			
Hospital 1	39 (19.7)	9 (9.4)	
Hospital 2	35 (17.7)	15 (15.6)	
Hospital 3	13 (6.6)	2 (2.1)	
Hospital 4	26 (13.1)	5 (5.2)	<0.001
Hospital 5	6 (3.0)	3 (3.1)	
Hospital 6	73 (36.9)	60 (62.5)	
Hospital 7	6 (3.0)	2 (2.1)	
Day of the week			
Monday	36 (18.2)	25 (26.0)	
Tuesday	44 (22.2)	17 (17.7)	
Wednesday	45 (22.7)	22 (22.9)	0.59
Thursday	40 (20.2)	18 (18.8)	
Friday	33 (16.7)	14 (14.6)	

Values are presented as median (25th and 75th percentiles) and frequency (%). P values for nonparametric continuous variables were calculated according to Wilcoxon rank-sum (Mann-Whitney U test) test, while categorical variables calculated using χ^2 . Due to rounding, percentages may not add up to 100%. MSU indicates mobile stroke unit; and NIHSS, National Institutes of Health Stroke Scale.

*Hospitals considered in our analysis and had stroke unit care in 2018.

adjusted odds ratio, 2.93 [95% CI, 1.46–5.88]; $P=0.01$; Table 2; Figure S5). The proportion of patients provided intravenous thrombolysis (IVT) was greater in the intervention group than in the control group (67% versus 26%; adjusted odds ratio, 5.43 [95% CI, 2.80–10.53]; $P<0.01$). The rate of intracerebral hemorrhage following thrombolysis was 5 (11.1%) in the control group and 2 (4.3%) in the intervention group ($P=0.22$). The intervention group received IVT after the onset of symptoms on average nearly 1 hour earlier than patients in the control group (98 [interquartile range (IQR), 109–204] minutes versus 144 [IQR, 76–124] minutes; adjusted difference in medians was –65 minutes [95% CI, –155.5 to 24.8]; $P=0.15$). Similarly, among those who were evaluated within 4.5 hours, the intervention group received IVT after the onset of symptoms on average nearly 1 hour earlier than patients in the control group (98 [IQR, 76–119] minutes versus 137 [IQR, 106–203]; adjusted difference in median was –64 minutes [95% CI, –101 to –27] minutes; $P=0.001$).

In the intervention group, all patients received IVT before their arrival to the hospital, while none of the control group received IVT before hospital arrival. Numerically, a greater proportion of patients received mechanical thrombectomy (MT) in the intervention group than in the control group (40% versus 23%; adjusted odds ratio, 1.94 [95% CI, 0.93–4.05]; $P=0.08$). The intervention group received MT after the onset of symptoms approximately an hour and a half earlier on average than patients in the control group (216 [IQR, 141–519] minutes versus 161

[IQR, 127–188] minutes; adjusted difference in medians was –91.5 minutes [95% CI, –235.2 to 52.1]; $P<0.21$). No statistically significant difference between groups in the proportion of patients receiving MT within 6 hours of stroke onset (66.7% versus 84%; adjusted odds ratio, 3.32 [95% CI, 0.68–16.18]; $P=0.14$) was observed. The majority of patients with ischemic stroke received MT in ≤ 60 minutes of arrival to the hospital, in the intervention group 24 (96%), compared with only 8 (22.9%) of patients in the control group (Fisher exact test $P<0.001$).

Hospitalization Costs From Stroke Event to 180 Days of Poststroke

Among the total matched sample, there were cost data available for 287 (98%; MSU=92 and standard ambulance users=195) participants (Table 3). The median cost of hospitalizations within 6 months postevent was \$44 548 (IQR, \$12 914–\$78 259) for the intervention group and \$24 418 (IQR, \$9840–\$48 596) for the control group (adjusted difference in medians was \$15 638 [95% CI, 3097–28 179]; $P=0.02$). Moreover, after adjusting the costs for age, sex, stroke diagnosis, hospital destination, day of admission, and NIHSS, the average cost difference between groups was \$17 949 (95% CI, \$4682–\$31 215); $P<0.01$). Similarly, the median cost was \$14 406 (95% CI, 2458–26 355); $P=0.02$) greater among patients with ischemic stroke and transported by MSU compared with those transported by standard ambulances. After adjusting the cost for age, sex, stroke diagnosis, stroke severity measured

Table 2. Time Metrics for Reperfusion Therapy Among Patients With Ischemic Stroke Within the Target Trial, Victoria, 2018

Characteristics	Standard ambulance, n=161	MSU group, n=70	P value
Symptom onset to first evaluation,* min; median (IQR)	173 (85–498)	131 (95–230)	0.34
Symptom onset to first evaluation* (≤ 4.5 h)	89 (56.3)†	53 (77.9)†	0.01
IVT, n (%)	45 (28.0)	47 (67.1)	<0.001
Onset to needle, min; median (IQR)	144.0 (107.8–204.0)†	98.0 (76.0–124.0)†	<0.001
Treated before hospital arrival	0 (0.0)†	39 (100%)†	<0.001
Treated ≤ 60 min, n (%) onset to needle	1 (2.3)†	5 (12.8)†	0.06
Treated ≤ 60 min, n (%) hospital arrival (door) to needle	26 (59.1)†	NA	NA
Adverse events after IVT‡	5 (11.1)	2 (4.3)	0.22
Mechanical thrombectomy (yes)	37 (23.0)	28 (40.0)	<0.01
Onset to MT, min; median (IQR)	216 (140.5–518.5)†	161 (127–188)†	0.07
MT ≤ 6 h, n (%) onset to MT	24 (66.7)†	21 (84.0)†	0.13
MT ≤ 60 min, n (%) hospital arrival to MT	8 (28.9)†	24 (96.0)†	<0.001§

P values for nonparametric continuous variables were calculated according to Wilcoxon rank-sum (Mann-Whitney U test) test, while categorical variables calculated using χ^2 . NA refers to not applicable, because, all participants received treatment before their arrival to the hospital. Symptom onset is defined as the time the patient was last known to be well or call received by an ambulance Victoria call center. All the tissue-type plasminogen activator and MT time metrics percentages and missing values are calculated among those receiving the treatments. Due to rounding, percentages may not add up to 100%. IQR indicates interquartile range; IVT, intravenous thrombolysis; MSU, mobile stroke unit, and MT, mechanical thrombectomy.

*Upon arrival at hospital for standard ambulance, and at the time of triage by the MSU for the MSU group.

†One to 8 missing observations.

‡Adverse events include intracranial hemorrhage, extracranial hemorrhage, and angioedema.

§Fisher exact test.

by NIHSS, hospital destination, and day of admission, costs were \$23 226 (95% CI, 6633–39 818; $P<0.01$) greater on average in the intervention group compared with the control group.

Health-Related Outcomes From Stroke Event to 90 to 180 Days of Poststroke

There were several differences in outcomes between groups. The intervention group significantly reduced the hazard of death during 180 days compared with the control group (hazard ratio, 0.53 [95% CI, 0.32–0.86]; Table 4). In addition, the intervention group improved nondisability rate compared with the control group (Figure S6) at the level of evidence against the null hypothesis of $P=0.05$ (odds ratio, 2.11 [95% CI, 1.07–4.18]). Further secondary analysis using linear regression of the utility-weighted mRS³⁴ indicated an adjusted beneficial effect in the intervention group, with an adjusted β of 0.081 (95% CI, 0.0–0.16; $P=0.05$). Patients in the intervention group had greater odds of hospital admissions after ischemic stroke than the control group (odds ratio, 1.94 [95% CI, 1.02–3.72]; $P=0.04$). In general, our causal analysis estimate indicated that the use of MSU service could reduce disability and mortality at 6 months.

DISCUSSION

In this study, we found that the Melbourne MSU facilitated faster and more frequent provision of reperfusion therapies. There were also improved rates

of nondisability and survival within 6 months of stroke, which is likely attributable to the increased access to reperfusion therapies. Moreover, patients with intracerebral hemorrhage are provided treatments such as antihypertensive therapy, hemostatic medications, and antiepileptic medications while in the MSU, and are transported directly to hospitals with neurosurgical capability. Reassuringly, the point estimates for the magnitude of the MSU effect are qualitatively similar in the overall patient cohort and the subset of patients with ischemic stroke. Therefore, it is more likely that the reduction in the sample size, when only patients with ischemic stroke are considered, leads to the reduced precision of the effect estimate. The findings related to improved provision of reperfusion therapies were consistent with other studies from similar settings with various research designs,^{11,12,35–37} but in some of these studies, no differences in nondisability rate^{36,38} or mortality^{35,38} were observed between groups. A greater cost of hospitalization per patient in the intervention group compared with the control group was observed and is likely due to the difference in receipt of reperfusion treatments between groups. Within the Australian context, providing MT is estimated to be approximately AUD 18 000 per procedure.³⁹

The proportion of patients provided MT was greater in the intervention group than in the control group, but there were no differences in onset-to-groin time in the first 6 hours, which is consistent with previous findings from a similar setting.¹² A modest reduction in onset-to-groin time was not statistically significant, and this could

Table 3. Hospitalization Costs From Stroke Event to 180 Days Poststroke Among Our Cohort Within the Target Trial, Victoria, 2018

	Standard ambulance (n=195)*	MSU group (n=80)*
	Costs AUD, median (IQR)	Costs AUD, median (IQR)
Stroke event		
ED presentations	1938 (1454–3208)	2225 (1173–2994)
Urgent	1951 (1484–3263)	2225 (1173–2994)
Nonurgent	1356 (890–2203)	...
Acute admission	10 137 (4587–20979)	13 365 (6051–26384)
After discharge		
ED presentations	1509 (908–2560)	1902 (894–3582)
Urgent	1736 (1149–2663)	1930 (1146–3833)
Nonurgent	809 (437–1301)	697 (66–1100)
Acute admissions	5017 (1499–15595)	12 613 (5575–32342)
Subacute admissions		
Admitted rehabilitation	23 940 (12394–39108)	44 895 (18171–77899)
Admitted palliative care	4125 (1120–4862)	1582 (869–3230)
Other subacute†	20 616 (12252–48201)	21 309 (12341–39049)

Subacute costs related to the stroke event for 1 patient in the standard ambulance group and 1 patient in the MSU group have not been shown. AUD indicates Australian Dollar in 2018 rate; ED, emergency department; IQR, interquartile range; and MSU, mobile stroke unit.

*While we used multiple imputation for hospital records with missing costs, only 195 of 198 patients in the control group and 80 of 96 patients in the intervention group were able to linked to hospital records.

†GEM, Geriatric evaluation and management, psychogeriatric, and admitted nonacute.

Table 4. Outcomes Associated With Use of Mobiles Stroke Units Compared With Standard Ambulance Users Among Our Cohort Within the Target Trial, Victoria, 2018

Characteristics	Standard ambulance, n=198	MSU group, n=96	Adjusted OR/HR (95% CI)	P value
Patients with ischemic/hemorrhagic stroke				
Nondisabled (0–1 mRS)	47 (23.7)	26 (27.1)	2.11 (1.07–4.18)	0.03
Independent (0–2 mRS)	68 (34.3)	33 (34.4)	1.70 (0.89–3.24)	0.11
Ambulatory (0–3 mRS)	104 (52.5)	46 (47.9)	1.29 (0.68–2.44)	0.44
Requires constant care or deceased (5–6 mRS)	76 (38.4)	37 (38.5)	0.56 (0.28–1.10)	0.09
Mortality (<180 d)	70 (35.4)	30 (31.3)	0.53 (0.32–0.86)	0.01
Hospital admission after stroke	114 (57.6)	61 (63.5)	1.56 (0.89–2.76)	0.12
Patients with ischemic stroke				
	n=161	n=70		
Nondisabled (0–1 mRS)	44 (27.3)	23 (32.9)	1.91 (0.93–3.92)	0.08
Independent (0–2 mRS)	63 (39.1)	30 (42.9)	1.63 (0.83–3.21)	0.16
Ambulatory (0–3 mRS)	93 (57.8)	41 (58.6)	1.35 (0.67–2.71)	0.40
Requires constant care or deceased (5–6 mRS)	55 (34.2)	20 (28.6)	0.50 (0.23–1.07)	0.07
Mortality (<180 d)	51 (31.7)	13 (18.6)	0.34 (0.17–0.66)	<0.01
Hospital admission after stroke	96 (59.6)	51 (72.9)	1.94 (1.02–3.72)	0.04

Values are presented as frequency (%). *P* values were calculated using χ^2 or logistic regression for disability status and hazard ratios for mortality outcome. The multivariable regressions were adjusted for age, sex, stroke diagnosis (ischemic stroke and hemorrhagic stroke) when applicable, day of admission, stroke severity as measured according to the National Institutes of Health Stroke Scale, and hospital destination of the patient. Due to rounding, percentages may not add up to 100%. HR indicates hazard ratio; mRS, modified Rankin Scale; MSU, mobile stroke unit; and OR, odds ratio.

be attributed to the challenges in integrating MSU into established systems,⁴⁰ and additional requests for new brain imaging (such as computed tomography perfusion for suspected large vessel occlusions with an onset >6 hours prior) or complexities in identifying candidates for MT, but further investigation using larger sample sizes is required. Obtaining brain imaging on board the MSU, administering reperfusion therapies on board the MSU, providing prenotification to the endovascular team, and directly transferring to MT-capable hospitals improved the timelines of MT and outcomes after stroke elsewhere.^{41,42} Every 15 minutes in time savings (ie, between the first 30 and 270 minutes) is associated with a 14% greater chance of ambulation by the time of discharge.⁴³ Conversely, for every 30-minute delay for MT, there is an associated 15% reduction in achieving a good clinical outcome. Moreover, every minute of delay in hospital-to-hospital transfer reduces the probability of patients receiving the therapy by nearly 3%.⁴⁴ Therefore, further investigation on the number and frequency of imaging, especially for patients with extended or unclear time windows, is required, especially for patients managed by an MSU suspected of having a large vessel occlusion.

The operational and maintenance expenses related to the Melbourne MSU have been published previously,¹⁰ and were not considered in this analysis. The outputs from this analysis will inform economic modeling of the costs and benefits of the Melbourne MSU over a patient's lifetime. There is evidence from economic modeling that the MSU as a model of prehospital stroke care can be cost-effective.^{7,10,17,19} Cost-effectiveness is improved when MSUs can serve a large population with frequent

cases of suspected stroke.^{7,45} Expanding the MSU service in Australia and other similar settings could be beneficial in saving lives and reducing disability, but the initial set-up and operational costs need to be strategically considered based on volume for optimizing the health system and ensuring more people get timely access to evidence-based care. Economic benefits from improving outcomes are likely to be seen beyond the time horizon used in this study (6 months), as observed in evaluations of IVT and MT.^{46,47} Economic impacts beyond the hospital perspective also need to be considered (such as ongoing health care utilization and productivity impacts over a lifetime) since reducing disability from reperfusion therapies may impact employment and long-term health service utilization. There is evidence from other studies that the upfront and ongoing financial investment in MSUs pays off over a patient's lifetime.^{7,17}

The strength of this evaluation is the use of real-world clinical quality registry data augmented using a linkage to government and ambulance services data sets to determine the costs and benefits of transporting a patient using MSU compared with the standard ambulances. There is only 1 ambulance provided for emergency care in Victoria, and the AuSCR is used across all hospitals within the MSU catchment. We also used strong statistical techniques to ensure reliable comparisons between groups with the ability to assess causality. However, due to our careful selection of patients for the target trial emulation, data from only a small proportion of patients were included, limiting statistical power and generalizability. The sample included for comparison was younger and more often male than that observed in the AuSCR median age of 79 versus 73 years and 58% versus

50% males.²⁴ Our study is one of the few to illustrate this approach in strengthening the use of observational, population-level data and shows that even with a large sample size, attrition of cases can be large.

A limitation of the study was that having a dual dispatch for MSU may have introduced a selection bias as evidenced by the greater NIHSS score among patients transported by MSU. Despite our efforts to control for differences between groups, there may still be selection bias during group assignment that might have been avoided with a prospective randomized controlled trial. Secondly, the precise timing of the mRS assessment at follow-up was variable, and the time frame we provided (ie, 90–180 days) is an approximate measure. Thirdly, we did not have information about where the pick-up location was (ie, home or elsewhere). Therefore, the pick-up postcode may not truly reflect a given patient's socio-economic position. Another limitation is that the type of adverse event after tissue-type plasminogen activator (eg, symptomatic intracerebral hemorrhage) was not reported in our findings, and it was not possible to investigate how these events might affect the cost of stroke. Hospitals included in this study were public hospitals, which is where the majority of patients with stroke are treated within the Australian health care system (particularly those that are transported by ambulance). The costs of private hospitals were not included and this may limit the generalizability of our results. Lastly, we acknowledge that costs after the initial hospitalization may also include costs that are unrelated to stroke.

CONCLUSIONS

With the context of Melbourne, Australia, we confirm that the MSU program provides greater access to reperfusion therapies, leading to improved functional outcomes and fewer deaths. Expanding the MSU service in Australia and other similar settings could be beneficial in saving lives and reducing disability. Cost implications that are likely related to increased provision of reperfusion therapies warranted further investigation.

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Supplemental Material

Supplemental Methods
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Appendix S1: STROBE Checklist
Appendix S2: CHEERS Checklist

REFERENCES

- Feigin VL, Stark BA, Johnson CO, Roth GA, Bisignano C, Abady GG, Abbasifard M, Abbasi-Kangevari M, Abd-Allah F, Abedi V, et al. Global, regional, and national burden of stroke and its risk factors, 1990–2013: a systematic analysis for the Global Burden of Disease Study 2019. *Lancet Neurol*. 2021;20:795–820. doi: 10.1016/S1474-4422(21)00252-0
- Greenberg SM, Ziai WC, Cordonnier C, Dowlathshahi D, Francis B, Goldstein JN, Hemphill JC, Johnson R, Keigher KM, Mack WJ, et al; American Heart Association/American Stroke Association. 2022 Guideline for the management of patients with spontaneous intracerebral hemorrhage: a guideline from the American Heart Association/American Stroke Association. *Stroke*. 2022;53:e282–e361. doi: 10.1161/STR.0000000000000407
- Cadilhac D, Dalli LL, Morrison J, Paice K, Carter H, Cambell B, Cloud GC, Kilkenny MF, Faux SG, Hill K, et al. The Australian Stroke Clinical Registry Annual Report 2022. 2024.
- Fassbender K, Grotta JC, Walter S, Grunwald IQ, Ragoschke-Schumm A, Saver JL. Mobile stroke units for prehospital thrombolysis, triage, and beyond: benefits and challenges. *Lancet Neurol*. 2017;16:227–237. doi: 10.1016/S1474-4422(17)30008-X
- Fassbender K, Merzou F, Lesmeister M, Walter S, Grunwald IQ, Ragoschke-Schumm A, Bertsch T, Grotta J. Impact of mobile stroke units. *J Neurol Neurosurg Psychiatry*. 2021;92:815–822. doi: 10.1136/jnnp-2020-324005
- Bowry R, Grotta JC. Mobile stroke units: current and future impact on stroke care. *Semin Neurol*. 2021;41:9–15. doi: 10.1055/s-0040-1722724
- Lund UH, Stoinska-Schneider A, Larsen K, Bache KG, Robberstad B. Cost-effectiveness of mobile stroke unit care in Norway. *Stroke*. 2022;53:3173–3181. doi: 10.1161/strokeaha.121.037491
- Walter S, Kostopoulos P, Haass A, Keller I, Lesmeister M, Schlechtriemen T, Roth C, Papanagiotou P, Grunwald I, Schumacher H, et al. Diagnosis and treatment of patients with stroke in a mobile stroke unit versus in hospital: a randomised controlled trial. *Lancet Neurol*. 2012;11:397–404. doi: 10.1016/S1474-4422(12)70057-1
- Zhao H, Coote S, Easton D, Langenberg F, Stephenson M, Smith K, Bernard S, Cadilhac DA, Kim J, Bladin CF, et al. Melbourne Mobile Stroke Unit and reperfusion therapy: greater clinical impact of thrombectomy than thrombolysis. *Stroke*. 2020;51:922–930. doi: 10.1161/STROKEAHA.119.027843
- Kim J, Easton D, Zhao H, Coote S, Sookram G, Smith K, Stephenson M, Bernard S, Parsons M W, Yan B, et al. Economic evaluation of the Melbourne mobile stroke unit. *Int J Stroke*. 2021;16:466–475. doi: 10.1177/1747493020929944
- Ebinger M, Siegerink B, Kunz A, Wendt M, Weber JE, Schwabauer E, Geisler F, Freitag E, Lange J, Behrens J, et al; Berlin_PRehospital Or Usual Delivery in stroke care (B_PROUD) Study Group. Association between dispatch of mobile stroke units and functional outcomes among patients with acute ischemic stroke in Berlin. *JAMA*. 2021;325:454–466. doi: 10.1001/jama.2020.26345

12. Grotta JC, Yamal JM, Parker SA, Rajan SS, Gonzales NR, Jones WJ, Alexandrov AW, Navi BB, Nour M, Spokoyny I, et al. Prospective, multicenter, controlled trial of mobile stroke units. *N Engl J Med*. 2021;385:971–981. doi: 10.1056/NEJMoa2103879
13. Turc G, Hadziahmetovic M, Walter S, Churilov L, Larsen K, Grotta JC, Yamal JM, Bowry R, Katsanos AH, Zhao H, et al. Comparison of mobile stroke unit with usual care for acute ischemic stroke management: a systematic review and meta-analysis. *JAMA Neurol*. 2022;79:281–290. doi: 10.1001/jamaneurol.2021.5321
14. Zhou T, Zhu L, Wang M, Li T, Li Y, Pei Q, Chen W, Zhao J, Wu H, Liu H, et al. Application of Mobile Stroke Unit in prehospital thrombolysis of acute stroke: experience from China. *Cerebrovasc Dis*. 2021;50:520–525. doi: 10.1159/000514370
15. Zheng B, Li Y, Gu G, Yang J, Jiang J, Chen Z, Fan Y, Wang S, Pei H, Wang J. Comparing 5G mobile stroke unit and emergency medical service in patients acute ischemic stroke eligible for t-PA treatment: a prospective, single-center clinical trial in Ya'an, China. *Brain Behav*. 2023;13:e3231. doi: 10.1002/brb3.3231
16. Nilanont Y, Chanyagorn P, Shukij K, Pengtong W, Kongmuangpuk M, Wongmayurachat K, Nittayaboon K, Wongsawat Y, Sirovetnukul R, Chakorn T, et al. Comparing performance measures and clinical outcomes between mobile stroke units and usual care in underserved areas. *Neurol Sci*. 2023;44:1261–1271. doi: 10.1007/s10072-022-06550-6
17. Rink JS, Froelich MF, Nour M, Saver JL, Szabo K, Hoyer C, Fassbender KC, Schoenberg SO, Tollens F. Lifetime economic potential of mobile stroke units in acute stroke care: a model-based analysis of the drivers of cost-effectiveness. *J Telemed Telecare*. 2024;30:1335–1344. doi: 10.1177/1357633x221140951
18. Reimer AP, Zafar A, Hustey FM, Kralovic D, Russman AN, Uchino K, Hussain MS, Udeh BL. Cost-consequence analysis of Mobile Stroke Units vs. Standard Prehospital Care and Transport. *Front Neurol*. 2019;10:1422. doi: 10.3389/fneur.2019.01422
19. Oliveira Gonçalves AS, Rohmann JL, Piccininni M, Kurth T, Ebinger M, Endres M, Freitag E, Harmel P, Lorenz-Meyer I, Rohrpascher-Napierkowski I, et al. Economic evaluation of a Mobile Stroke Unit service in Germany. *Ann Neurol*. 2023;93:942–951. doi: 10.1002/ana.26602
20. Navi BB, Audebert HJ, Alexandrov AW, Cadilhac DA, Grotta JC, Group PW. Mobile stroke units: evidence, gaps, and next steps. *Stroke*. 2022;53:2103–2113. doi: 10.1161/STROKEAHA.121.037376
21. Hernán MA, Robins JM. Using big data to emulate a target trial when a randomized trial is not available. *Am J Epidemiol*. 2016;183:758–764. doi: 10.1093/aje/kwv254
22. von Elm E, Altman DG, Egger M, Pocock SJ, Gøtzsche PC, Vandenbroucke JP. The Strengthening of Reporting of Observational Studies in Epidemiology (STROBE) statement: guidelines for reporting observational studies. *Prev Med*. 2007;45:247–251. doi: 10.1016/j.ypmed.2007.08.012
23. Husereau D, Drummond M, Augustovski F, de Bekker-Grob E, Briggs AH, Carswell C, Caulley L, Chaiyakunapruk N, Greenberg D, Loder E, et al; CHEERS 2022 ISPOR Good Research Practices Task Force. Consolidated Health Economic Evaluation Reporting Standards 2022 (CHEERS 2022) statement: updated reporting guidance for health economic evaluations. *BMC Med*. 2022;20:23. doi: 10.1186/s12916-021-02204-0
24. Eliakundu AL, Smith K, Kilkenny MF, Kim J, Bagot KL, Andrew E, Cox S, Bladin CF, Cadilhac DA. Linking data from the Australian Stroke Clinical Registry with ambulance and emergency administrative data in Victoria. *Inquiry*. 2022;59:469580221102200. doi: 10.1177/00469580221102200
25. Cadilhac DA, Lannin NA, Anderson CS, Levi CR, Faux S, Price C, Middleton S, Lim J, Thrift AG, Donnan GA. Protocol and pilot data for establishing the Australian Stroke Clinical Registry. *Int J Stroke*. 2010;5:217–226. doi: 10.1111/j.1747-4949.2010.00430.x
26. Breen S, Cadilhac D, Lannin N, Kim J, Dalli L, Anderson C, Kilkenny M, Shehata S, Faux S, Dewey H, et al. The Australian Stroke Clinical Registry Annual Report 2018. Accessed December 2019. <https://auscr.com.au/>
27. Zhu Y, Hubbard RA, Chubak J, Roy J, Mitra N. Core concepts in pharmacoeconomics: violations of the positivity assumption in the causal analysis of observational data: consequences and statistical approaches. *Pharmacoeconomics Drug Saf*. 2021;30:1471–1485. doi: 10.1002/pds.5338
28. Zakrisson TL, Austin PC, McCredie VA. A systematic review of propensity score methods in the acute care surgery literature: avoiding the pitfalls and proposing a set of reporting guidelines. *Eur J Trauma Emerg Surg*. 2018;44:385–395. doi: 10.1007/s00068-017-0786-6
29. Zhang Z, Kim HJ, Lonjon G, Zhu Y; written on behalf of AME Big-Data Clinical Trial Collaborative Group. Balance diagnostics after propensity score matching. *Ann Transl Med*. 2019;7:16. doi: 10.21037/atm.2018.12.10
30. Austin PC. Balance diagnostics for comparing the distribution of baseline covariates between treatment groups in propensity-score matched samples. *Stat Med*. 2009;28:3083–3107. doi: 10.1002/sim.3697
31. Saver JL, Chaisinanunkul N, Campbell BCV, Grotta JC, Hill MD, Khatri P, Landen J, Lansberg MG, Venkatasubramanian C, Albers GW; Xlth Stroke Treatment Academic Industry Roundtable. Standardized nomenclature for modified Rankin Scale global disability outcomes: consensus recommendations from stroke therapy academic industry roundtable XI. *Stroke*. 2021;52:3054–3062. doi: 10.1161/STROKEAHA.121.034480
32. Banks JL, Marotta CA. Outcomes validity and reliability of the modified Rankin Scale: implications for stroke clinical trials. *Stroke*. 2007;38:1091–1096. doi: 10.1161/01.STR.0000258355.23810.c6
33. van Swieten JC, Koudstaal PJ, Visser MC, Schouten HJ, van Gijn J. Interobserver agreement for the assessment of handicap in stroke patients. *Stroke*. 1988;19:604–607. doi: 10.1161/01.str.19.5.604
34. Dijkland SA, Voormolen DC, Venema E, Roozenbeek B, Polinder S, Haagsma JA, Nieboer D, Chalos V, Yoo AJ, Schreuders J, et al; MR CLEAN Investigators. Utility-weighted modified Rankin Scale as primary outcome in stroke trials: a simulation study. *Stroke*. 2018;49:965–971. doi: 10.1161/STROKEAHA.117.020194
35. Ebinger M, Winter B, Wendt M, Weber JE, Waldschmidt C, Rozanski M, Kunz A, Koch P, Kellner PA, Gierhake D, et al; STEMO Consortium. Effect of the use of ambulance-based thrombolysis on time to thrombolysis in acute ischemic stroke: a randomized clinical trial. *JAMA*. 2014;311:1622–1631. doi: 10.1001/jama.2014.2850
36. Bender MT, Mattingly TK, Rahmani R, Proper D, Burnett WA, Burgett JL, LEsperance J, Cushman JT, Pilcher WH, Benesch CG, et al. Mobile stroke care expedites intravenous thrombolysis and endovascular thrombectomy. *Stroke Vasc Neurol*. 2022;7:209–214. doi: 10.1136/svn-2021-001119
37. Kunz A, Ebinger M, Geisler F, Rozanski M, Waldschmidt C, Weber JE, Wendt M, Winter B, Zieschang K, Fiebach JB, et al. Functional outcomes of pre-hospital thrombolysis in a mobile stroke treatment unit compared with conventional care: an observational registry study. *Lancet Neurol*. 2016;15:1035–1043. doi: 10.1016/S1474-4422(16)30129-6
38. Tsvigoulis G, Geisler F, Katsanos AH, Körn J, Kunz A, Mikulik R, Rozanski M, Wendt M, Audebert HJ. Ultraearly intravenous thrombolysis for acute ischemic stroke in Mobile Stroke Unit and hospital settings. *Stroke*. 2018;49:1996–1999. doi: 10.1161/STROKEAHA.118.021536
39. Arora N, Makino K, Tilden D, Lobotesis K, Mitchell P, Gillespie J. Cost-effectiveness of mechanical thrombectomy for acute ischemic stroke: an Australian payer perspective. *J Med Econ*. 2018;21:799–809. doi: 10.1080/13696998.2018.1474746
40. Bagot KL, Purvis T, Hancock S, Zhao H, Coote S, Easton D, Campbell BC, Davis SM, Donnan GA, Foster S, et al. Sustaining a new model of acute stroke care: a mixed-method process evaluation of the Melbourne Mobile Stroke Unit. *Int J Health Policy Manag*. 2023;12:1–13.
41. Czap AL, Singh N, Bowry R, Jagolino-Cole A, Parker SA, Phan K, Wang M, Sheth SA, Rajan SS, Yamal JM, et al. Mobile Stroke Unit computed tomography angiography substantially shortens door-to-puncture time. *Stroke*. 2020;51:1613–1615. doi: 10.1161/STROKEAHA.119.028626
42. Requena M, Olivé-Gadea M, Muchada M, Hernández D, Rubiera M, Boned S, Piñana C, Deck M, García-Tornel A, Díaz-Silva H, et al. Direct to angiography suite without stopping for computed tomography imaging for patients with acute stroke: a randomized clinical trial. *JAMA Neurol*. 2021;78:1099–1107. doi: 10.1001/jamaneurol.2021.2385
43. Jahan R, Saver JL, Schwamm LH, Fonarow GC, Liang L, Matsouka RA, Xian Y, Holmes DN, Peterson ED, Yavagal D, et al. Association between time to treatment with endovascular reperfusion therapy and outcomes in patients with acute ischemic stroke treated in clinical practice. *JAMA*. 2019;322:252–263. doi: 10.1001/jama.2019.8286
44. Prabhakaran S, Ward E, John S, Lopes DK, Chen M, Temes RE, Mohammad Y, Lee VH, Bleck TP. Transfer delay is a major factor limiting the use of intra-arterial treatment in acute ischemic stroke. *Stroke*. 2011;42:1626–1630. doi: 10.1161/STROKEAHA.110.609750
45. Gyrdd-Hansen D, Olsen KR, Bollweg K, Kronborg C, Ebinger M, Audebert HJ. Cost-effectiveness estimate of prehospital thrombolysis: results of the PHANTOM-S study. *Neurology*. 2015;84:1090–1097. doi: 10.1212/WNL.0000000000001366
46. Joo H, Wang G, George MG. A literature review of cost-effectiveness of intravenous recombinant tissue plasminogen activator for treating acute ischaemic stroke. *Stroke Vasc Neurol*. 2017;2:73–83. doi: 10.1136/svn-2016-000063
47. Wu X, Khunte M, Gandhi D, Sanelli P, Forman HP, Malhotra A. A systematic review of cost-effectiveness analyses on endovascular thrombectomy in ischemic stroke patients. *Eur Radiol*. 2022;32:3757–3766. doi: 10.1007/s00330-022-08671-0